

Accounting for Post-Crisis Inflation and Employment: A Retro Analysis*

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Abstract

What accounts for inflation after 2008? We use the prominent pre-crisis Smets-Wouters (2007) model to address this question. We find that due to price markup shocks alone inflation would have been 1% higher than observed and 0.5% higher than the long-run average. Their standard deviation is similar to its pre-crisis level. Price markup shocks were also responsible for the slow recovery of employment, though not for the initial drop. Monetary policy shocks predict an inflation rate 0.5% below average. Government expenditure innovations do not contribute much either to inflation or to employment dynamics.

**Technical Appendix — NOT FOR
PUBLICATION**

Table 1: Summary statistics

| | 1948Q1 - 2004Q4 | | 1948Q1 - 2014Q2 | |
|---------------|-----------------|---------|-----------------|---------|
| | mean | std dev | mean | std dev |
| consumption | 0.535 | 0.846 | 0.480 | 0.834 |
| investment | 0.565 | 2.380 | 0.442 | 2.414 |
| output | 0.514 | 1.002 | 0.455 | 0.973 |
| labor | 0.825 | 2.900 | 0.000 | 3.650 |
| inflation | 0.8400 | 0.652 | 0.789 | 0.624 |
| wage rate | 0.478 | 0.627 | 0.433 | 0.721 |
| interest rate | 1.325 | 0.867 | 1.192 | 0.884 |

Growth rates of consumption, investment, output and real wage. Consumption, investment, output and hours worked are net of labor force growth, normalized to 1 in 2005Q1. Variables are deflated using the GDP deflator.

A Re-Estimating the Smets-Wouters model

Following Smets and Wouters, we adopt a Bayesian approach and we estimate the posterior distribution using a MH algorithm with 250,000 simulations. Priors are the same as the ones assumed in Smets-Wouters. Like them, we use 7 time series to estimate the model: output, consumption, investment, inflation, hours worked, real wages, and the (Federal Fund) interest rate. We compute consumption, investment and hours worked per-capita using a population index normalized to 1 in 1992Q4. We use the GDP deflator as a measure of inflation. Consumption, investment, output and wage are at constant prices. Variables are net of the labor force growth, normalized to 1 in 2005Q1. For the interest rate, we use the Federal Funds Rate.

We use data from U.S. Bureau of Economic Analysis and U.S. Bureau of Labor Statistics for the period 1948Q1-2014Q2. Table 1 collects some summary statistics for the the sample used by Smets and Wouter (left column) and our sample (right column). In general, the mean is smaller and unsurprisingly the standard deviation is larger when we consider the extended

sample. The average of labor hours relative to the hours worked in 2005 Q2 (2005Q2=0) decreases considerably, from 1.348 to 0.663 with the inclusion of the observations from 2005Q1 to 2014Q2. Contrary to the original paper, our sample starts from 1948Q1 instead of 1947Q1 because these data points are not available for the most recent vintages of the time series.

Table 2: Posterior distributions (Metropolis Hasting 250000 simulations). Comparison between our sample and the sample used by Smets and Wouters.

| | | 1948Q1 - 2014Q2 | | | 1948Q1 - 2004Q4 | | |
|-------------------|----------------|-----------------|---------|---------|-----------------|---------|---------|
| | | mean | HDP inf | HPD sup | mean | HDP inf | HPD sup |
| $\sigma(e_Z)$ | TFP | 0.507 | 0.468 | 0.551 | 0.522 | 0.467 | 0.579 |
| $\sigma(e_{b^2})$ | risk premium | 0.121 | 0.096 | 0.146 | 0.112 | 0.069 | 0.153 |
| $\sigma(e_g)$ | gov't exp | 0.645 | 0.597 | 0.689 | 0.734 | 0.671 | 0.794 |
| $\sigma(e_\mu)$ | | 0.477 | 0.412 | 0.536 | 0.435 | 0.333 | 0.529 |
| $\sigma(e_{ms})$ | monetary shock | 0.224 | 0.205 | 0.244 | 0.236 | 0.212 | 0.258 |
| $\sigma(e_p)$ | price shock | 0.204 | 0.176 | 0.233 | 0.216 | 0.184 | 0.247 |
| $\sigma(e_w)$ | wage shock | 0.344 | 0.313 | 0.378 | 0.250 | 0.219 | 0.281 |
| ρ_z | | 0.985 | 0.977 | 0.991 | 0.974 | 0.958 | 0.991 |
| ρ_{b^2} | | 0.806 | 0.740 | 0.877 | 0.852 | 0.769 | 0.934 |
| ρ_g | | 0.977 | 0.967 | 0.988 | 0.978 | 0.964 | 0.991 |
| ρ_μ | | 0.801 | 0.717 | 0.886 | 0.673 | 0.565 | 0.784 |
| ρ_{ms} | | 0.160 | 0.071 | 0.247 | 0.137 | 0.049 | 0.220 |
| ρ_p | | 0.977 | 0.959 | 0.996 | 0.979 | 0.961 | 0.997 |
| ρ_w | | 0.964 | 0.939 | 0.989 | 0.965 | 0.944 | 0.988 |
| θ_p | price MA | 0.873 | 0.809 | 0.934 | 0.872 | 0.805 | 0.941 |
| θ_w | wage MA | 0.932 | 0.894 | 0.971 | 0.890 | 0.838 | 0.943 |

Continued on next page

Table 2 – continued from previous page

| | | 1948Q1 - 2014Q2 | | | 1948Q1 - 2004Q4 | | |
|---------------|-------------------|-----------------|---------|---------|-----------------|---------|---------|
| | | mean | HDP inf | HPD sup | mean | HDP inf | HPD sup |
| $S''(\gamma)$ | inv. adj cost | 3.819 | 2.437 | 5.141 | 6.747 | 5.094 | 8.391 |
| σ_c | intertemp elast | 2.042 | 1.634 | 2.453 | 1.450 | 1.082 | 1.817 |
| h | habits | 0.348 | 0.265 | 0.427 | 0.506 | 0.421 | 0.585 |
| ζ_w | wage stickiness | 0.781 | 0.713 | 0.847 | 0.749 | 0.679 | 0.816 |
| ν_L | labor | 1.471 | 0.774 | 2.095 | 1.971 | 1.135 | 2.770 |
| ζ_p | price stickiness | 0.667 | 0.586 | 0.749 | 0.608 | 0.531 | 0.680 |
| ι_w | wage index | 0.494 | 0.317 | 0.676 | 0.523 | 0.359 | 0.682 |
| ι_p | price index | 0.300 | 0.163 | 0.433 | 0.311 | 0.168 | 0.451 |
| $czcap$ | capital utiliz. | 0.7525 | 0.608 | 0.893 | 0.453 | 0.316 | 0.596 |
| Φ | fixed cost | 1.608 | 1.497 | 1.718 | 1.441 | 1.318 | 1.569 |
| ψ_1 | monetary p. | 2.034 | 1.771 | 2.299 | 1.982 | 1.691 | 2.266 |
| ρ_R | monetary p. | 0.880 | 0.855 | 0.907 | 0.884 | 0.858 | 0.910 |
| ψ_2 | monetary p. | 0.082 | 0.044 | 0.122 | 0.084 | 0.015 | 0.146 |
| ψ_3 | monetary p. | 0.189 | 0.155 | 0.221 | 0.223 | 0.178 | 0.269 |
| $const_\pi$ | measurem eq | 0.752 | 0.627 | 0.878 | 0.725 | 0.574 | 0.871 |
| $constbeta$ | measurem eq | 0.261 | 0.090 | 0.410 | 0.251 | 0.094 | 0.403 |
| $const_L$ | measurem eq | -0.583 | -2.747 | 1.491 | 0.558 | -1.719 | 2.750 |
| $ctrend$ | growth rate | 0.460 | 0.433 | 0.488 | 0.469 | 0.440 | 0.499 |
| ρ_{gZ} | gov't exp | 0.622 | 0.496 | 0.761 | 0.629 | 0.495 | 0.761 |
| α | $\frac{k_*}{y_*}$ | 0.181 | 0.127 | 0.237 | 0.192 | 0.129 | 0.252 |

There are some changes in the estimated posteriors once we extend the time period, but they appear to be minor and not particularly remarkable. A comparison is in table ???. There, HPD inf and HPD low are respectively the

lower bound and the upper bound of a 90% HPD interval. In particular, the intertemporal elasticity of substitution σ_c falls a bit, and the habit level as well as the adjustment costs rise. Robustness checks and estimation results for different subsamples and specifications of the model are reported in a technical appendix.

Technical Appendix — Not for Publication

B The model

The model is described in detail in [?]. This is a review, for convenience. Households maximize the sum of the discounted flow of per-period utility with respect to a consumption good, labor, investment, effective capital and capital utilization. The utility function has external habits. Households can transfer resources in the future by investing in physical capital or by buying government bonds. Capital adjustments come at a cost that depends on the investment growth.

Final good producers use a continuum with mass 1 of imperfectly substitutable intermediate goods to produce a final consumption good. They solve a static problem in which they maximize profits subject to a production function and a zero-profit condition. They take the price of the final good and of the intermediate goods as given. Intermediate firms maximize the sum of the discounted flow of profits assuming a Calvo price setting with partial indexation to inflation subject to the demand of intermediate goods by the final good producers. They produce intermediate goods using capital and labor.

Households provide labor to the labor unions. Labor unions sell labor to labor packers that resell the labor to the intermediate firms. Labor packers operate in a competitive market. They buy labor from unions and repackage it according to a Dixit-Stiglitz aggregator. They maximize profits taking as given the intermediate wage and the final wage and subject to a zero profit condition. Labor unions maximize the sum of the discounted flow of profits assuming a Calvo price setting with partial indexation to inflation subject to the demand of labor by the labor packers.

Finally, a Taylor rule and a stochastic process for government expenditure

close the model. Following Smets and Wouters, the government expenditure depends on its lag, an exogenous shock and the technological shock.

B.1 Household

$$\hat{c}_t = \frac{1}{1+h/\gamma} E_t \hat{c}_{t+1} + \frac{h/\gamma}{1+h/\gamma} \hat{c}_{t-1} - \frac{1-h/\gamma}{\sigma_c(1+h/\gamma)} (\hat{R}_t - E_t \hat{\pi}_{t+1}) - \frac{(\sigma_c-1)(w_*^h L_*/c_*)}{\sigma_c(1+h/\gamma)} (E_t \hat{L}_{t+1} - \hat{L}_t) + b_t^2 \quad (1)$$

$$\hat{i}_t = \frac{1}{1+\bar{\beta}\gamma} (\hat{i}_{t-1} + \hat{\beta}\gamma E_t \hat{i}_{t+1} + \frac{1}{\gamma^2 S''} \hat{Q}_t) + \hat{\mu}_t \quad (2)$$

$$\hat{Q}_t = \frac{\sigma_c(1+h/\gamma)}{1-h/\gamma} \hat{b}_t^2 - \hat{R}_t + E_t \hat{\pi}_{t+1} + \frac{r_*^k}{r_*^k + 1 - \delta} E_t \hat{r}_{t+1}^k + \frac{1-\delta}{r_*^k + 1 - \delta} E_t \hat{Q}_{t+1} \quad (3)$$

$$\hat{u}_t = \frac{r_*^k}{a''(1)} \hat{r}_t^k \quad (4)$$

$$\hat{k}_t = \hat{u}_t + \hat{\bar{k}}_{t-1} \quad (5)$$

$$\hat{\bar{k}}_t = (1 - \frac{i_*}{\bar{k}_*}) \hat{\bar{k}}_{t-1} + \frac{i_*}{\bar{k}_*} \gamma^2 S''(\gamma) \mu_t + \frac{i_*}{\bar{k}_*} \hat{i}_t \quad (6)$$

B.2 Firms

$$\hat{y}_t = \Phi \hat{Z}_t + \alpha \Phi \hat{k}_t + (1-\alpha) \Phi \hat{L}_t \quad (7)$$

$$\hat{k}_t = \hat{w}_t - \hat{r}_t^k + \hat{L}_t \quad (8)$$

$$\hat{m}c_t = \alpha \hat{r}_t^k + (1-\alpha) \hat{w}_t - \hat{Z}_t \quad (9)$$

$$\hat{\pi}_t = \frac{(1-\zeta_p \bar{\beta} \gamma)(1-\zeta_p)}{(1+\bar{\beta} \gamma \iota_p) \zeta_p} \frac{1}{(\frac{y_*+\Phi}{y_*}-1) * curvp + 1} \hat{m}c_t + \hat{\lambda}_{p,t} + \frac{\iota_p}{(1+\bar{\beta} \gamma \iota_p)} \hat{\pi}_{t-1} + \frac{\bar{\beta} \gamma}{(1+\bar{\beta} \gamma \iota_p)} E_t \hat{\pi}_{t+1} \quad (10)$$

B.3 Labor unions

$$\begin{aligned}
& (1 + \bar{\beta}\gamma)\hat{w}_t - \hat{w}_{t-1} - \bar{\beta}\gamma E_t \hat{w}_{t+1} = \\
& = \frac{(1 - \zeta_w \bar{\beta}\gamma)(1 - \zeta_w)}{\zeta_w} \frac{1}{(\lambda_w - 1) * curvw + 1} \left[\frac{1}{1 - h/\gamma} \hat{c}_t - \frac{h/\gamma}{1 - h/\gamma} \hat{c}_{t-1} + \nu_L \hat{L}_t - \hat{w}_t \right] \\
& \quad - (1 + \bar{\beta}\gamma \iota_w) \hat{\pi}_t + \iota_w \hat{\pi}_{t-1} + \bar{\beta}\gamma E_t \hat{\pi}_{t+1} + \hat{\lambda}_{w,t}
\end{aligned}$$

B.4 Government and Monetary Policy

$$\hat{R}_t = \rho_R \hat{R}_{t-1} + (1 - \rho_R)(\psi_1 \hat{\pi}_t + \psi_2(\hat{y}_t - \hat{y}_t^{flex})) + \psi_3(\hat{y}_t - \hat{y}_{t-1} - (\hat{y}_t^{flex} - \hat{y}_{t-1}^{flex})) + ms_t \quad (11)$$

B.5 Aggregate resource constraints

$$\frac{c_*}{y_*} \hat{c}_t + \frac{i_*}{y_*} \hat{i}_t + \hat{g}_t + \frac{r_*^k k_*}{y_*} \hat{u}_t = \hat{y}_t \quad (12)$$

B.6 Natural output level

B.6.1 Household

$$\hat{c}_t = \frac{1}{1 + h/\gamma} E_t \hat{c}_{t+1} + \frac{h/\gamma}{1 + h/\gamma} \hat{c}_{t-1} - \frac{1 - h/\gamma}{\sigma_c(1 + h/\gamma)} \hat{R}_t - \frac{(\sigma_c - 1)(w_*^h L_*/c_*)}{\sigma_c(1 + h/\gamma)} (E_t \hat{L}_{t+1} - \hat{L}_t) + b_t^2 \quad (13)$$

$$\hat{i}_t = \frac{1}{1 + \bar{\beta}\gamma} (\hat{i}_{t-1} + \bar{\beta}\gamma E_t \hat{i}_{t+1} + \frac{1}{\gamma^2 S''} \hat{Q}_t) + \hat{\mu}_t \quad (14)$$

$$\hat{Q}_t = \frac{\sigma_c(1 + h/\gamma)}{1 - h/\gamma} \hat{b}_t^2 - \hat{R}_t + \frac{r_*^k}{r_*^k + 1 - \delta} E_t \hat{r}_{t+1}^k + \frac{1 - \delta}{r_*^k + 1 - \delta} E_t \hat{Q}_{t+1} \quad (15)$$

$$\hat{u}_t = \frac{r_*^k}{a''(1)} \hat{r}_t^k \quad (16)$$

$$\hat{k}_t = \hat{u}_t + \hat{\bar{k}}_{t-1} \quad (17)$$

$$\hat{k}_t = (1 - \frac{i_*}{\bar{k}_*}) \hat{k}_{t-1} + \frac{i_*}{\bar{k}_*} \gamma^2 S''(\gamma) \mu_t + \frac{i_*}{\bar{k}_*} \hat{i}_t \quad (18)$$

B.6.2 Firms

$$\hat{y}_t = \Phi \hat{Z}_t + \alpha \Phi \hat{k}_t + (1 - \alpha) \Phi \hat{L}_t \quad (19)$$

$$\hat{k}_t = \hat{w}_t - \hat{r}_t^k + \hat{L}_t \quad (20)$$

$$0 = \alpha \hat{r}_t^k + (1 - \alpha) \hat{w}_t - \hat{Z}_t \quad (21)$$

B.6.3 Labor unions

$$\hat{w}_t = \frac{1}{1 - h/\gamma} \hat{c}_t - \frac{h/\gamma}{1 - h/\gamma} \hat{c}_{t-1} + \nu_L \hat{L}_t \quad (22)$$

B.6.4 Aggregate resource constraints

$$\frac{c_*}{y_*} \hat{c}_t + \frac{i_*}{y_*} \hat{i}_t + \hat{g}_t + \frac{r_*^k k_*}{y_*} \hat{u}_t = \hat{y}_t \quad (23)$$

B.7 Exogenous processes

$$Z_t = \rho_Z Z_{t-1} + e_Z \quad (24)$$

$$b_t^2 = \rho_{b^2} b_{t-1}^2 + e_{b^2} \quad (25)$$

$$g_t = \rho_g g_{t-1} + e_g + \rho_g e_Z \quad (26)$$

$$\mu_t = \rho_\mu \mu_{t-1} + e_\mu \quad (27)$$

$$\lambda_{p,t} = \rho_p \lambda_{p,t-1} + e_p - \theta_p e_{p,t-1} \quad (28)$$

$$\lambda_{w,t} = \rho_w \lambda_{w,t-1} + e_w - \theta_w e_{w,t-1} \quad (29)$$

$$ms_t = \rho_{ms} ms_{t-1} + e_{ms} \quad (30)$$

B.8 Measurement equations

$$dy_t = y_t - y_{t-1} + ctrend \quad (31)$$

$$dc_t = c_t - c_{t-1} + ctrend \quad (32)$$

$$dinve_t = i_t - i_{t-1} + ctrend \quad (33)$$

$$dw_t = w_t - w_{t-1} + ctrend \quad (34)$$

$$pinfobst = \pi_t + const_\pi \quad (35)$$

$$robst = R_t + const_R \quad (36)$$

$$labobst = L_t + const_L \quad (37)$$

C The data

We list the data used in our estimation:

- **GDPC96**: Real Gross Domestic Product, 3 decimal - Billions of Chained 2009 Dollars, Seasonally Adjusted Annual Rate. Source: U.S. Department of Commerce, Bureau of Economic Analysis
- **GDPDEF**: Gross Domestic Product - Implicit Price Deflator - 2009=100, Seasonally Adjusted. Source: U.S. Department of Commerce, Bureau of Economic Analysis
- **PCEC**: Personal Consumption Expenditures - Billions of Dollars, Seasonally Adjusted Annual Rate. Source: U.S. Department of Commerce, Bureau of Economic Analysis
- **FPI** : Fixed Private Investment - Billions of Dollars, Seasonally Adjusted Annual Rate Source: U.S. Department of Commerce, Bureau of Economic Analysis. We are using the latest vintage date, *FPI_20130731* (Source for alternative vintage dates: ALFRED).
- **CE16OV** : Civilian Employment: Sixteen Years & Over, Thousands, Seasonally Adjusted. Source: U.S. Department of Labor: Bureau of Labor Statistics. Start date: 1948Q1.
- **CE16OV index** : CE16OV (2009:1)=1
- **Effective Federal Funds Rate** : Percentage, Averages of Daily Figures Percent, Quarterly. Not Seasonally Adjusted. Source: Board of Governors of the Federal Reserve System. (Before 1954: 3-Month Treasury Bill Rate, Secondary Market Averages of Business Days, Discount Basis. For this time series, we used the data from FRED, TB3MS, 3-Month Treasury Bill: Secondary Market Rate (TB3MS), Percent, Quarterly, Not Seasonally Adjusted.)

- **LNS10000000** : Labor Force Status : Civilian noninstitutional population - Age : 16 years and over - Seasonally Adjusted - Number in thousands. Source: U.S. Bureau of Labor Statistics. (Before 1976: CNP16OV : Civilian Noninstitutional Population level - 16 Years and Older)
- **LNSindex** : LNS10000000(2005:1)=1
- **PRS85006023** - Nonfarm Business, All Persons, Average Weekly Hours Duration : index, 2005 = 100, Seasonally Adjusted. Source : U.S. Department of Labor.
- **COMPNFB** :Nonfarm Business Sector: Compensation Per Hour, Index 2005=100, Quarterly, Seasonally Adjusted. Source : U.S. Department of Labor

Definition of data variables in the measurement equations:

$$consumption = dc = \Delta LN((PCEC/GDPDEF)/LNSindex) * 100 \quad (38)$$

$$investment = dinve = \Delta LN((FPI/GDPDEF)/LNSindex) * 100 \quad (39)$$

$$output = dy = \Delta LN(GDPC96/LNSindex) * 100 \quad (40)$$

$$hours = labobs = LN((PRS85006023 * CE16OV/100)/LNSindex) * 100 \quad (41)$$

$$inflation = pinfobs = LN(GDPDEF/GDPDEF(-1)) * 100 \quad (42)$$

$$realwage = dw = \Delta LN(COMPFB/GDPDEF) * 100 \quad (43)$$

$$interestrate = robs = FederalFundsRate/4 \quad (44)$$

D Estimation procedure

We use a Metropolis-Hasting algorithm with 250,000 iterations to compute the posterior distributions. We used the same priors as Smets-Wouters 2006.

D.1 List of estimated parameters

The following are the parameters estimated in the model.

Parameters for the stochastic processes:

$std\ err\ e_Z, std\ err\ e_{b^2}, std\ err\ e_g, std\ err\ e_\mu, std\ err\ e_{ms}, std\ err\ e_p, std\ err\ e_w,$
 $\rho_Z, \rho_{b^2}, \rho_g, \rho_{gZ}, \rho_\mu, \rho_{ms}, \rho_p, \rho_w, \theta_p, \theta_w$

Investment adjustment cost: $S''(\gamma)$

Utility function: σ_c, h, ν_L ,

Capital utilization parameter: $czcap$

Inflation indexation parameters: ι_w, ι_p

Calvo probability of price resetting: ζ_w, ζ_p

Fixed production cost: Φ

Capital share: α

Taylor rule: $\psi_1, \rho_R, \psi_2, \psi_3$

Measurement equations: $const_\pi, constbeta, const_L, ctrend$

D.2 List of calibrated parameters

The following parameters have been calibrated.

Capital depreciation: $\delta = .025$

Steady state labor markup: $\lambda_{w,*} = 0.5$

Steady state government expenditure: $g_* = 0.18$

Generalization of the Dixit-Stigler aggregators for the intermediate good firms and the labor unions: $curvp = curvw = 10$

D.3 List of derived parameters

The following are function of other parameters:

$$\bar{\beta} = \beta\gamma^{-\sigma_c} \quad (45)$$

$$ctrend = 100(\gamma - 1) \quad (46)$$

$$\lambda_{p,*} = \Phi - 1 \quad (47)$$

$$cr = \frac{cpie}{\beta\gamma^{-\sigma_c}} \quad (48)$$

$$r_{k,*} = \beta^{-1}\gamma^{\sigma_c} - (1 - \delta) \quad (49)$$

$$w_* = \left(\frac{\alpha^\alpha(1-\alpha)^{1-\alpha}}{\Phi r_{k,*}^\alpha} \right)^{\frac{1}{1-\alpha}} \quad (50)$$

$$\frac{i_*}{\bar{k}_*} = 1 - \frac{1 - \delta}{\gamma} \quad (51)$$

$$\frac{k_*}{L_*} = \frac{\alpha}{1 - \alpha} \frac{w_*}{r_{k,*}} \quad (52)$$

$$\frac{k_*}{y_*} = \Phi \left(\frac{k_*}{L_*} \right)^{1-\alpha} \quad (53)$$

$$\frac{i_*}{y_*} = \left(1 - \frac{1 - \delta}{\gamma} \right) \gamma \frac{k_*}{y_*} \quad (54)$$

$$\frac{c_*}{y_*} = 1 - g_* - \frac{i_*}{y_*} \quad (55)$$

$$\frac{w_*^h L_*}{c_*} = \frac{1}{1 + \lambda_{w,*}} \frac{1 - \alpha}{\alpha} r_{k,*} \frac{k_*}{c_*} \quad (56)$$

$$const_R = (cr - 1) * 100 \quad (57)$$

$$const_\pi = (cpie - 1) * 100 \quad (58)$$

$$a'(1) = \frac{1 - czcap}{czcap} r_{k,*} \quad (59)$$

E Robustness checks

In this section we summarize the results from our robustness checks.

Firstly, we want to evaluate if introducing the years of the recession for the estimation significantly modifies the estimates for the posterior distributions. We compared the posterior distribution for two samples: 1948Q1-2004Q1, the sample used by the authors (except for the first 4 data points that we dropped as discussed in the paper) and 1948Q1-2014Q2, the extended sample. We do not estimate the posterior distributions separately for the last part of the sample because 2007-2014 would be too short of a sample period relative to the full sample to guarantee comparability of the results.

Parameters with the largest deviations are: $std\ err\ e_g$, ρ_{b^2} , ρ_μ , $S''(\gamma)$ (investment adjustment cost), σ_c , h , czcap (function of the capital utilization adjustment cost), Φ , α . However, due to the large number of data points in common, the difference is small relative to the confidence intervals of the corresponding parameters.

We estimated the parameters for smaller subsamples, to investigate whether there are major changes in the posterior distributions.

When we consider the periods 1988-1996, 1996-2004 and 2005-2014, the estimate of the coefficient of autocorrelation of the real interest rate becomes progressively larger as we move forward in time. The AR and the MA coefficients for the price markup and the wage markup decrease as we estimate them using subsamples towards the end of the considered period.

Moreover there is some variation in the estimates of the elasticity of intertemporal substitution and the habit formation parameter across subsamples. For the elasticity of intertemporal substitution the lowest value for the mean (0.860 compared to the whole sample estimate of 1.984) is attained in the sample 1988-1996, while in the sample 2005-2014 the mean for the habit formation parameter is twice as large as the corresponding value estimated

with the whole sample.

Following Smets and Wouters, we used a generalization of the Dixit-Stiglitz index for the general specification of the model. However, we tested the implications for the posterior distributions of changing this assumption and we conclude that there are no main changes in the posterior distributions of the parameters if we use the Dixit-Stiglitz aggregator.

Table 3: Posterior distributions (Metropolis Hasting 250000 simulations). Comparison between our sample and the sample used by Smets and Wouters.

| | | 1948Q1 - 2014Q2 | | | 1948Q1 - 2004Q4 | | |
|-------------------|----------------|-----------------|---------|---------|-----------------|---------|---------|
| | | mean | HDP inf | HPD sup | mean | HDP inf | HPD sup |
| $\sigma(e_Z)$ | TFP | 0.503 | 0.458 | 0.547 | 0.522 | 0.467 | 0.579 |
| $\sigma(e_{b^2})$ | risk premium | 0.136 | 0.099 | 0.172 | 0.112 | 0.069 | 0.153 |
| $\sigma(e_g)$ | gov't exp | 0.647 | 0.602 | 0.695 | 0.734 | 0.671 | 0.794 |
| $\sigma(e_\mu)$ | | 0.488 | 0.421 | 0.554 | 0.435 | 0.333 | 0.529 |
| $\sigma(e_{ms})$ | monetary shock | 0.226 | 0.206 | 0.247 | 0.236 | 0.212 | 0.258 |
| $\sigma(e_p)$ | price shock | 0.208 | 0.180 | 0.235 | 0.216 | 0.184 | 0.247 |
| $\sigma(e_w)$ | wage shock | 0.278 | 0.249 | 0.307 | 0.250 | 0.219 | 0.281 |
| ρ_Z | | 0.985 | 0.978 | 0.992 | 0.974 | 0.958 | 0.991 |
| ρ_{b^2} | | 0.765 | 0.658 | 0.871 | 0.852 | 0.769 | 0.934 |
| ρ_g | | 0.977 | 0.965 | 0.988 | 0.978 | 0.964 | 0.991 |
| ρ_μ | | 0.780 | 0.694 | 0.873 | 0.673 | 0.565 | 0.784 |
| ρ_{ms} | | 0.167 | 0.076 | 0.256 | 0.137 | 0.049 | 0.220 |
| ρ_p | | 0.975 | 0.957 | 0.996 | 0.979 | 0.961 | 0.997 |
| ρ_w | | 0.960 | 0.936 | 0.986 | 0.965 | 0.944 | 0.988 |
| θ_p | price MA | 0.882 | 0.823 | 0.944 | 0.872 | 0.805 | 0.941 |

Continued on next page

Table 3 – continued from previous page

| | | 1948Q1 - 2014Q2 | | | 1948Q1 - 2004Q4 | | |
|---------------|-------------------|-----------------|---------|---------|-----------------|---------|---------|
| | | mean | HDP inf | HPD sup | mean | HDP inf | HPD sup |
| θ_w | wage MA | 0.922 | 0.880 | 0.965 | 0.890 | 0.838 | 0.943 |
| $S''(\gamma)$ | inv. adj cost | 4.063 | 2.572 | 5.514 | 6.747 | 5.094 | 8.391 |
| σ_c | intertemp elast | 1.984 | 1.528 | 2.443 | 1.450 | 1.082 | 1.817 |
| h | habits | 0.384 | 0.287 | 0.483 | 0.506 | 0.421 | 0.585 |
| ζ_w | wage stickiness | 0.803 | 0.739 | 0.866 | 0.749 | 0.679 | 0.816 |
| ν_L | labor | 1.638 | 0.862 | 2.391 | 1.971 | 1.135 | 2.770 |
| ζ_p | price stickiness | 0.675 | 0.584 | 0.761 | 0.608 | 0.531 | 0.680 |
| ι_w | wage index | 0.489 | 0.319 | 0.656 | 0.523 | 0.359 | 0.682 |
| ι_p | price index | 0.290 | 0.162 | 0.415 | 0.311 | 0.168 | 0.451 |
| $czcap$ | capital utiliz. | 0.690 | 0.528 | 0.854 | 0.453 | 0.316 | 0.596 |
| Φ | fixed cost | 1.609 | 1.495 | 1.721 | 1.441 | 1.318 | 1.569 |
| ψ_1 | monetary p. | 1.969 | 1.705 | 2.226 | 1.982 | 1.691 | 2.266 |
| ρ_R | monetary p. | 0.881 | 0.853 | 0.909 | 0.884 | 0.858 | 0.910 |
| ψ_2 | monetary p. | 0.084 | 0.043 | 0.123 | 0.084 | 0.015 | 0.146 |
| ψ_3 | monetary p. | 0.187 | 0.152 | 0.221 | 0.223 | 0.178 | 0.269 |
| $const_\pi$ | measurem eq | 0.740 | 0.612 | 0.868 | 0.725 | 0.574 | 0.871 |
| $constbeta$ | measurem eq | 0.248 | 0.091 | 0.398 | 0.251 | 0.094 | 0.403 |
| $const_L$ | measurem eq | -0.021 | -2.097 | 2.128 | 0.558 | -1.719 | 2.750 |
| $ctrend$ | growth rate | 0.469 | 0.440 | 0.499 | 0.469 | 0.440 | 0.499 |
| ρ_{gZ} | gov't exp | 0.629 | 0.495 | 0.761 | 0.629 | 0.495 | 0.761 |
| α | $\frac{k_*}{y_*}$ | 0.192 | 0.129 | 0.252 | 0.192 | 0.129 | 0.252 |

Note: HPD inf and HPD low are respectively the lower bound and the upper bound of a 90% HPD interval.

Table 4: Posterior distributions in subsamples
 (Metropolis Hasting 25000 simulations)

| | 1964Q3 | 2004Q4 | |
|-------------------|--------|---------|---------|
| | mean | HPD inf | HPD sup |
| $\sigma(e_Z)$ | 0.4291 | 0.387 | 0.4763 |
| $\sigma(e_{b^2})$ | 0.2315 | 0.1885 | 0.2725 |
| $\sigma(e_g)$ | 0.5148 | 0.4663 | 0.5631 |
| $\sigma(e_\mu)$ | 0.4 | 0.3349 | 0.4581 |
| $\sigma(e_{ms})$ | 0.2407 | 0.2167 | 0.2626 |
| $\sigma(e_p)$ | 0.1255 | 0.1018 | 0.1489 |
| $\sigma(e_w)$ | 0.2615 | 0.2225 | 0.3019 |
| ρ_z | 0.9551 | 0.9358 | 0.972 |
| ρ_{b^2} | 0.2503 | 0.0985 | 0.387 |
| ρ_g | 0.9622 | 0.9473 | 0.9783 |
| ρ_μ | 0.8022 | 0.7075 | 0.8962 |
| ρ_{ms} | 0.1588 | 0.0541 | 0.2667 |
| ρ_p | 0.8913 | 0.8181 | 0.9628 |
| ρ_w | 0.945 | 0.912 | 0.9774 |
| θ_p | 0.7311 | 0.5947 | 0.8682 |
| θ_w | 0.798 | 0.6775 | 0.9217 |
| $S''(\gamma)$ | 5.1598 | 3.423 | 6.9416 |
| σ_c | 1.559 | 1.2911 | 1.8157 |
| h | 0.656 | 0.5741 | 0.7373 |
| ζ_w | 0.6916 | 0.5841 | 0.805 |
| ν_L | 1.6361 | 0.613 | 2.6136 |
| ζ_p | 0.7103 | 0.6449 | 0.7788 |

Continued on next page

Table 4 – continued from previous page

| | 1964Q3 | 2004Q4 | HPD inf | HPD sup |
|---------------|--------|---------|---------|---------|
| | mean | mode | | |
| ι_w | 0.6333 | 0.4523 | 0.8503 | |
| ι_p | 0.3351 | 0.1581 | 0.512 | |
| czcap | 0.4927 | 0.341 | 0.6416 | |
| Φ | 1.6933 | 1.5646 | 1.826 | |
| ψ_1 | 2.0242 | 1.7617 | 2.3184 | |
| ρ_R | 0.8107 | 0.7726 | 0.8538 | |
| ψ_2 | 0.0826 | 0.047 | 0.1218 | |
| ψ_3 | 0.2203 | 0.1752 | 0.2636 | |
| $const_{\pi}$ | 0.7029 | 0.5892 | 0.8139 | |
| constebeta | 0.2492 | 0.0954 | 0.3969 | |
| $const_L$ | 0.4429 | -1.2289 | 2.0248 | |
| ctrend | 0.4388 | 0.407 | 0.4737 | |
| ρ_{gZ} | 0.6125 | 0.4584 | 0.7703 | |
| α | 0.3058 | 0.2376 | 0.3662 | |

Note: HPD inf and HPD low are respectively the lower bound and the upper bound of a 90% HPD interval.

Table 5: Posterior distributions in subsamples
 (Metropolis Hasting 25000 simulations)

| | 1964Q3 | 2014Q2 | |
|-------------------|--------|---------|---------|
| | mean | HPD inf | HPD sup |
| $\sigma(e_Z)$ | 0.4503 | 0.4058 | 0.4917 |
| $\sigma(e_{b^2})$ | 0.1695 | 0.1187 | 0.2141 |
| $\sigma(e_g)$ | 0.5016 | 0.4628 | 0.541 |
| $\sigma(e_\mu)$ | 0.3922 | 0.3191 | 0.4585 |
| $\sigma(e_{ms})$ | 0.2307 | 0.2063 | 0.2518 |
| $\sigma(e_p)$ | 0.1244 | 0.1054 | 0.1459 |
| $\sigma(e_w)$ | 0.3707 | 0.3304 | 0.4085 |
| ρ_z | 0.9683 | 0.9538 | 0.9826 |
| ρ_{b^2} | 0.5116 | 0.3251 | 0.7038 |
| ρ_g | 0.9687 | 0.957 | 0.9811 |
| ρ_μ | 0.8581 | 0.7915 | 0.9304 |
| ρ_{ms} | 0.1936 | 0.0922 | 0.3074 |
| ρ_p | 0.9166 | 0.8642 | 0.9723 |
| ρ_w | 0.9563 | 0.9251 | 0.986 |
| θ_p | 0.7819 | 0.6806 | 0.8804 |
| θ_w | 0.9239 | 0.8741 | 0.9712 |
| $S''(\gamma)$ | 4.156 | 2.5194 | 5.7709 |
| σ_c | 1.6607 | 1.3621 | 1.9754 |
| h | 0.5599 | 0.4544 | 0.6695 |
| ζ_w | 0.7859 | 0.7022 | 0.8657 |
| ν_L | 1.3969 | 0.5682 | 2.2562 |
| ζ_p | 0.766 | 0.7085 | 0.8258 |

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Table 5 – continued from previous page

| | 1964Q3 2014Q2 | mean | HPD inf | HPD sup |
|--------------------|---------------|---------|---------|---------|
| ι_w | | 0.6317 | 0.4331 | 0.8272 |
| ι_p | | 0.2461 | 0.1189 | 0.3704 |
| czcap | | 0.8195 | 0.7152 | 0.9267 |
| Φ | | 1.6489 | 1.5227 | 1.7826 |
| ψ_1 | | 1.8612 | 1.5969 | 2.1239 |
| ρ_R | | 0.805 | 0.766 | 0.8486 |
| ψ_2 | | 0.0589 | 0.0334 | 0.0879 |
| ψ_3 | | 0.2311 | 0.1872 | 0.2747 |
| $const_{\pi}$ | | 0.735 | 0.6263 | 0.8319 |
| constebeta | | 0.2659 | 0.0865 | 0.4361 |
| $const_L$ | | -1.3316 | -3.0709 | 0.442 |
| ctrend 0.4224 | | 0.3942 | 0.4508 | |
| ρ_{gZ} 0.5644 | | 0.4437 | 0.7014 | |
| α | | 0.2043 | 0.1498 | 0.2613 |

Note: HPD inf and HPD low are respectively the lower bound and the upper bound of a 90% HPD interval.

Table 6: Posterior distributions in subsamples
 (Metropolis Hasting 25000 simulations)

| | 1988Q3 | 1996Q3 | |
|-------------------|--------|---------|---------|
| | mean | HPD inf | HPD sup |
| $\sigma(e_Z)$ | 0.2882 | 0.219 | 0.3528 |
| $\sigma(e_{b^2})$ | 0.1246 | 0.067 | 0.1865 |
| $\sigma(e_g)$ | 0.3155 | 0.25 | 0.3739 |
| $\sigma(e_\mu)$ | 0.4558 | 0.3271 | 0.5832 |
| $\sigma(e_{ms})$ | 0.0925 | 0.0708 | 0.1124 |
| $\sigma(e_p)$ | 0.083 | 0.0516 | 0.1129 |
| $\sigma(e_w)$ | 0.2204 | 0.1461 | 0.2872 |
| ρ_Z | 0.7458 | 0.5487 | 0.9732 |
| ρ_{b^2} | 0.5977 | 0.2279 | 0.8847 |
| ρ_g | 0.9267 | 0.8719 | 0.9817 |
| ρ_μ | 0.3911 | 0.1262 | 0.6495 |
| ρ_{ms} | 0.4781 | 0.2682 | 0.6737 |
| ρ_p | 0.5854 | 0.3498 | 0.8421 |
| ρ_w | 0.7952 | 0.6795 | 0.9101 |
| θ_p | 0.4058 | 0.1308 | 0.6579 |
| θ_w | 0.301 | 0.0822 | 0.5021 |
| $S''(\gamma)$ | 4.7921 | 2.7207 | 6.6622 |
| σ_c | 1.072 | 0.7006 | 1.4933 |
| h | 0.656 | 0.5142 | 0.79 |
| ζ_w | 0.5462 | 0.422 | 0.6718 |
| ν_L | 1.8372 | 0.7643 | 2.9239 |
| ζ_p | 0.7658 | 0.6753 | 0.8646 |

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Table 6 – continued from previous page

| | 1988Q3 | 1996Q3 | |
|---------------|--------|---------|---------|
| | mean | HPD inf | HPD sup |
| ι_w | 0.5012 | 0.2545 | 0.7297 |
| ι_p | 0.4688 | 0.2266 | 0.7011 |
| czcap | 0.4872 | 0.2836 | 0.6949 |
| Φ | 1.3721 | 1.2155 | 1.5335 |
| ψ_1 | 1.671 | 1.3066 | 2.0548 |
| ρ_R | 0.8119 | 0.7427 | 0.8883 |
| ψ_2 | 0.1129 | 0.0242 | 0.1984 |
| ψ_3 | 0.1085 | 0.0511 | 0.1635 |
| $const_{\pi}$ | 0.5856 | 0.4756 | 0.6945 |
| constbeta | 0.2589 | 0.1059 | 0.4173 |
| $const_L$ | 0.9035 | -0.4823 | 2.2403 |
| ctrend | 0.4496 | 0.3738 | 0.517 |
| ρ_{gZ} | 0.3218 | 0.0555 | 0.5523 |
| α | 0.2974 | 0.2228 | 0.3678 |

Note: HPD inf and HPD low are respectively the lower bound and the upper bound of a 90% HPD interval.

Table 7: Posterior distributions in subsamples
 (Metropolis Hasting 25000 simulations)

| | 1996Q4 | 2004Q4 | |
|-------------------|--------|---------|---------|
| | mean | HPD inf | HPD sup |
| $\sigma(e_Z)$ | 0.4079 | 0.3126 | 0.4967 |
| $\sigma(e_{b^2})$ | 0.1256 | 0.0638 | 0.1938 |
| $\sigma(e_g)$ | 0.41 | 0.3038 | 0.5113 |
| $\sigma(e_\mu)$ | 0.3707 | 0.234 | 0.5004 |
| $\sigma(e_{ms})$ | 0.1022 | 0.0781 | 0.1256 |
| $\sigma(e_p)$ | 0.0804 | 0.0485 | 0.1073 |
| $\sigma(e_w)$ | 0.5068 | 0.3536 | 0.6731 |
| ρ_Z | 0.6237 | 0.3747 | 0.8924 |
| ρ_{b^2} | 0.5997 | 0.2702 | 0.8714 |
| ρ_g | 0.8407 | 0.742 | 0.9358 |
| ρ_μ | 0.6198 | 0.4028 | 0.8204 |
| ρ_{ms} | 0.3761 | 0.136 | 0.6149 |
| ρ_p | 0.7787 | 0.5951 | 0.9646 |
| ρ_w | 0.4959 | 0.2056 | 0.7961 |
| θ_p | 0.4963 | 0.1895 | 0.7981 |
| θ_w | 0.5532 | 0.2601 | 0.8866 |
| $S''(\gamma)$ | 5.5894 | 3.4611 | 7.4793 |
| σ_c | 1.3417 | 0.907 | 1.7607 |
| h | 0.6367 | 0.5228 | 0.7628 |
| ζ_w | 0.6473 | 0.5315 | 0.7749 |
| ν_L | 1.3849 | 0.3243 | 2.2253 |
| ζ_p | 0.688 | 0.5803 | 0.798 |

Continued on next page

Table 7 – continued from previous page

| | 1996Q4 | 2004Q4 | |
|-------------|--------|---------|---------|
| | mean | HPD inf | HPD sup |
| ι_w | 0.4919 | 0.2539 | 0.7359 |
| ι_p | 0.4385 | 0.2116 | 0.6862 |
| czcap | 0.6921 | 0.5116 | 0.879 |
| Φ | 1.3714 | 1.2141 | 1.5314 |
| ψ_1 | 1.5669 | 1.1107 | 1.9673 |
| ρ_R | 0.8162 | 0.7387 | 0.8881 |
| ψ_2 | 0.1493 | 0.0795 | 0.2259 |
| ψ_3 | 0.1259 | 0.0677 | 0.185 |
| $const_\pi$ | 0.6719 | 0.5296 | 0.8102 |
| constebeta | 0.2519 | 0.0841 | 0.4186 |
| $const_L$ | 1.2548 | -0.2489 | 2.6742 |
| ctrend | 0.6078 | 0.5391 | 0.678 |
| ρ_{gZ} | 0.7038 | 0.4368 | 0.9379 |
| α | 0.2529 | 0.1685 | 0.3335 |

Note: HPD inf and HPD low are respectively the lower bound and the upper bound of a 90% HPD interval.

Table 8: Posterior distributions in subsamples
 (Metropolis Hasting 25000 simulations)

| | 2005Q1 | 2014Q2 | |
|-------------------|--------|---------|---------|
| | mean | HPD inf | HPD sup |
| $\sigma(e_Z)$ | 0.5226 | 0.4086 | 0.6298 |
| $\sigma(e_{b^2})$ | 0.0869 | 0.055 | 0.12 |
| $\sigma(e_g)$ | 0.4499 | 0.3565 | 0.5381 |
| $\sigma(e_\mu)$ | 0.344 | 0.1989 | 0.4766 |
| $\sigma(e_{ms})$ | 0.0851 | 0.0649 | 0.1051 |
| $\sigma(e_p)$ | 0.1483 | 0.0926 | 0.1979 |
| $\sigma(e_w)$ | 0.7066 | 0.5553 | 0.8587 |
| ρ_Z | 0.9266 | 0.8847 | 0.9734 |
| ρ_{b^2} | 0.8576 | 0.7896 | 0.9289 |
| ρ_g | 0.8718 | 0.8052 | 0.9391 |
| ρ_μ | 0.7338 | 0.5287 | 0.9479 |
| ρ_{ms} | 0.6126 | 0.4404 | 0.786 |
| ρ_p | 0.5598 | 0.248 | 0.8725 |
| ρ_w | 0.3815 | 0.0676 | 0.6421 |
| θ_p | 0.5462 | 0.2787 | 0.8617 |
| θ_w | 0.5465 | 0.3404 | 0.7756 |
| $S''(\gamma)$ | 4.9614 | 2.8923 | 6.8457 |
| σ_c | 1.0194 | 0.7983 | 1.2607 |
| h | 0.7195 | 0.6236 | 0.8228 |
| ζ_w | 0.7329 | 0.6073 | 0.8664 |
| ν_L | 1.14 | 0.2503 | 2.1553 |
| ζ_p | 0.7789 | 0.6841 | 0.8626 |

Continued on next page

Table 8 – continued from previous page

| | 2005Q1 | 2014Q2 | |
|---------------|--------|---------|---------|
| | mean | HPD inf | HPD sup |
| ι_w | 0.4249 | 0.1939 | 0.6651 |
| ι_p | 0.321 | 0.1365 | 0.4771 |
| czcap | 0.8369 | 0.7358 | 0.94 |
| Φ | 1.347 | 1.1744 | 1.5085 |
| ψ_1 | 1.4027 | 1 | 1.7051 |
| ρ_R | 0.8623 | 0.8036 | 0.91 |
| ψ_2 | 0.1086 | 0.0493 | 0.1657 |
| ψ_3 | 0.1135 | 0.0693 | 0.1569 |
| $const_{\pi}$ | 0.641 | 0.5075 | 0.7785 |
| constbeta | 0.2469 | 0.0897 | 0.386 |
| $const_L$ | -0.35 | -2.2096 | 1.6797 |
| ctrend | 0.3863 | 0.2745 | 0.4882 |
| ρ_{gZ} | 0.5888 | 0.3347 | 0.8105 |
| α | 0.2469 | 0.1732 | 0.3271 |

Note: HPD inf and HPD low are respectively the lower bound and the upper bound of a 90% HPD interval.

Table 9: Posterior distributions with Dixit-Stiglitz aggregator (Metropolis Hasting 25000 simulations)

| | DixitStiglitz aggregator 1948Q1 2014Q2 | | |
|-------------------|--|---------|---------|
| | mean | HPD inf | HPD sup |
| $\sigma(e_Z)$ | 0.5107 | 0.467 | 0.5557 |
| $\sigma(e_{b^2})$ | 0.1196 | 0.0947 | 0.1455 |
| $\sigma(e_g)$ | 0.6458 | 0.6007 | 0.6904 |
| $\sigma(e_\mu)$ | 0.4915 | 0.4281 | 0.5615 |
| $\sigma(e_{ms})$ | 0.2254 | 0.2034 | 0.2467 |
| $\sigma(e_p)$ | 0.21 | 0.1826 | 0.2371 |
| $\sigma(e_w)$ | 0.3495 | 0.3141 | 0.381 |
| ρ_Z | 0.985 | 0.9772 | 0.9922 |
| ρ_{b^2} | 0.8084 | 0.7436 | 0.878 |
| ρ_g | 0.9779 | 0.968 | 0.9884 |
| ρ_μ | 0.8168 | 0.7266 | 0.9193 |
| ρ_{ms} | 0.1602 | 0.0753 | 0.2439 |
| ρ_p | 0.9802 | 0.9651 | 0.9955 |
| ρ_w | 0.966 | 0.9439 | 0.9885 |
| θ_p | 0.8474 | 0.7787 | 0.917 |
| θ_w | 0.925 | 0.8811 | 0.9671 |
| $S''(\gamma)$ | 3.4535 | 2.001 | 4.668 |
| σ_c | 2.0222 | 1.5757 | 2.4105 |
| h | 0.3417 | 0.2519 | 0.4256 |
| ζ_w | 0.8866 | 0.8488 | 0.9211 |
| ν_L | 1.3977 | 0.6634 | 2.1122 |

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Table 9 – continued from previous page

DixitStiglitz aggregator 1948Q1 2014Q2

| | mean | HPD inf | HPD sup |
|-------------|---------|---------|---------|
| ζ_p | 0.8143 | 0.7659 | 0.8614 |
| ι_w | 0.4954 | 0.3109 | 0.6757 |
| ι_p | 0.3426 | 0.1936 | 0.4948 |
| czcap | 0.7796 | 0.6425 | 0.9161 |
| Φ | 1.591 | 1.4855 | 1.7084 |
| ψ_1 | 2.0434 | 1.7827 | 2.2888 |
| ρ_R | 0.8698 | 0.8422 | 0.898 |
| ψ_2 | 0.0732 | 0.0392 | 0.1112 |
| ψ_3 | 0.1852 | 0.1533 | 0.2216 |
| $const_\pi$ | 0.7639 | 0.6401 | 0.8935 |
| constebeta | 0.2482 | 0.0836 | 0.4018 |
| $const_L$ | -0.8031 | -2.9312 | 1.4902 |
| ctrend | 0.458 | 0.4288 | 0.487 |
| ρ_{gZ} | 0.627 | 0.4899 | 0.7577 |
| α | 0.1774 | 0.1166 | 0.2373 |

Note: HPD inf and HPD low are respectively the lower bound and the upper bound of a 90% HPD interval.

We report in Figure 8 the time series of the estimated shocks. The price and markup shocks do not appear particularly large, whereas it seems that during the financial crisis exceptionally large realizations of the shocks to the risk premium, TFP and monetary policy occurred.

F Figures

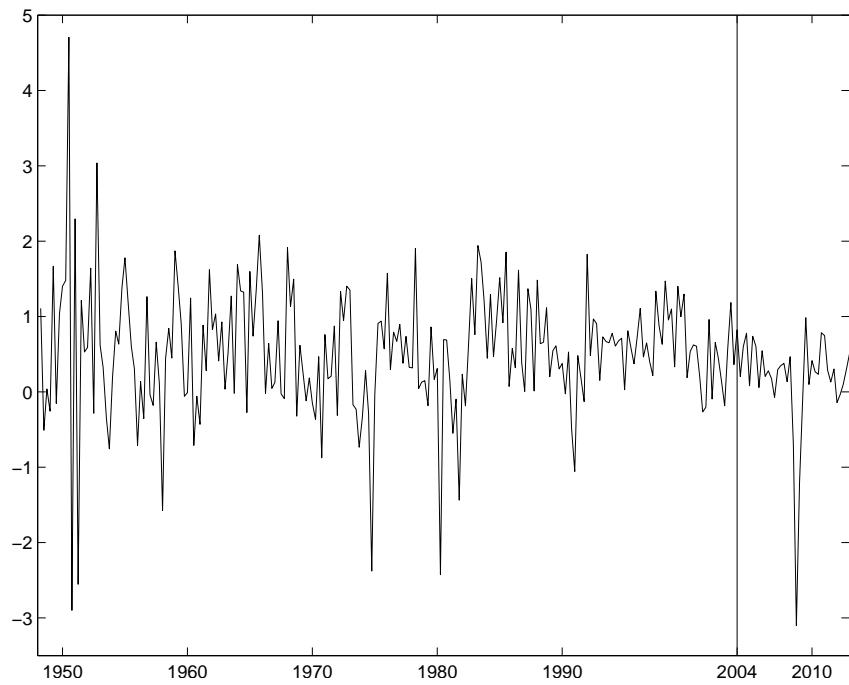


Figure 1: Consumption (percentage change). The line highlights the data points in addition to Smets, Wouters (2006)

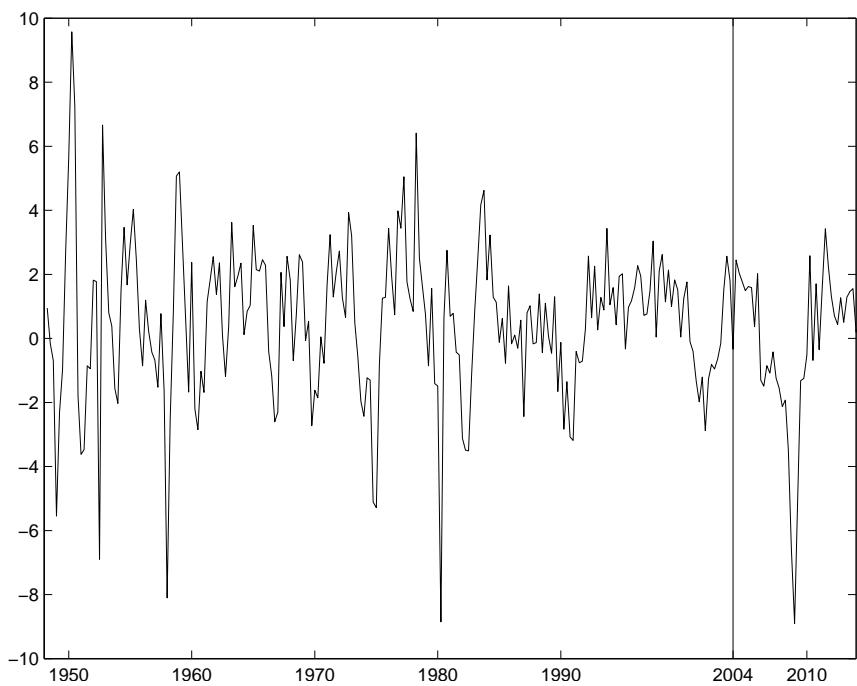


Figure 2: Investment (percentage change). The line highlights the data points in addition to Smets, Wouters (2006)

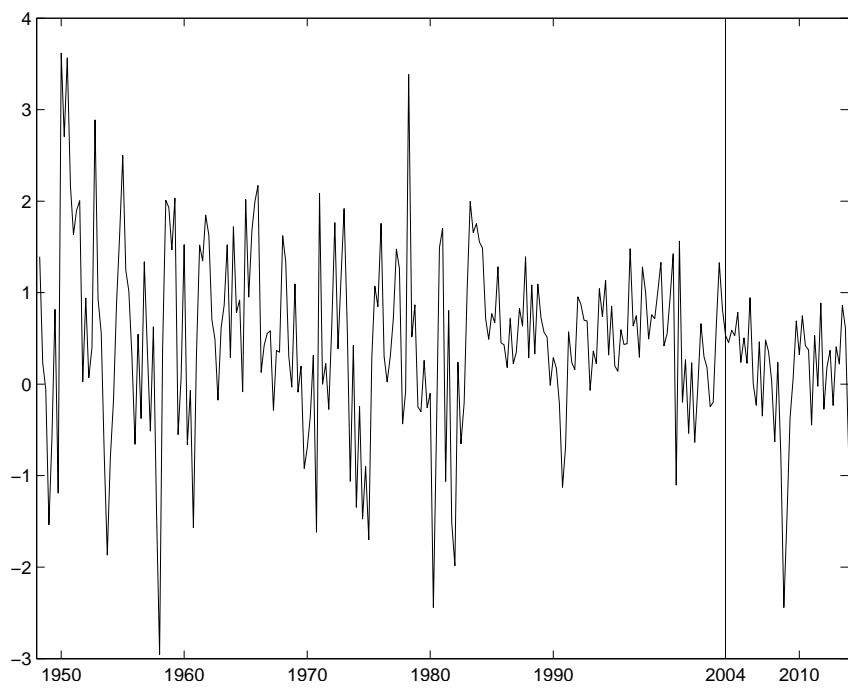


Figure 3: Output (percentage change). The line highlights the data points in addition to Smets, Wouters (2006)

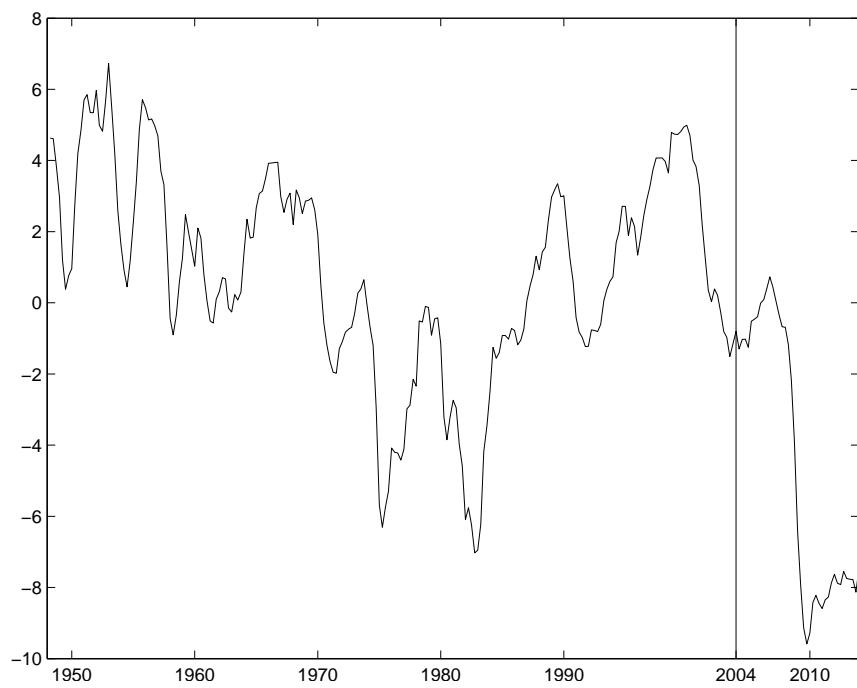


Figure 4: Labor (index). The line highlights the data points in addition to Smets, Wouters (2006)

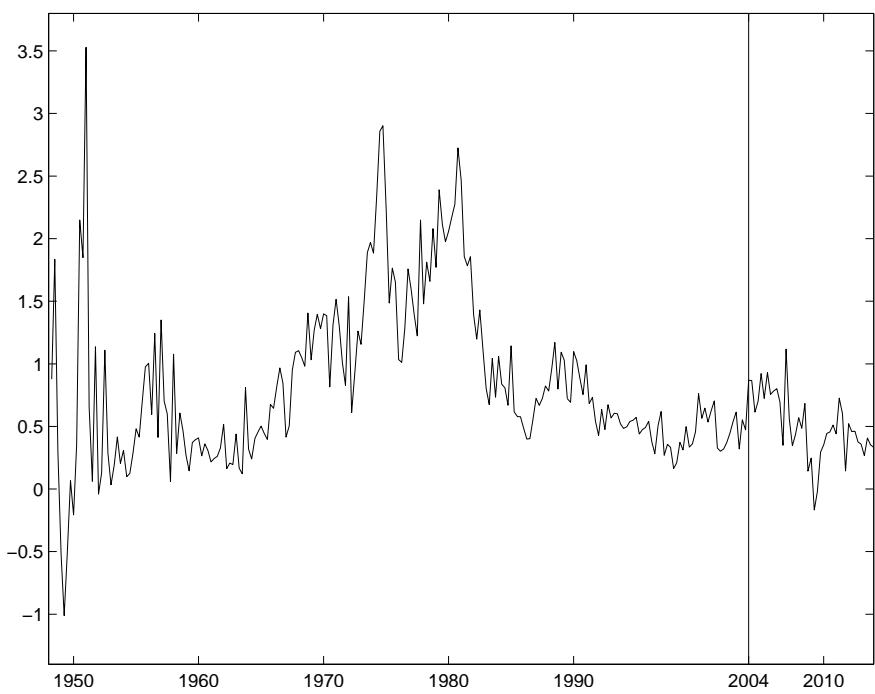


Figure 5: Inflation. The line highlights the data points in addition to Smets, Wouters (2006)

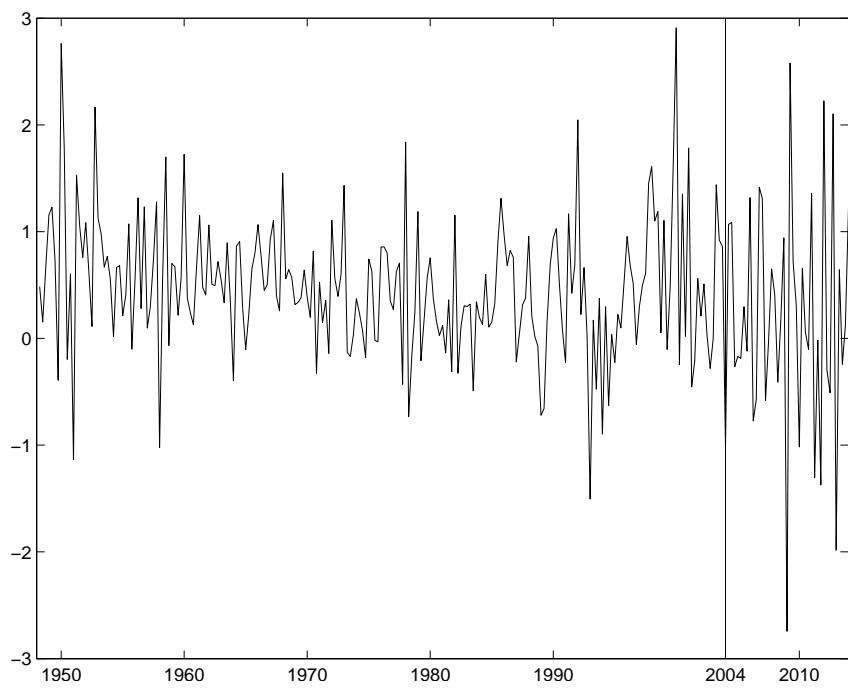


Figure 6: Wage rate (percentage change). The line highlights the data points in addition to Smets, Wouters (2006)

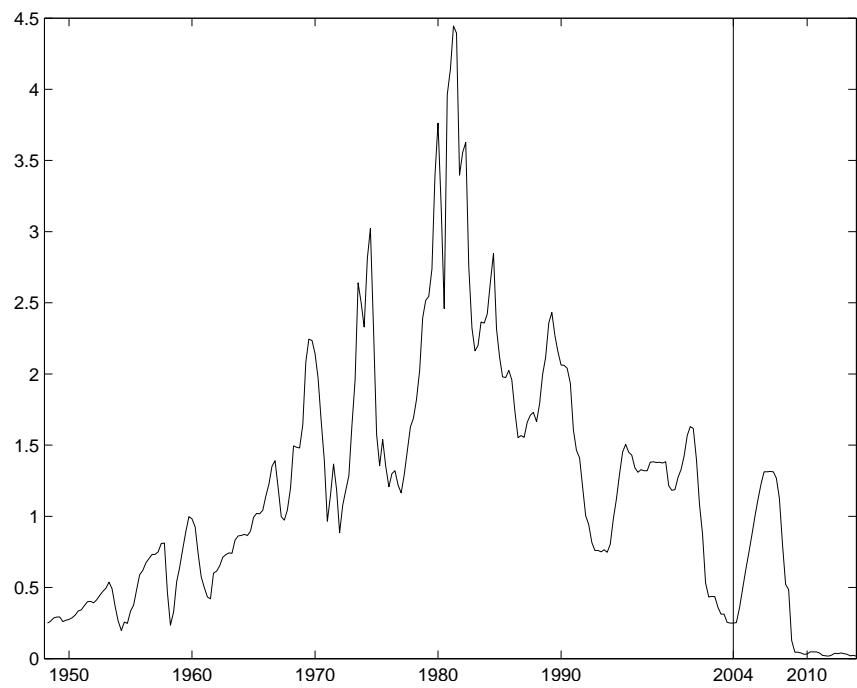


Figure 7: Interest rate. The line highlights the data points in addition to Smets, Wouters (2006)

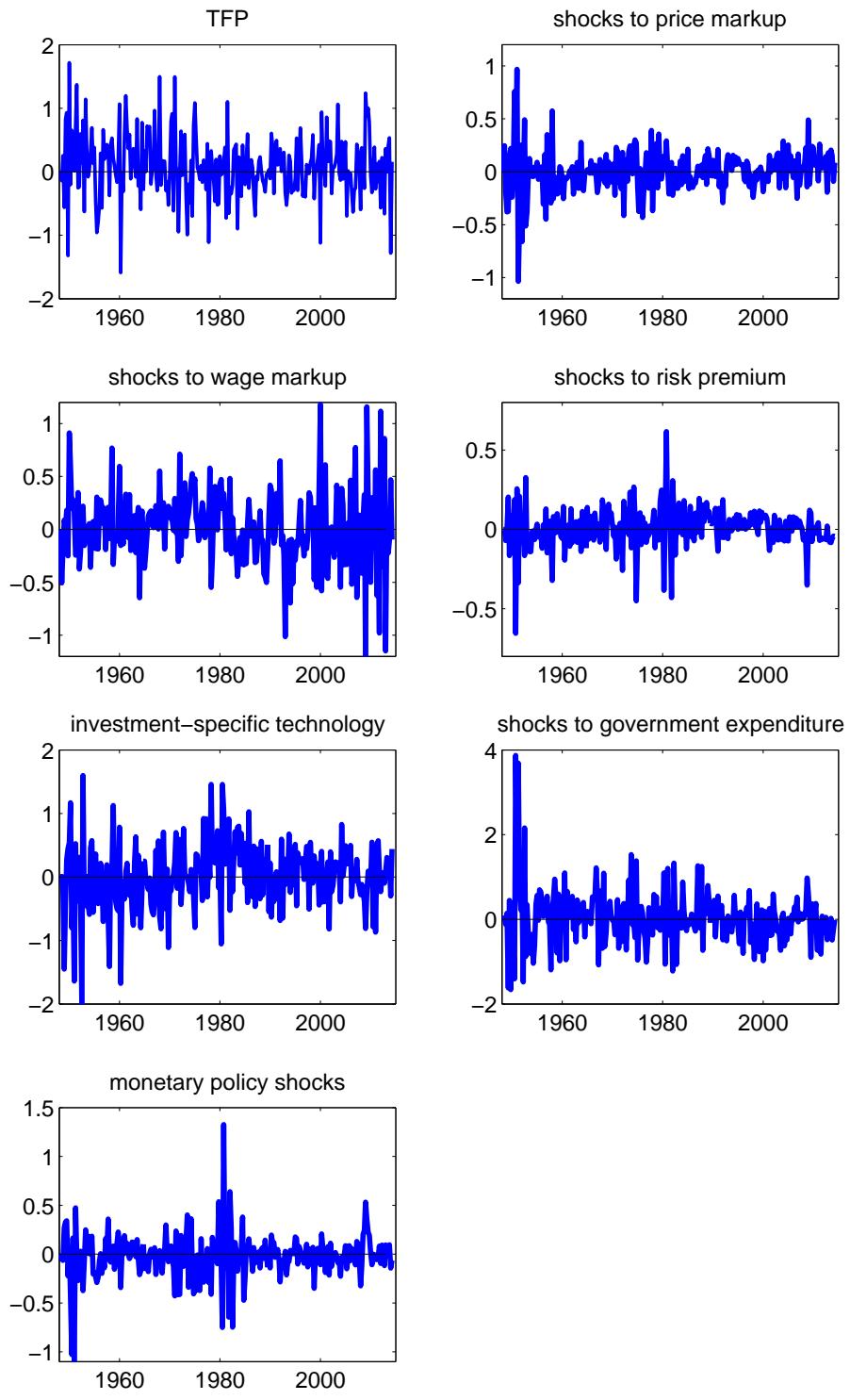


Figure 8: Time series of the estimated shocks

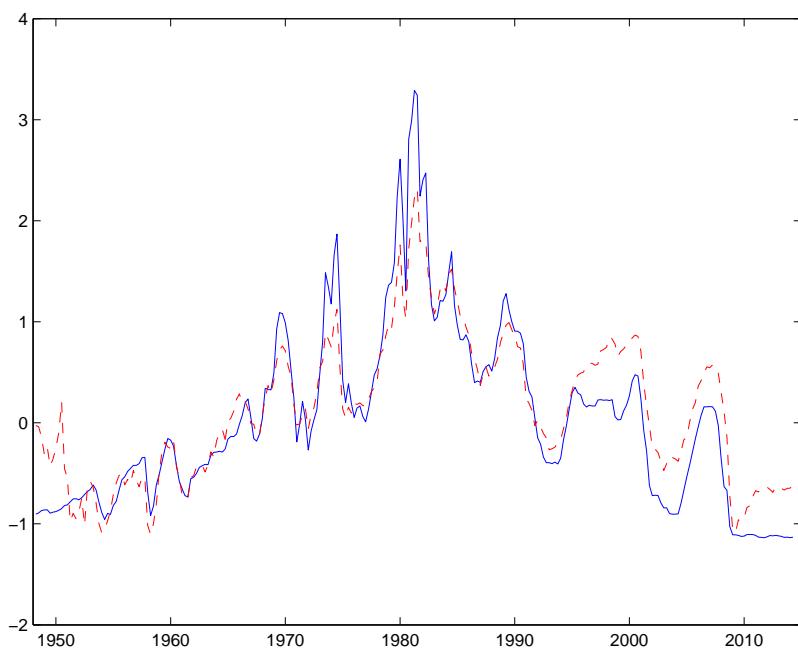


Figure 9: Historical shock decomposition of the interest rate (relative to long-run constant). The solid line is actual interest rate. The dashed line is interest rate generated by a combination of the estimated shocks to the risk premium and to investment adjustment cost.

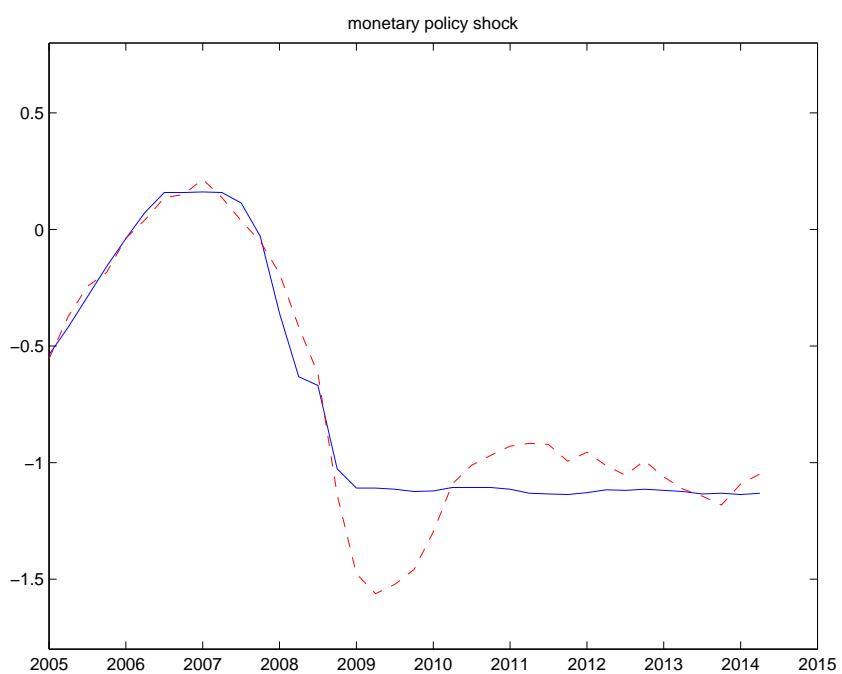


Figure 10: Historical shock decomposition of the interest rate (relative to long-run constant). The solid line is actual interest rate. The dashed line is the predicted interest rate if there were no monetary policy shocks.

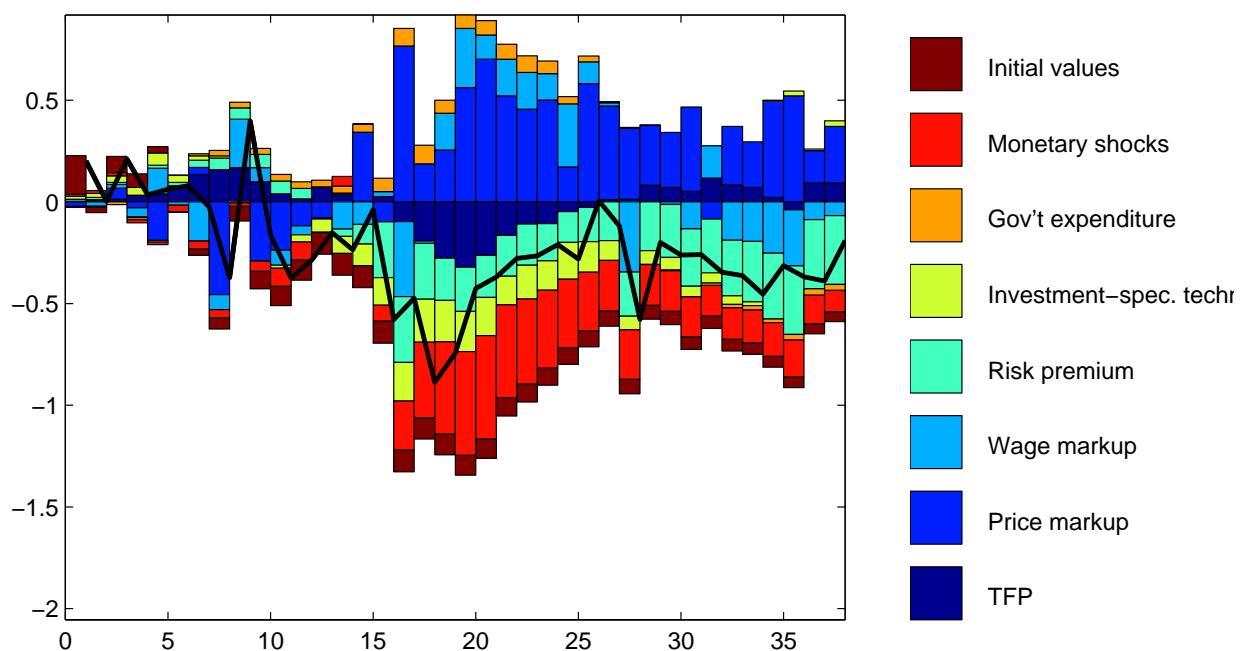


Figure 11: Historical shock decomposition of inflation

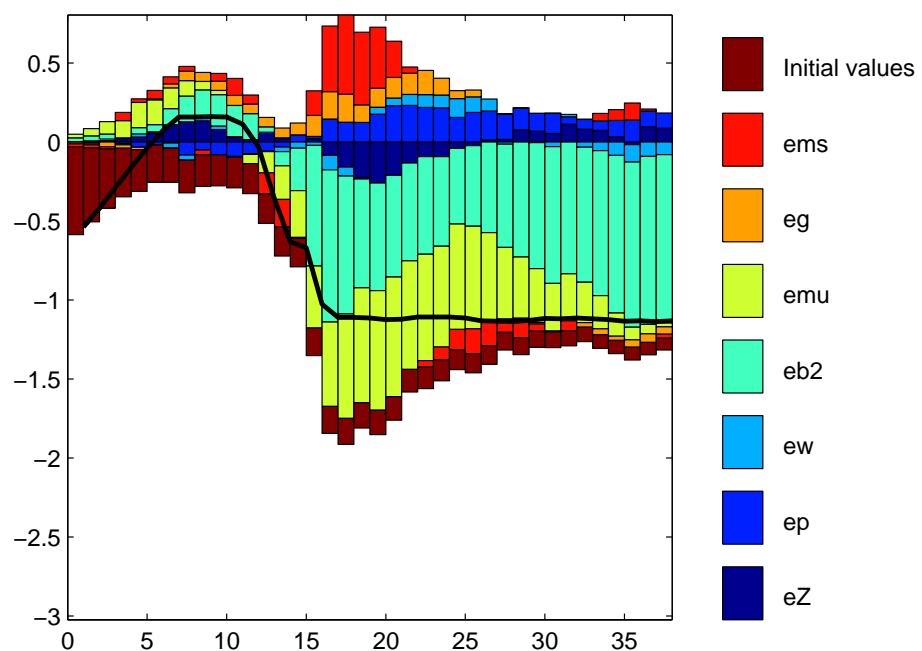


Figure 12: Historical shock decomposition of the interest rate

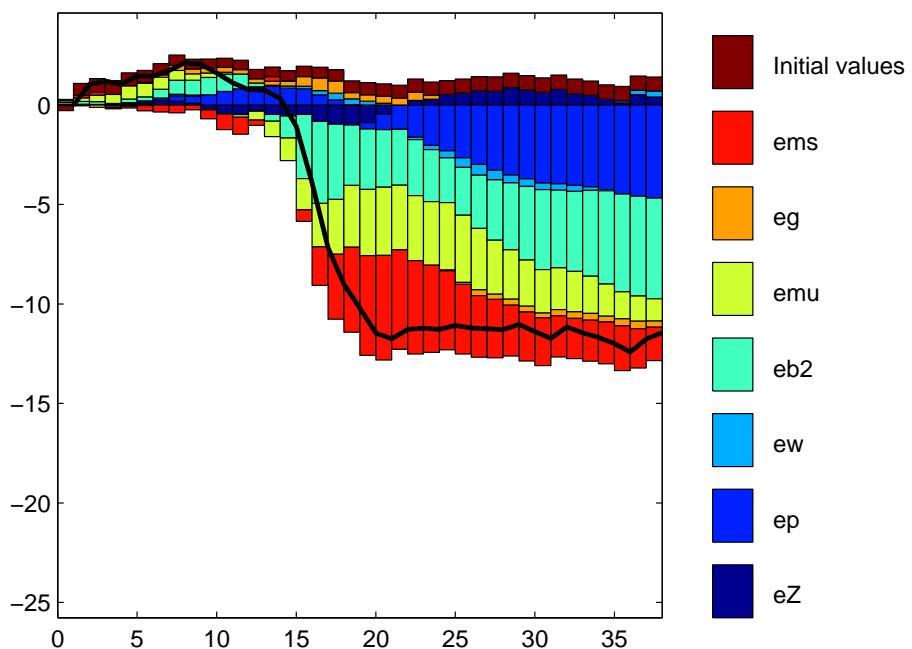


Figure 13: Historical shock decomposition of the output gap

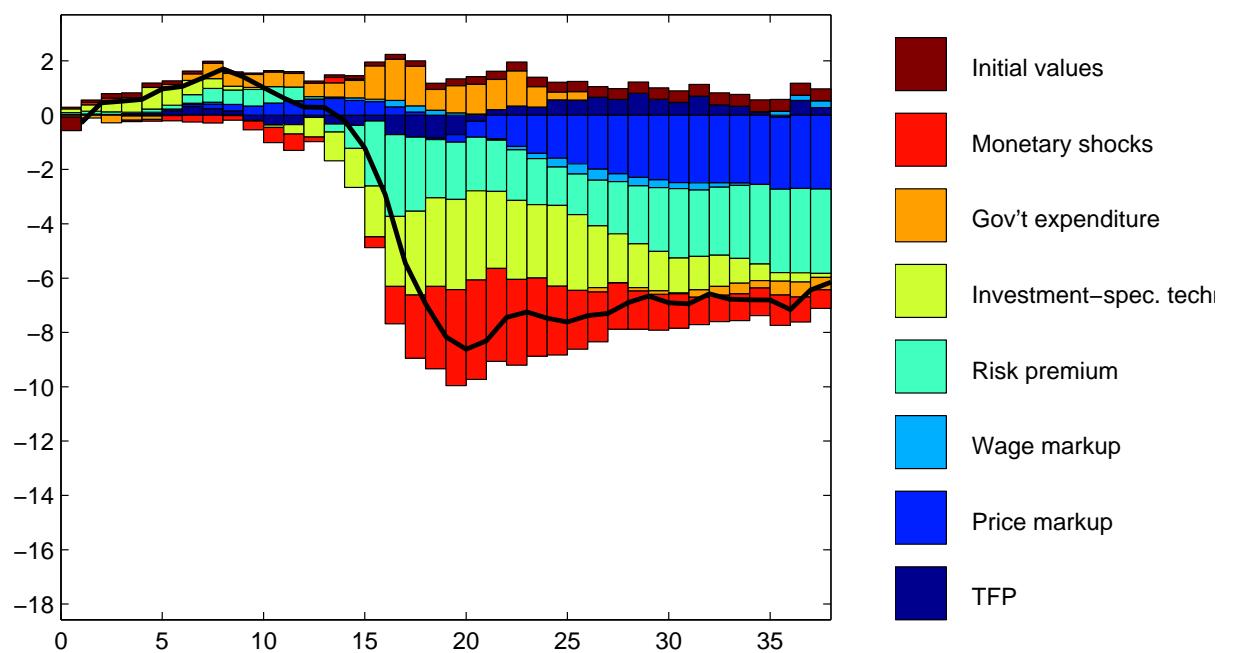


Figure 14: Historical shock decomposition of the labor

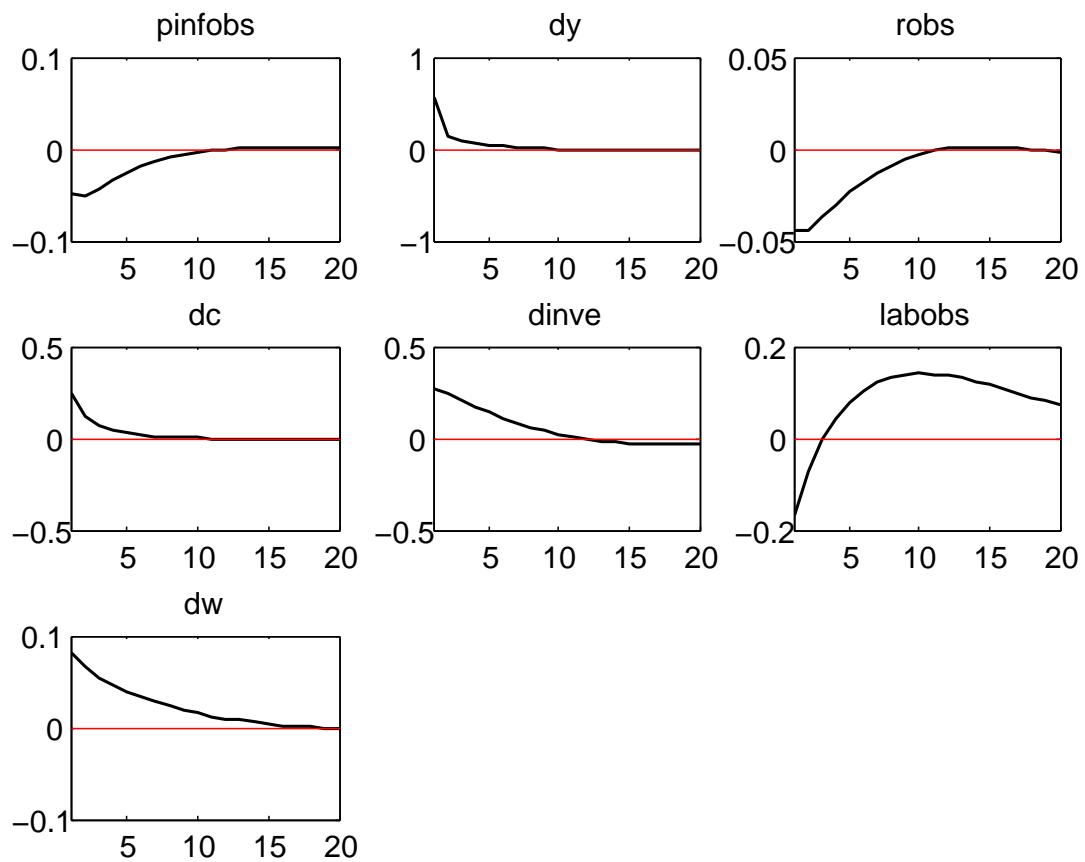


Figure 15: Impulse Response Functions to a shock to TFP

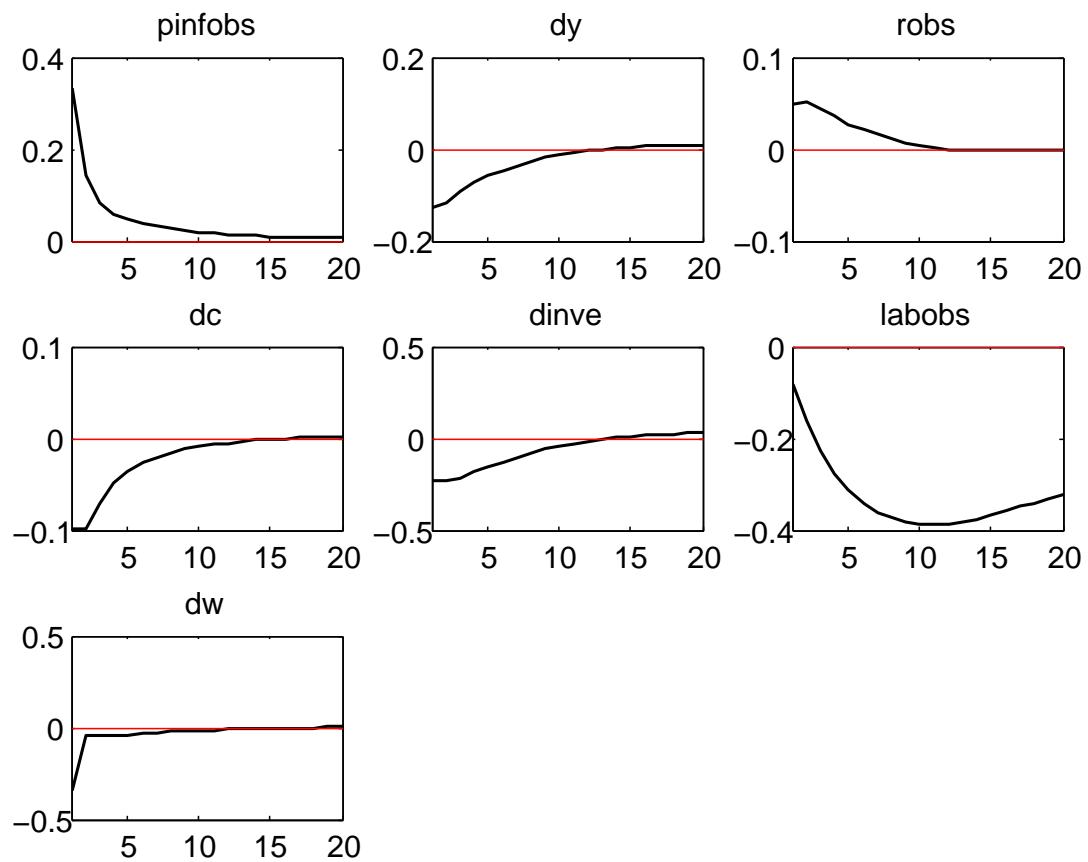


Figure 16: Impulse Response Functions to a shock to price markup

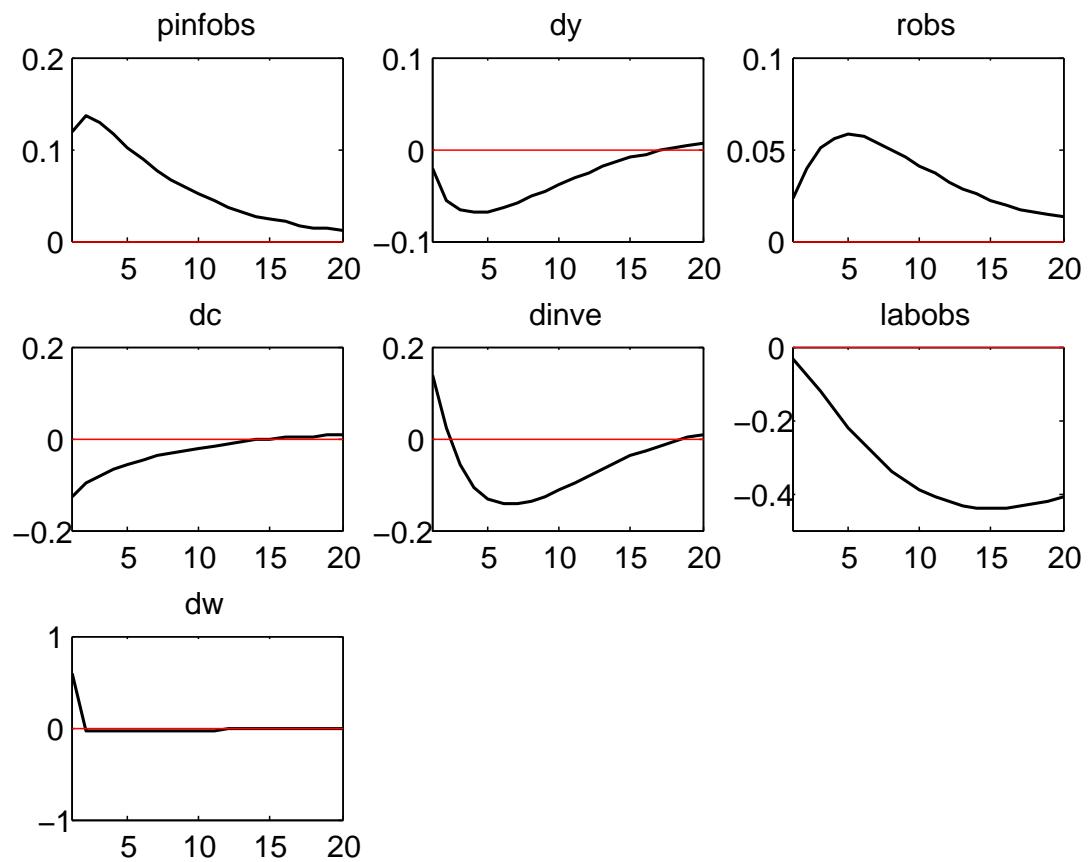


Figure 17: Impulse Response Functions to a shock to wage markup

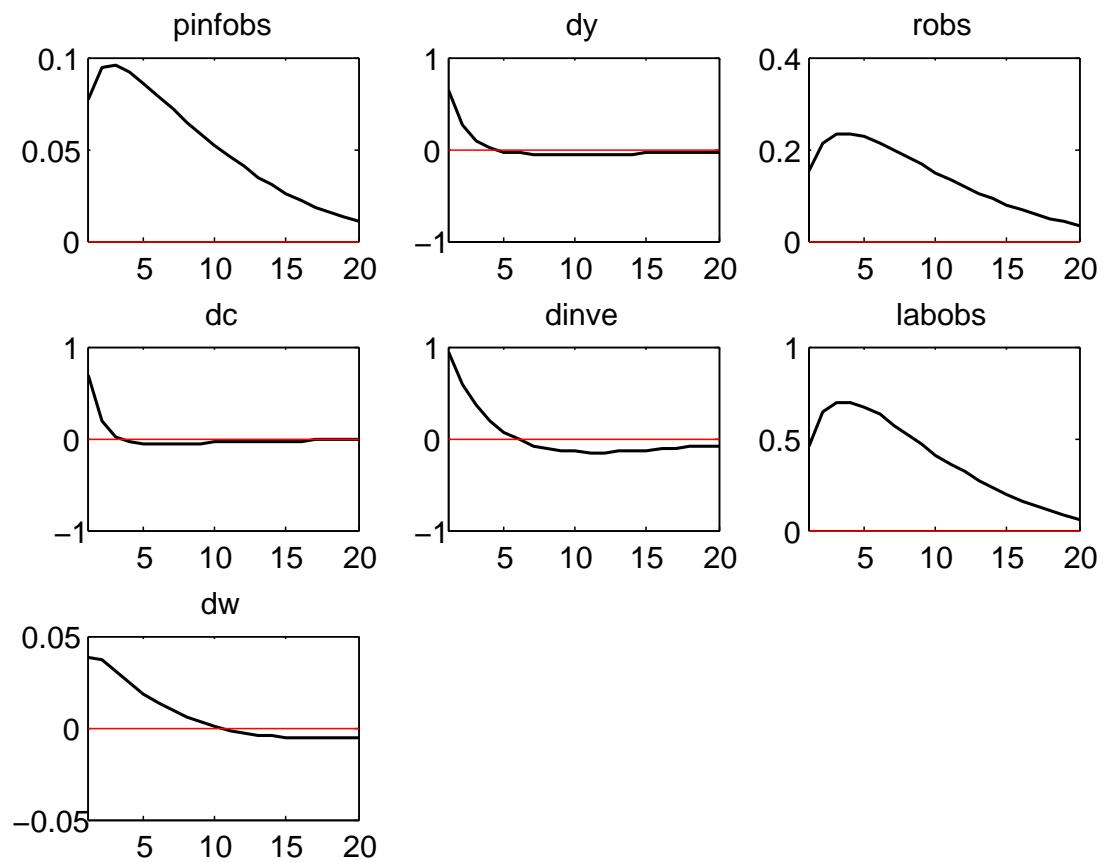


Figure 18: Impulse Response Functions to a shock to the risk premium

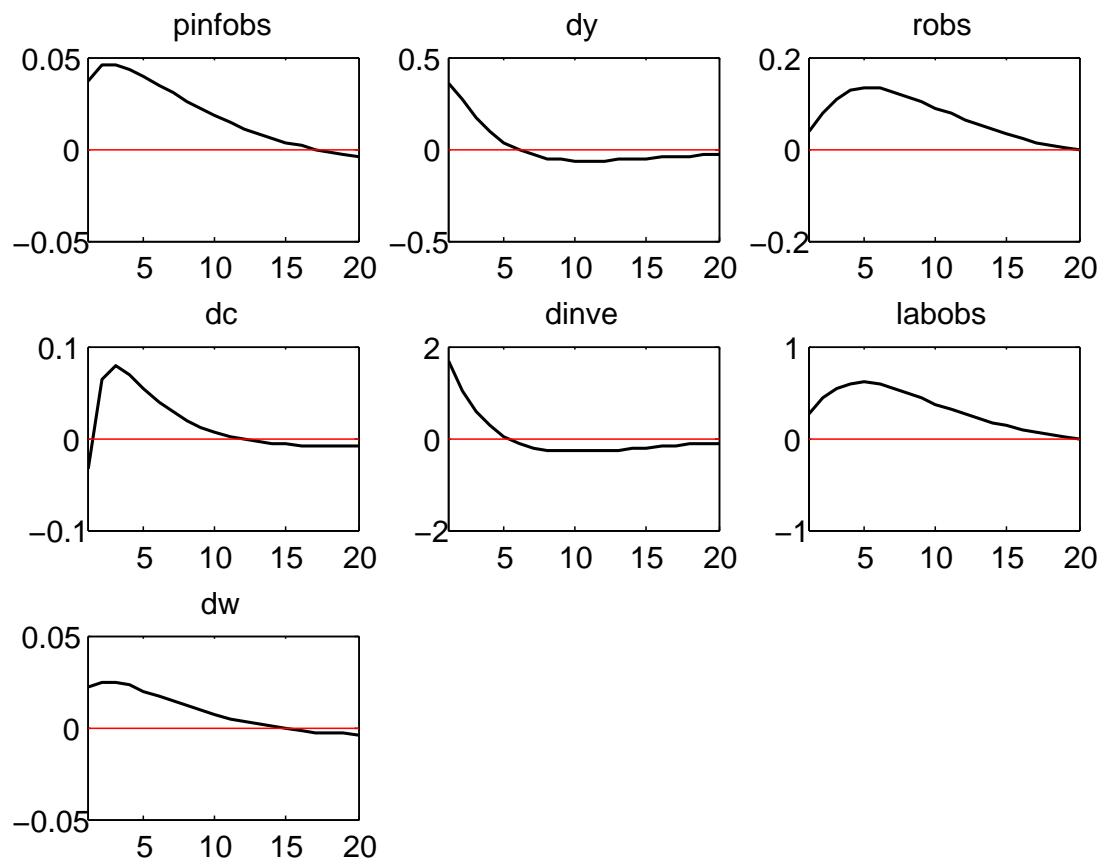


Figure 19: Impulse Response Functions to a shock to investment-specific technology

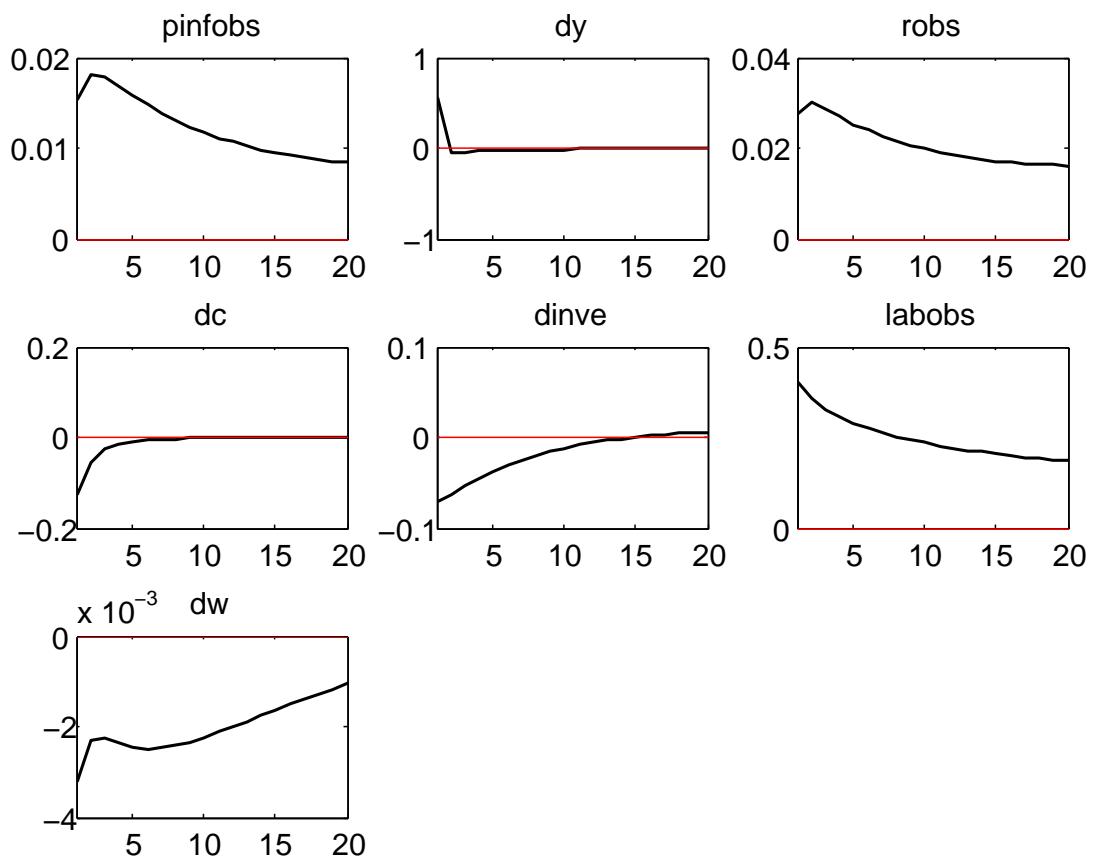


Figure 20: Impulse Response Functions to a shock to government expenditure

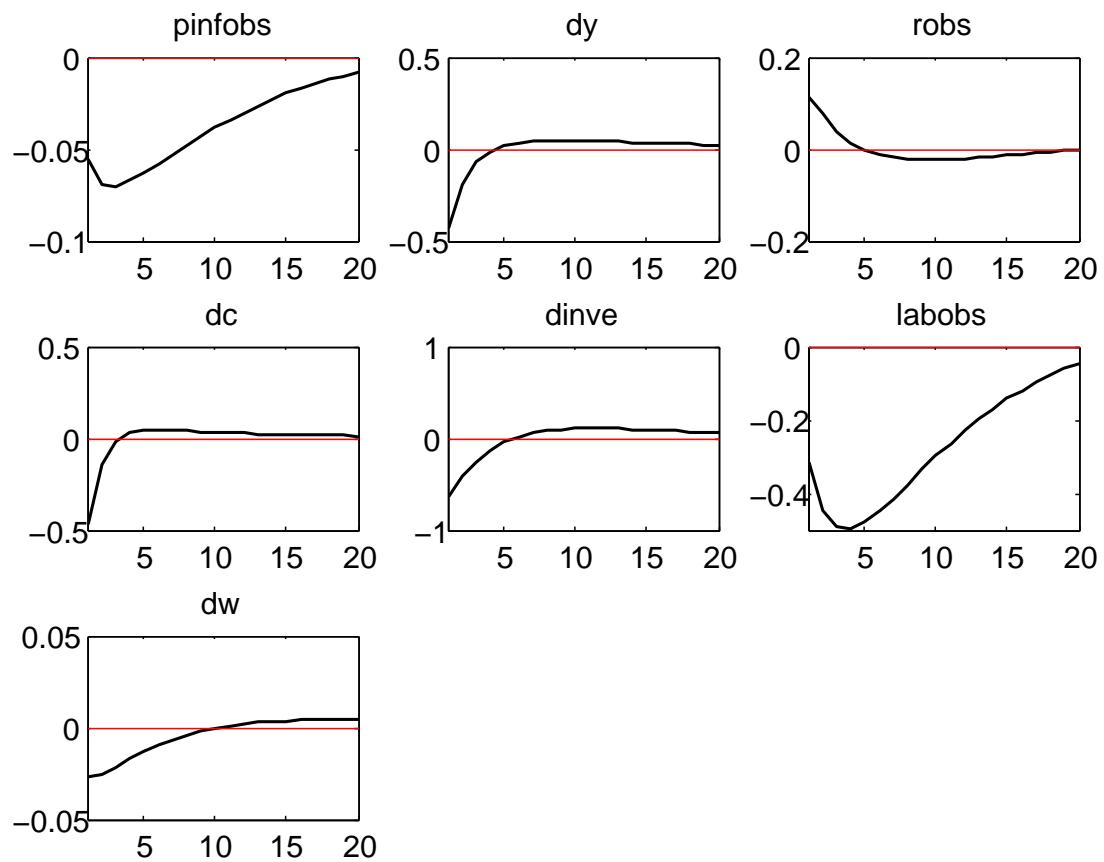


Figure 21: Impulse Response Functions to a monetary policy shock