## ONLINE APPENDIX

## A Additional Data Source and Methods Tables and Figures

In Section 2 we discussed the construction of our eligible-workers frame. Here we provide further details on which workers are excluded from the all-workers frame to arrive at the eligible-workers frame and how this impacts the earnings coverage of LEHD when compared to NIPA.

## A. 1 All-Workers Frame

The all-workers frame contains earnings for all jobs reported on the UI data for each date regime in the relevant years from 1990-2013, as noted in Figure 1 and summarized in Table A. 1 below.

Using the person level earnings, $e_{i t}$, an estimate of annual earnings for the all-workers frame in year $t$ is calculated as follows:

$$
E_{t}^{A W}=\sum_{i \in A W_{t}} e_{i t}
$$

where $A W_{t}$ is the set of workers in the all-workers frame in year $t$.

Table A.1: LEHD Regimes

| Count | State | $\begin{gathered} \text { First } \\ \text { YYYYQ } \end{gathered}$ | $\begin{gathered} \text { Last } \\ \text { YYYYQ } \end{gathered}$ | Entry Order | Pct 2012Q1 QCEW Emp |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Regime 1-1990Q1 to 2013Q4-19.35\% of 2012Q1 QCEW Employment |  |  |  |  |  |
| 1 | Maryland | 1985Q2 | 2014Q3 | 1 | 1.83\% |
| 2 | Alaska | 1990Q1 | 2014Q3 | 2 | 0.22\% |
| 3 | Colorado | 1990Q1 | 2014Q3 | 3 | 1.70\% |
| 4 | Idaho | 1990Q1 | 2014Q3 | 4 | 0.45\% |
| 5 | Illinois | 1990Q1 | 2014Q3 | 5 | 4.38\% |
| 6 | Indiana | 1990Q1 | 2014Q3 | 6 | 2.19\% |
| 7 | Kansas | 1990Q1 | 2013Q4 | 7 | 0.98\% |
| 8 | Louisiana | 1990Q1 | 2014Q2 | 8 | 1.41\% |
| 9 | Missouri | 1990Q1 | 2014Q3 | 9 | 1.99\% |
| 10 | Washington | 1990Q1 | 2014Q3 | 10 | 2.12\% |
| 11 | Wisconsin | 1990Q1 | 2014Q3 | 11 | 2.08\% |
| Regime 2-1995Q1 to 2013Q4-48.28\% of 2012Q1 QCEW Employment |  |  |  |  |  |
| 12 | North Carolina | 1991Q1 | 2014Q3 | 1 | 2.92\% |
| 13 | Oregon | 1991Q1 | 2014Q3 | 2 | 1.23\% |
| 14 | Pennsylvania | 1991Q1 | 2014Q3 | 3 | 4.44\% |
| 15 | California | 1991Q3 | 2014Q3 | 4 | 11.37\% |
| 16 | Arizona | 1992Q1 | 2014Q3 | 5 | 1.85\% |
| 17 | Wyoming | 1992Q1 | 2014Q3 | 6 | 0.19\% |
| 18 | Florida | 1992Q4 | 2014Q2 | 7 | 5.78\% |
| 19 | Montana | 1993Q1 | 2014Q3 | 8 | 0.31\% |
| 20 | Georgia | 1994Q1 | 2014Q3 | 9 | 2.90\% |
| 21 | South Dakota | 1994Q1 | 2014Q2 | 10 | 0.30\% |
| 22 | Minnesota | 1994Q3 | 2014Q3 | 11 | 2.05\% |
| 23 | New York | 1995Q1 | 2014Q3 | 12 | 6.49\% |
| 24 | Rhode Island | 1995Q1 | 2014Q3 | 13 | 0.35\% |
| 25 | Texas | 1995Q1 | 2014Q3 | 14 | 8.10\% |
| Regime 3-1998Q1 to 2013Q4-17.66\% of 2012Q1 QCEW Employment |  |  |  |  |  |
| 26 | New Mexico | 1995Q3 | 2014Q3 | 1 | 0.55\% |
| 27 | Hawaii | 1995Q4 | 2014Q3 | 2 | 0.44\% |
| 28 | Connecticut | 1996Q1 | 2014Q3 | 3 | 1.26\% |
| 29 | Maine | 1996Q1 | 2014Q3 | 4 | 0.43\% |
| 30 | New Jersey | 1996Q1 | 2014Q3 | 5 | 2.87\% |
| 31 | Kentucky | 1996Q4 | 2014Q3 | 6 | 1.32\% |
| 32 | West Virginia | 1997Q1 | 2014Q3 | 7 | 0.52\% |
| 33 | Michigan | 1998Q1 | 2014Q3 | 8 | 3.04\% |
| 34 | Nevada | 1998Q1 | 2014Q3 | 9 | 0.89\% |
| 35 | North Dakota | 1998Q1 | 2014Q3 | 10 | 0.31\% |
| 36 | South Carolina | 1998Q1 | 2014Q3 | 11 | 1.35\% |
| 37 | Tennessee | 1998Q1 | 2014Q3 | 12 | 2.03\% |
| 38 | Virginia | 1998Q1 | 2014Q2 | 13 | 2.65\% |
| Regime 4-2004Q1 to 2013Q4-14.71\% of 2012Q1 QCEW Employment |  |  |  |  |  |
| 39 | Delaware | 1998Q3 | 2014Q3 | 1 | 0.31\% |
| 40 | Iowa | 1998Q4 | 2014Q3 | 2 | 1.12\% |
| 41 | Nebraska | 1999Q1 | 2014Q3 | 3 | 0.69\% |
| 42 | Utah | 1999Q1 | 2014Q3 | 4 | 0.91\% |
| 43 | Ohio | 2000Q1 | 2014Q3 | 5 | 3.93\% |
| 44 | Oklahoma | 2000Q1 | 2014Q3 | 6 | 1.11\% |
| 45 | Vermont | 2000Q1 | 2014Q3 | 7 | 0.22\% |
| 46 | Alabama | 2001Q1 | 2014Q3 | 8 | 1.34\% |
| 47 | Massachusetts | 2002Q1 | 2014Q2 | 9 | 2.55\% |
| 48 | District of Columbia | 2002Q2 | 2014Q3 | 10 | 0.43\% |
| 49 | Arkansas | 2002Q3 | 2014Q3 | 11 | 0.86\% |
| 50 | New Hampshire | 2003Q1 | 2014Q2 | 12 | 0.47\% |
| 51 | Mississippi | 2003Q3 | 2014Q3 | 13 | 0.77\% |

Notes: This table presents information on the states that make up each date regime. Each panel gives the first and last quarter of data available, the entry order, and the employment coverage (percent of 2012Q1 private QCEW employment) of each state in the regime. OPM data for federal workers is not shown in this table, but is available beginning in 2000 Q 1 .

Figure A.1: Immigrant Candidates - Excluded Earnings Records


Notes: This figure presents the count of earnings records excluded from the eligible-workers frame each year, disaggregated by the different eligibility requirements the record failed to meet: (i) Invalid SSN are records that are only on the UI; (ii) Age<5 are records where the SNN is valid, but the age of the worker is less than 5; (iii) $5 \leq A g e<13$ are records where the worker is between 5 and 13 years old; (iv) $13 \leq$ Age $<18$ are records where the worker is between 13 and 18 years old; (v) Age> 70 are records where the worker is more than 70 years old; (vi) \#Jobs $>12$ are records where the worker has more than 12 jobs a year; and (vii) Other are records that fail to meet the other eligibility requirements, such as the year is greater than or equal to the SSN year of issue and less than year of death (when available).

Table A.2: Immigrant Candidates - Excluded Earnings Records

| Year | Total | Invalid SSN | Age<5 | 5 $\geq$ Age $<13$ | 13 $\leq$ Age $<18$ | Age $>70$ | \#Jobs>12 | Other |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1990 | 2,173,054 | 131,768 | 92,173 | 115,966 | 1,383,852 | 302,791 | 61,336 | 85,168 |
| 1991 | 2,029,041 | 156,980 | 96,503 | 110,535 | 1,228,937 | 300,232 | 53,311 | 82,543 |
| 1992 | 2,024,225 | 161,800 | 99,380 | 111,528 | 1,199,329 | 310,526 | 55,873 | 85,789 |
| 1993 | 2,227,908 | 204,299 | 123,925 | 122,587 | 1,294,809 | 333,303 | 59,024 | 89,961 |
| 1994 | 2,546,460 | 228,963 | 145,038 | 136,015 | 1,500,927 | 363,506 | 74,634 | 97,377 |
| 1995 | 9,875,811 | 1,240,177 | 939,315 | 676,532 | 4,536,074 | 1,695,371 | 337,545 | 450,797 |
| 1996 | 10,144,571 | 1,282,244 | 1,020,460 | 731,340 | 4,625,974 | 1,649,645 | 377,807 | 457,101 |
| 1997 | 10,560,373 | 1,318,787 | 1,051,685 | 773,013 | 4,802,606 | 1,737,019 | 408,080 | 469,183 |
| 1998 | 13,680,138 | 1,579,419 | 1,227,565 | 942,868 | 6,460,058 | 2,308,455 | 571,745 | 590,028 |
| 1999 | 14,850,424 | 1,801,636 | 1,328,052 | 1,059,582 | 6,864,218 | 2,559,284 | 617,195 | 620,457 |
| 2000 | 15,909,402 | 2,087,866 | 1,441,233 | 1,147,779 | 7,084,996 | 2,826,633 | 671,695 | 649,200 |
| 2001 | 15,142,444 | 2,313,768 | 1,354,067 | 1,109,587 | 6,313,180 | 2,864,144 | 565,342 | 622,356 |
| 2002 | 13,646,946 | 2,030,273 | 1,168,828 | 988,866 | 5,573,020 | 2,784,977 | 519,677 | 581,305 |
| 2003 | 13,105,529 | 2,260,426 | 1,059,202 | 965,151 | 4,979,593 | 2,776,405 | 493,455 | 571,297 |
| 2004 | 15,254,789 | 2,628,435 | 1,099,414 | 1,087,743 | 5,976,072 | 3,254,876 | 561,150 | 647,099 |
| 2005 | 16,109,360 | 2,881,580 | 1,030,810 | 1,240,576 | 6,271,025 | 3,383,095 | 626,426 | 675,848 |
| 2006 | 16,830,576 | 3,071,079 | 959,130 | 1,332,606 | 6,513,877 | 3,564,841 | 686,925 | 702,118 |
| 2007 | 16,464,027 | 3,109,359 | 860,258 | 1,254,957 | 6,233,964 | 3,605,470 | 712,999 | 687,020 |
| 2008 | 14,509,746 | 2,909,378 | 683,388 | 1,081,938 | 5,135,680 | 3,478,821 | 564,086 | 656,455 |
| 2009 | 11,701,711 | 2,484,829 | 471,798 | 884,181 | 3,620,311 | 3,240,941 | 390,427 | 609,224 |
| 2010 | 11,019,697 | 2,328,456 | 382,395 | 816,592 | 3,283,378 | 3,210,027 | 402,839 | 596,010 |
| 2011 | 10,942,606 | 2,307,310 | 315,743 | 767,636 | 3,269,325 | 3,224,106 | 450,244 | 608,242 |
| 2012 | 11,556,277 | 2,822,199 | 240,123 | 742,658 | 3,386,957 | 3,282,004 | 449,498 | 632,838 |
| 2013 | 13,216,695 | 4,157,518 | 178,979 | 671,775 | 3,622,084 | 3,409,276 | 492,710 | 684,353 |

Notes: The first column presents of the total number of earnings records excluded from the eligible-workers frame each year. The remaining columns disaggregate this count by the different eligibility requirements the record failed to meet: (i) Invalid SSN are records that are only on the UI; (ii) Age $<5$ are records where the SNN is valid, but the age of the worker is less than 5 ; (iii) $5 \leq$ Age $<13$ are records where the worker is between 5 and 13 years old; (iv) $13 \leq$ Age $<18$ are records where the worker is between 13 and 18 years old; (v) Age $>70$ are records where the worker is more than 70 years old; (vi) \#Jobs $>12$ are records where the worker has more than 12 jobs a year; and (vii) Other are records that fail to meet the other eligibility requirements, such as the year is greater than or equal to the SSN year of issue and less than year of death (when available). The frame is complete from 2004 forward.

## A. 2 Comparison to NIPA

The BEA NIPA estimates are based primarily on the BLS Quarterly Census of Employment and Wages (QCEW), an alternative source of employment and earnings with similar coverage as the UI based job level data used in this paper. A firm typically files a QCEW firm-level report in conjunction with the UI job-level data each quarter. The QCEW report is sent to BLS where it undergoes edits and imputations before the final statistics are released. ${ }^{53}$ These data are then used by BEA as the primary input when estimating the wage and salary component of the NIPA tables. ${ }^{54}$

Table A. 3 presents a comparison of our estimates of annual earnings with the BEA NIPA data. Figure A. 2 plots this comparison. Our estimates of total annual earnings using the all-workers frame vary from $16.5 \%$ of NIPA wage and salary estimates in 1990, the beginning of LEHD date regime 1 ; to $60.1 \%$ in 1995, the beginning of date regime 2 ; to $76.4 \%$ in 1998, the beginning of date regime 3 ; to $90.6 \%$ in 2004, the beginning of date regime 4 . Once LEHD data are complete in 2004, the two series track almost exactly. By 2013 the all-workers estimate is about $91.7 \%$ of

[^0]the NIPA wage and salary estimates. The eligible-workers estimates follow a similar pattern as the all-workers estimates, with about two percentage points lower coverage relative to the all-workers frame after 2004.

The coverage of both the all-workers and eligible-workers frames is very low relative to the NIPA estimates in the early 1990s but increases dramatically in 1995 once the historical data for the more populous states (CA, FL, NY, and TX) have entered the LEHD infrastructure files. When the frame is complete (date regime 4), there is an apparent coverage gap of about eight to nine percentage points for the all-workers frame and ten to 11 percentage points for the eligible-workers frame. About half of this gap is due to differences between the statutory-employer population for UI wage records and the NIPA definition of wage and salary income. When comparing frames with similar coverage definitions (UI wage records vs. QCEW), our results suggest that the gap between the two frames is about four to five percentage points for the all-workers frame and five to six percentage points for the eligible-workers frame.

Table A.3: Earnings Measures-National Income and Product Accounts versus LEHD Data

| Year | NIPA Wage and Salary | LEHD <br> Total | Eligible <br> Workers | Immigrant Candidates |
| :---: | :---: | :---: | :---: | :---: |
| 1990 | 3,611.6 | 594.7 | 587.4 | 7.3 |
| 1991 | 3,558.4 | 593.2 | 585.7 | 7.5 |
| 1992 | 3,639.8 | 611.2 | 603.8 | 7.4 |
| 1993 | 3,669.6 | 609.3 | 601.6 | 7.7 |
| 1994 | 3,760.7 | 642.6 | 633.9 | 8.7 |
| 1995 | 3,862.1 | 2,319.3 | 2,279.7 | 39.6 |
| 1996 | 3,969.2 | 2,336.3 | 2,294.5 | 41.8 |
| 1997 | 4,159.4 | 2,494.2 | 2,448.4 | 45.8 |
| 1998 | 4,417.6 | 3,374.6 | 3,312.9 | 61.7 |
| 1999 | 4,607.8 | 3,539.3 | 3,469.8 | 69.5 |
| 2000 | 4,825.9 | 3,770.5 | 3,694.7 | 75.8 |
| 2001 | 4,817.3 | 3,785.9 | 3,707.7 | 78.2 |
| 2002 | 4,782.5 | 3,743.2 | 3,666.4 | 76.8 |
| 2003 | 4,808.3 | 3,739.8 | 3,663.8 | 76.0 |
| 2004 | 4,942.6 | $4,478.7$ | 4,387.3 | 91.4 |
| 2005 | 5,018.8 | 4,565.8 | 4,469.3 | 96.5 |
| 2006 | 5,174.0 | 4,716.5 | 4,613.0 | 103.5 |
| 2007 | 5,312.4 | 4,842.3 | 4,736.2 | 106.1 |
| 2008 | 5,224.3 | 4,767.6 | 4,667.3 | 100.3 |
| 2009 | 5,018.6 | 4,579.8 | 4,489.4 | 90.4 |
| 2010 | 5,037.6 | 4,593.4 | 4,503.7 | 89.7 |
| 2011 | 5,078.9 | 4,630.1 | 4,539.8 | 90.3 |
| 2012 | 5,197.7 | 4,750.8 | 4,652.5 | 98.3 |
| 2013 | 5,257.9 | 4,822.0 | 4,706.0 | 116.0 |

Notes: This table compares total earnings as measured in the BEA NIPA to earnings computed from LEHD. LEHD Total presents total annual earnings for the all-workers frame. This total is decomposed into earnings attributed to workers included in the eligible-workers frame (Eligible Workers) and to workers who are not included (Immigrant Candidates). Units are in billions of real (2000) dollars, converted using CPI-U. The frame is complete from 2004 forward.

Figure A.2: NIPA


Notes: This figure compares total earnings as measured in BEA NIPA (blue line) to earnings computed from LEHD using all workers (red line).

## A. 3 Estimation of the Earnings/Inactivity Distribution

Table A.4: Labor Force Activity of Workers in Each Earnings Bin

| Quarters <br> Worked | Longest Job | Workers |  | Jobs (Avg) | Earnings (Avg) |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Counts | Percent |  |  |
| Bottom 20\% of Earnings Distribution |  |  |  |  |  |
| 1 | 1 | 8,543,957 | 30.6\% | 1.066 | \$1,366 |
| 2 | 1 | 1,883,159 | 6.7\% | 1.996 | \$2,187 |
| 2 | 2 | 5,806,138 | 20.8\% | 1.213 | \$2,824 |
| 3 | 1 | 520,324 | 1.9\% | 2.594 | \$3,029 |
| 3 | 2 | 2,467,851 | 8.8\% | 2.297 | \$3,480 |
| 3 | 3 | 2,591,936 | 9.3\% | 1.263 | \$3,726 |
| 4 | 1 | 58,758 | 0.2\% | 4.542 | \$3,480 |
| 4 | 2 | 949,367 | 3.4\% | 3.429 | \$4,274 |
| 4 | 3 | 932,150 | 3.3\% | 2.602 | \$4,544 |
| 4 | 4 | 187,115 | 0.7\% | 1.716 | \$4,161 |
| 4 | 5 | 1,078,088 | 3.9\% | 1.440 | \$4,178 |
| 4 | 6 | 2,893,038 | 10.4\% | 1.251 | \$4,227 |
| Middle 60\% of Earnings Distribution |  |  |  |  |  |
| 1 | 1 | 853,497 | 1.0\% | 1.051 | \$13,637 |
| 2 | 1 | 489,513 | 0.6\% | 1.643 | \$14,924 |
| 2 | 2 | 2,697,567 | 3.2\% | 1.176 | \$14,375 |
| 3 | 1 | 680,994 | 0.8\% | 1.475 | \$19,879 |
| 3 | 2 | 2,409,536 | 2.9\% | 2.119 | \$15,891 |
| 3 | 3 | 4,976,450 | 5.9\% | 1.233 | \$17,446 |
| 4 | 1 | 52,620 | 0.1\% | 3.726 | \$17,579 |
| 4 | 2 | 2,746,891 | 3.3\% | 3.287 | \$17,604 |
| 4 | 3 | 7,105,740 | 8.5\% | 2.592 | \$20,563 |
| 4 | 4 | 841,481 | 1.0\% | 2.109 | \$19,230 |
| 4 | 5 | 8,869,511 | 10.6\% | 1.602 | \$22,405 |
| 4 | 6 | 52,012,001 | 62.1\% | 1.212 | \$26,107 |
| Top 20\% of Earnings Distribution |  |  |  |  |  |
| 1 | 1 | 75,101 | 0.3\% | 1.038 | \$146,574 |
| 2 | 1 | 34,381 | 0.1\% | 1.361 | \$138,531 |
| 2 | 2 | 112,925 | 0.4\% | 1.096 | \$102,246 |
| 3 | 1 | 94,047 | 0.3\% | 1.178 | \$92,110 |
| 3 | 2 | 171,999 | 0.6\% | 1.605 | \$95,079 |
| 3 | 3 | 434,213 | 1.6\% | 1.128 | \$89,432 |
| 4 | 1 | 7,589 | 0.0\% | 2.608 | \$90,693 |
| 4 | 2 | 312,325 | 1.1\% | 2.752 | \$84,965 |
| 4 | 3 | 1,383,555 | 5.0\% | 2.323 | \$87,727 |
| 4 | 4 | 139,347 | 0.5\% | 1.993 | \$90,280 |
| 4 | 5 | 2,493,150 | 8.9\% | 1.500 | \$92,054 |
| 4 | 6 | 22,653,328 | 81.2\% | 1.181 | \$88,447 |

Notes: Each row in the table represents a specific combination of quarters worked and number of quarters in the longest job. A five quarter longest job is active in either the fourth quarter of the previous year or the first quarter of the subsequent year, while a six quarter longest job is active in both. The number of quarters in the longest job takes on values from one to six. The counts are averages per year.

## B Inequality Trends in the LEHD All-Workers Frame (1990-2013)

In Section 3 we discussed the trends in earnings inequality observed in the eligible-workers frame. Here, we detail the inequality trends in the all-workers frame, and analyze how they differ from the trends observed in the eligible-workers frame.

With a better understanding of how the exclusion of specific workers affects the distribution of earnings, we then turn our attention to earnings inequality. We analyze how various measurements of the gap between the top and bottom of the earnings distribution have changed over time and how the trends change as we move from the all-workers to the eligible-workers frame.

Figure B. 1 plots selected percentiles for the two worker frames: the solid lines are the percentiles computed from the all-workers frame, while the dotted lines are the percentiles computed from the eligible-workers frame. Comparing the solid and dotted lines in Figure B.1, it is clear that the main consequence of shifting the frame from all workers to eligible workers is an increase in the percentile values, particularly at the bottom of the earnings distribution.

Figure B.1: Percentiles of the Earnings Distribution by Worker Frame


Notes: This figure plots the $10^{t h}, 20^{t h}, 50^{t h}, 80^{t h}, 90^{t h}$, and $95^{t h}$ percentiles of the earnings distribution by worker frame and year. The solid lines are the percentiles of the earnings distribution of all workers (AW) by year. The dotted lines are the percentiles of the earnings distribution of eligible workers (EW) by year.

Figure B. 2 plots the ratio of the $90^{\text {th }}$ percentile to the $10^{\text {th }}$ percentile for each date regime using the all-workers frame. The figure confirms that there are some differences in the levels of these curves but the trend analysis is largely unchanged.

Figure B.2: Ratio of the 90th to the 10th Percentile of the Earnings Distribution


Notes: This figure plots the ratio of the 90 th to the 10 th percentile for all workers by date regime.
To see this more clearly, Table B. 1 presents the average percentile values from 1995-2013 for both the all-workers and the eligible-workers frames, and the last row computes their ratio (eligible workers to all workers). First, notice that the ratio is always above one, meaning that each percentile computed from the eligible-workers frame is greater than the equivalent percentile computed from the all-workers frame. Removing the immigrant candidates from the all-workers frame to construct the eligible-workers frame eliminates an unknown number of individuals who make very low earnings and, thus, tend to be at the bottom of the all-workers earnings distribution. For example, in 2006, immigrant candidates held about $8 \%$ of all jobs, but only contributed about $2 \%$ to total earnings. Furthermore, average earnings for immigrant candidates were about $\$ 6,150$ in 2006 as compared with $\$ 32,865$ for eligible workers. Thus, the removal of these low-earnings workers from the all-workers frame makes the ratio of EW to AW percentiles in Table B. 1 higher towards the bottom of the earnings distribution. Specifically, notice that the $1^{\text {st }}$ percentile in the eligible-workers earnings distribution is, on average, about $32 \%$ greater than the $1^{\text {st }}$ percentile in the all-workers earnings distribution; the $5^{t h}$ percentile is about $41 \%$ greater, the $10^{t h}$ percentile is about $36 \%$ greater, and the $20^{\text {th }}$ percentile is about $26 \%$ greater. From the median onwards, while the absolute differences in the percentile values are large, the relative differences are not as stark, with the percentiles in the eligible-workers earnings distribution being about $2 \%$ to $8 \%$ greater than the corresponding percentile in the all-workers earnings distribution. Finally, notice that regardless of the worker frame used, there is a large number of workers with very low earnings in LEHD, with the average $10^{\text {th }}$ percentile at only $\$ 1,858$ in the all-workers frame and $\$ 2,527$ in the eligible-workers frame.

Table B.1: Average Percentiles of the Earnings Distribution by Worker Frame (1995-2013)

| Frame | Percentiles |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 st | 5th | 10th | 20th | 50th | 80th | 90th | 95th | 99th |
| All Workers | 100 | 713 | 1,858 | 5,141 | 20,093 | 43,741 | 62,277 | 84,012 | 173,847 |
| Eligible Workers | 132 | 1,005 | 2,527 | 6,463 | 21,762 | 45,343 | 64,021 | 86,108 | 178,304 |
| Ratio of EW to AW | 1.3195 | 1.4088 | 1.3605 | 1.2572 | 1.0831 | 1.0366 | 1.0280 | 1.0249 | 1.0256 |

Notes: The first row presents the average percentile values from the earnings distribution of all workers in all states from 1995-2013. The second row presents the average percentile values from the earnings distribution of eligible workers in all states from 1995-2013. The last row computes the ratio of each percentile from the eligible-workers frame to all-workers frame.

Starting with the all-workers frame in Figure B.3, notice that all the measures show a decline in earnings inequality from 1995 to 2000. This can also be seen in Table B.2. The first row presents the average of each ratio from 1995-1999. Notice that they are all above one, meaning that earnings inequality was greater in the late 1990s than in 2000 . Then, after 2000, except for the $99 / 1$ ratio, which has a slight upward trend, all other measures of earnings inequality remain relatively stable. The second row of Table B. 2 presents the average of each ratio (relative to 2000) from 2001-2013. Notice that aside from the $99 / 1$ ratio, which on average increased by about $5 \%$ after 2000 , the other measures have remained around their 2000 levels. Thus, aside from differences at the very top or the very bottom of the earnings distribution, earnings inequality among all workers has apparently seen little or no change over the last 10 years.

Figure B.3: Selected Inequality Measures 1990-2013, Relative to 2000 (All Workers)


Notes: Subplot (a) presents measures of earnings inequality for all workers in all states relative to 2000 from 1990-2013. Subplot (b) presents measures of earnings inequality for eligible workers in all states relative to 2000 from 1990-2013. The measures of earnings inequality considered are (i) P99 to P1: the ratio of the 99th to the 1st percentile; (ii) P95 to P5: the ratio of the 95th to the 5th percentile; (iii) P90 to P10: the ratio of the 90th to the 10th percentile; (iv) P80 to P20 the ratio of the 80 th to the 20 th percentile; and (v) Variance: the variance of $\log$ annual earnings.

Table B.2: Inequality Measures Relative to 2000 by Worker Frame

|  | Inequality Measures |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | 99th/1st | 95th/5th | 90th/10th | 80th/20th | Variance |
| All Workers |  |  |  |  |  |
| Pre-2000 | 1.038 | 1.099 | 1.092 | 1.075 | 1.036 |
| Post-2000 | 1.050 | 1.010 | 1.004 | 0.983 | 1.004 |
| Pre-GR | 1.001 | 1.003 | 1.005 | 0.992 | 1.001 |
| $G R$ | 1.059 | 1.009 | 1.005 | 0.978 | 1.006 |
| Post-GR | 1.131 | 1.022 | 1.002 | 0.968 | 1.007 |
| Eligible Workers |  |  |  |  |  |
| Pre-2000 | 1.085 | 1.119 | 1.103 | 1.080 | 1.047 |
| Post-2000 | 1.154 | 1.136 | 1.114 | 1.064 | 1.054 |
| Pre-GR | 1.063 | 1.075 | 1.067 | 1.039 | 1.031 |
| $G R$ | 1.209 | 1.181 | 1.151 | 1.084 | 1.073 |
| Post-GR | 1.286 | 1.222 | 1.178 | 1.099 | 1.086 |

Notes: The first panel presents measures of earnings inequality for all workers in all states relative to 2000, while the second panel presents the same measures for eligible workers. The measures of earnings inequality considered are (i) $99 t h / 1 s t$ : the ratio of the 99 th to the 1st percentile; (ii) $95 t h / 5$ th: the ratio of the 95 th to the 5 th percentile; (iii) 90th/10th: the ratio of the 90th to the 10th percentile; (iv) 80th/20th the ratio of the 80th to the 20th percentile; and (v) Variance: the variance of log annual earnings. The values in the table are averages before and after 2000: (i) pre-2000: 1995-1999; and (ii) post-2000: 2001-2013. The post-2000 years are further subdivided into three periods: (i) pre-GR: 2001-2007; (ii) GR: 2008-2009; and (iii) post-GR 2010-2013.

Figure B.4: Selected Inequality Measures for the Top and Bottom of the Earnings Distribution 1990-2013, Relative to 2000 (All Workers)

(a) All Workers - Top
(b) All Workers - Bottom

Notes: Subplots (a) and (b) decompose the $99 / 1$ ratio, the $95 / 5$ ratio, the $90 / 10$ ratio, and the $80 / 20$ for eligible workers in all states relative to 2000 from 1990-2013 relative to the median. Subplot (a) plots the following ratios for the top half of the earnings distribution: (i) P99 to P50: the ratio of the $99^{t h}$ to the $50^{t h}$ percentile; (ii) P95 to P50: the ratio of the $95^{t h}$ to the $50^{t h}$ percentile; (iii) P90 to $P 50$ : the ratio of the $90^{t h}$ to the $50^{t h}$ percentile; and (iv) P80 to $P 50$ the ratio of the $80^{t h}$ to the $50^{t h}$ percentile. Subplot (b) plots the following ratios for the bottom half of the earnings distribution: (i) P50 to P1: the ratio of the $50^{t h}$ to the $1^{\text {st }}$ percentile; (ii) P50 to P5: the ratio of the $50^{t h}$ to the $5^{t h}$ percentile; (iii) P50 to P10: the ratio of the $50^{t h}$ to the $10^{t h}$ percentile; and (iv) P50 to P20 the ratio of the $50^{t h}$ to the $20^{t h}$ percentile. The estimates are based on the all-workers frame from the LEHD infrastructure files.

Figure B.5: Percentile Ratios of the Earnings Distribution by Worker Frame


Notes: This figure plots ratios of top and bottom percentiles for all workers (solid lines) and for eligible workers (dotted lines). Subplot (a) plots the ratio of the 99th to the 1st percentile by worker frame. This ratio is decomposed into the ratio of the 99th to the 50 th percentile in subplot (c) and the ratio of the 50 th to the 1 st percentile in subplot (e). Subplot (b) plots the ratio of the 95 th to the 15 th percentile by worker frame. This ratio is decomposed into the ratio of the 95 th to the 50 th percentile in subplot (d) and the ratio of the 50 th to the 5 th percentile in subplot (f).

Figure B.6: Percentile Ratios of the Earnings Distribution by Worker Frame


Notes: This figure plots ratios of top and bottom percentiles for all workers (solid lines) and for eligible workers (dotted lines). Subplot (a) plots the ratio of the 90th to the 10 th percentile by worker frame. This ratio is decomposed into the ratio of the 90 th to the 50 th percentile in subplot (c) and the ratio of the 50 th to the 10 th percentile in subplot (e). Subplot (b) plots the ratio of the 80th to the 20 th percentile by worker frame. This ratio is decomposed into the ratio of the 80 th to the 50 th percentile in subplot (d) and the ratio of the 50th to the 20 th percentile in subplot (f).

## C Comparison with Household Surveys

In Section 3, we discussed the trends in earnings inequality based on our analysis of the eligibleworkers frame, which we constructed using the LEHD infrastructure data and supplementary information from the Census Bureau's enhanced version of SSA's Numident file. Section 3.1 discussed the highlights of the comparison of our data to the Current Population Survey and the American Community Survey. To put our inequality measures in the context of a broader literature, we compare results based on the administrative data frame discussed in main text with similar measures constructed using household survey data. ${ }^{55}$

## C. 1 Household Survey Data

To create our household survey analysis file, we use the following records from the Current Population Survey-Annual Social and Economic Supplement (March) and the American Community Survey:

- Current Population Survey Annual Social and Economic Supplement (CPS-ASEC): all persons from survey years 1990-2004
- American Community Survey (ACS): all persons from survey years 2000 to 2013

In the CPS-ASEC, the respondent is surveyed in March and reports earnings for the previous calendar year. We date the earnings accordingly. However, in the ACS the respondent reports earnings for the past 12 months and the survey is in the field continuously throughout the year. Our approach in this case is to date the earnings with the calendar year containing the majority of the months covered by the response, with ties going to the more recent year. As in the LEHD data, nominal earnings are deflated to real 2000 dollars using the CPI-U. In all cases, we used the internal (confidential) versions of the CPS-ASEC and ACS. None of the household survey data are topcoded. We did not replace the Census Bureau's edit and imputation routines with our own. We used the allocated values in the files.

Similar to the workers in LEHD, we consider two samples of individuals from the household surveys. The first includes all individuals. The second isolates workers whose employment should be covered under Unemployment Insurance (including federal employees) and who should, therefore, appear in the LEHD administrative data. We designate a survey respondent as a "covered worker" if he or she meets the following conditions:

- Person interviewed is not living in group quarters
- Individual is employed at a private firm, the local/state/federal government, or is selfemployed in an incorporated firm
- Labor earnings are positive
- Individual is between 18 and 70 years old, inclusive.

The last two restrictions combined are included to match the earnings and age restrictions used to identify active eligible workers in the LEHD data.

Finally, in most of the results to follow, we do not report results separately for CPS-ASEC and ACS individuals. Instead, in the overlapping years (2000-2003), we interpolate estimates computed from the CPS-ASEC and the ACS to create a single time series using the method in Abowd and Vilhuber (2011).

[^1]
## C. 2 Comparison of Aggregate Summaries

We start by analyzing how the earnings distribution in household surveys compares to the one computed from administrative records. In the household survey data, the estimated percentile values tend to be greater for covered workers than for all workers. Figure C. 1 presents the percentiles of the earnings distribution for all and covered workers in the CPS/ACS. Comparing these values to the ones estimated from the LEHD data, shown in Figure B.1, notice that for percentiles above the median, the values from the eligible-workers frame are fairly close to the ones from the household surveys. Below the median, however, the differences are greater, with the percentiles estimated from the household surveys being much greater than the percentiles estimated from LEHD. For example, notice that earnings associated with the $10^{\text {th }}$ percentile in the CPS/ACS data are close to the $20^{t h}$ percentile in LEHD.

Figure C.1: Percentiles of the Earnings Distribution for All and Covered Workers by Year


Notes: This figure plots the $10^{t h}, 20^{t h}, 50^{t h}, 80^{t h}, 90^{t h}$, and $95^{t h}$ percentiles of the earnings distribution for all workers (ALL) and for covered workers (CW) in the CPS-ASEC (1990-2003) and the ACS (2000-2013) surveys by year. Covered workers include respondents whose employment relation should be covered as a statutory employee under state UI or as a federal employee, and therefore appear in the LEHD data.

To see these differences in percentiles more clearly, Figure C. 2 plots the ratio of the percentiles of the earnings distribution measured using the LEHD eligible-workers frame to the same percentiles measured from the covered workers in the combined CPS/ACS data. First, notice that all the ratios in Figure C. 2 are below one, meaning that the percentiles estimated from the household surveys are always greater than the corresponding percentiles estimated from the administrative data. However, the magnitude of this difference varies greatly across the percentiles of the earnings distribution. Specifically, notice that the relative differences in the $95^{t h}, 90^{t h}$, and $80^{t h}$ percentiles are very small compared to the relative differences in the $5^{t h}, 10^{t h}$, and $20^{t h}$ percentiles.

In the main text, Table 4 presents averages of the percentiles from 1995 to 2013 for CPS/ACS and LEHD. Notice that on average the earnings associated with the $80^{t h}, 90^{t h}$, and $95^{t h}$ percentiles are about $\$ 3,500$ less in the LEHD data than in CPS/ACS data. Further, as can be seen in

Figure C.2, this gap is decreasing over time, such that in 2013 the difference in the $95^{\text {th }}$ percentile is only $\$ 264$. In the bottom half of the earnings distribution, however, a CPS/ACS covered worker earns about $\$ 4,000$ more than an LEHD worker at the same point in the earnings distribution. While this absolute difference may not seem that large, relatively, a CPS/ACS worker at the $10^{\text {th }}$ percentile is making 2.54 times more than his LEHD counterpart, and 3.40 times more for a CPS/ACS worker in the $5^{t h}$ percentile. This means that the survey data include more low-earning jobs that are not statutory employment relationships, or are not reported as such to state UI systems. Lastly, notice that the percentiles in LEHD increase faster than their CPS/ACS counterparts since all the ratios exhibit an upward trend, especially after the Great Recession.

Figure C.2: Ratio of UI Earnings to Household Survey Reported Earnings by Percentile


Notes: This figure presents the ratio of earnings for eligible workers in LEHD to the earnings reported by covered workers in CPS/ACS for the $5^{t h}, 10^{t h}, 20^{t h}, 50^{t h}, 80^{t h}, 90^{t h}$, and $95^{t h}$ percentile.

To see whether differences in the earnings distribution between workers in CPS/ACS and eligible workers in the LEHD data translate into differences in trends in inequality, we compute various measures of earnings inequality in CPS/ACS and compare them to their LEHD counterparts. In particular, we compute the $95 / 5,90 / 10$, and $80 / 20$ ratios, and the variance of $\log$ annual earnings. We plot their time series in Figures 5 a and 5 b for all workers and covered workers, respectively. Both the all-workers and the covered-workers samples show a decline in earnings inequality during the late 1990s that reverses after 2000. However, in the all-workers sample, the magnitude of this increase in inequality in the post-2000 period strongly depends on the measure considered. For example, from 2000 to 2013 , the $95 / 5$ ratio increased by $66 \%$ from 36.30 to 60.26 , while the $90 / 10$ ratio increased by $42 \%$ from 12.95 to 18.35 . On the other hand, the $80 / 20$ ratio increased by $26 \%$ from 4.64 to 5.86 , while the variance of log earnings increased by only $5 \%$ from 1.23 to 1.26. Thus, while the measures are all trending upwards after 2000 in the all-workers sample, it is unclear whether this increase has been large or small. In the covered-workers sample, earning inequality has also been increasing after 2000, however the magnitude of this increase is relatively consistent across the different measures of earnings inequality. The $95 / 5$ ratio increased by $32 \%$ from 2000 to 2013 , while the $90 / 10$ ratio increased by $26 \%$. The $80 / 20$ ratio and the variance in $\log$ earnings increased less over this period, by about $13 \%$ and $14 \%$, respectively. On the other hand, notice that the decline in inequality in the 1990s is very similar across the various measures and samples.

These trends in earnings inequality are very similar to the ones observed among eligible
workers in the LEHD data. Specifically, comparing the time series of earnings inequality for covered workers in CPS/ACS (Figure 5b) to the one for eligible workers in LEHD (Figure 3), notice that the general patterns are very similar. Both of these figures show a decline in inequality during the 1990s and a steady increase in inequality after 2000. The magnitude of this increase is also similar between the covered workers in CPS/ACS and the eligible workers in the LEHD data. Compare the second panel of Table C. 1 to the second panel of Table B.2. The second row in both tables shows the average of the $95 / 5$ ratio, the $90 / 10$, ratio, the $80 / 20$ ratio, and the variance of log earnings (relative to 2000) after 2000. Both the covered workers in CPS/ACS and the eligible workers in LEHD saw an increase in the $95 / 5$ ratio and the $90 / 10$ ratio above $10 \%$, and an increase in the $80 / 20$ ratio and the variance of $\log$ earnings around $5-6 \%$. Furthermore, most of this increase occurred during or after the Great Recession.

Table C.1: Inequality Measures Relative to 2000 for Workers in Household Surveys

|  | 95th/5th | nequality <br> 90th/10th | Measures 80th/20th | Variance |
| :---: | :---: | :---: | :---: | :---: |
| All Workers |  |  |  |  |
| Pre-2000 | 1.126 | 1.118 | 1.041 | 1.099 |
| Post-2000 | 1.406 | 1.273 | 1.174 | 1.001 |
| Pre-GR | 1.280 | 1.171 | 1.099 | 0.976 |
| $G R$ | 1.429 | 1.331 | 1.231 | 0.998 |
| Post-GR | 1.616 | 1.422 | 1.278 | 1.047 |
| Covered Workers |  |  |  |  |
| Pre-2000 | 1.156 | 1.122 | 1.035 | 1.082 |
| Post-2000 | 1.168 | 1.129 | 1.056 | 1.064 |
| Pre-GR | 1.106 | 1.071 | 1.016 | 1.040 |
| $G R$ | 1.135 | 1.147 | 1.060 | 1.044 |
| Post-GR | 1.293 | 1.221 | 1.125 | 1.117 |

Notes: The first panel presents measures of earnings inequality for all workers in CPS/ACS, while the second panel presents the same measures for covered workers. The measures of earnings inequality considered are (i) $95^{t h} / 5^{\text {th }}$ : the ratio of the $95^{t h}$ to the $5^{t h}$ percentile; (ii) $90^{t h} / 10^{t h}$ : the ratio of the $90^{\text {th }}$ to the $10^{t h}$ percentile; (iii) $80^{t h} / 20^{t h}$ the ratio of the $80^{t h}$ to the $20^{t h}$ percentile; and (iv) Variance: the variance of log annual earnings. The values in the table are averages before and after 2000: (i) pre-2000: 1995-1999; and (ii) post-2000: 2001-2013. The post-2000 years are further subdivided into three periods: (i) pre-GR: 2001-2007; (ii) GR: 2008-2009; and (iii) post-GR 2010-2013. All measures are 1.00 in 2000.

To see whether it is changes in the top or bottom half of the earnings distribution that are driving these trends, we decompose these ratios around the median, as we did using the two worker frames from LEHD. Notice that since 2000 the ratio of the top percentiles to the median has been gradually increasing for both the all-workers sample and the covered-workers sample (Figures C.3a and C.3c). The bottom of the earnings distribution, however, has evolved differently across these two samples. In the all-workers sample, there has been a substantial rise in inequality (Figure C.3b). Among the covered workers, the rise has been much more mild (Figures C.3d). In fact, the trends in earnings inequality among the covered workers is very similar to those observed among the eligible workers in LEHD both in terms of the correlation of the times series and the magnitude of the changes. However, one notable difference is the change in earnings inequality around the Great Recession. In LEHD, inequality increased dramatically during the Great Recession. In CPS/ACS, inequality actually drops substantially just prior to the onset of the Great Recession. However, these gains are lost during the recession years as inequality quickly increases back to trend. Thus, while both the household surveys and the administrative data highlight the
sensitivity of the bottom of the earnings distribution to the Great Recession, the precise cyclical patterns are not consistent across these two data sources.

Figure C.3: Selected Inequality Measures 1990-2013 for the Top and Bottom of the Earnings Distribution, Relative to 2000 (CPS/ACS)


Notes: Subplots (a) and (b) decompose the 99/1 ratio, the 95/5 ratio, the 90/10 ratio, and the $80 / 20$ for all workers in CPS/ACS relative to 2000 from 1990-2013 relative to the median. Subplot (a) plots the following ratios for the top half of the earnings distribution: (i) P95 to P50: the ratio of the $95^{t h}$ to the $50^{t h}$ percentile; (ii) P90 to P50: the ratio of the $90^{t h}$ to the $50^{t h}$ percentile; and (iii) P80 to $P 50$ the ratio of the $80^{t h}$ to the $50^{t h}$ percentile. Subplot (b) plots the following ratios for the bottom half of the earnings distribution: (i) P50 to P5: the ratio of the 50th to the 5th percentile; (ii) P50 to P10: the ratio of the $50^{t h}$ to the $10^{t h}$ percentile; and (iii) P50 to P20 the ratio of the $50^{\text {th }}$ to the $20^{\text {th }}$ percentile. Subplots (c) and (d) present the same decomposition for covered workers in CPS/ACS relative to 2000 from 1990-2013.

Figure C.4: Comparison of Percentiles in the ACS and LEHD


Notes: This figure plots the 10th, 20th, 80th, and 90th percentiles of the earnings distributions from four samples of the ACS: (i) individuals with positive UI earnings, but no reported ACS earnings (dotted red line); (ii) individuals with positive reported ACS earnings, but no UI earnings (solid red line with diamonds); (iii) individuals with positive reported ACS earnings and positive UI earnings using ACS earnings to compute the earnings distribution (solid green line); and (iv) individuals with positive reported ACS earnings and positive UI earnings using UI earnings to compute the earnings distribution (dotted green line with circles). Subplots (a), (c), (e), and $(g)$ are the percentiles for all workers in ACS. Subplots (b), (d), (f), and (h) are the percentiles for the covered workers in ACS. These are compared to the same percentiles from the eligible workers frame in LEHD (dotted blue line with squares).

## C. 3 Detailed Analysis of Linked Records

To understand where the discrepancies between the administrative and household survey earnings distributions occur, we analyze the individual ACS records from 2005-2013-linking them to LEHD UI records from the eligible-workers frame using a crosswalk between the two person identifiers
developed and maintained by the Census Bureau. This allows us to see how earnings differ among workers who do and do not match to the LEHD individual data. We focus on records from 2005 forward because, for these years, both the ACS and LEHD are fully national.

For an individual in the ACS, there are three types of matches to the eligible-workers frame in the LEHD data: (i) reported earnings are positive in ACS, but UI earnings are zero; (ii) no reported earnings in ACS, but UI earnings are positive; (iii) both ACS reported earnings and UI earnings are positive. We present these match results in Table C.2. The left panel presents the statistics for all individuals in ACS and the right panel presents the same statistics for covered workers in ACS. When we include all individuals in ACS, about $96 \%$ report positive earnings when surveyed in ACS. However, $21 \%$ have no UI earnings. A very small fraction of the individuals in ACS, the remaining $4 \%$, have positive UI earnings but did not report any earnings when surveyed. When we consider only covered ACS workers, all these individuals should report positive earnings in ACS. Of these covered workers, $85 \%$ also have positive UI earnings and $15 \%$ do not match to any UI records.

Table C.2: ACS/UI Match Comparison

| All Individuals |  |  |  | Covered Workers |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $A C S$ | $U I$ | Percent |  | $A C S$ |  | $U I$ | Percent |
| earn $>0$ | earn $>0$ | $75 \%$ |  | earn $>0$ | earn $>0$ | $85 \%$ |  |
| earn $=0$ | earn $>0$ | $4 \%$ |  | earn $=0$ | earn $>0$ | $0 \%$ |  |
| earn $>0$ | earn $=0$ | $21 \%$ |  | earn $>0$ | earn $=0$ | $15 \%$ |  |

Notes: The first row reports the fraction of individuals in ACS that report positive earnings when surveyed and match to the eligible-workers frame in the LEHD data, and therefore have positive UI earnings. The second row reports the fraction of individuals in ACS that do not report earnings when surveyed, but match to the LEHD data, and, therefore, have positive UI earnings. The third row reports the fraction of individuals in ACS who do not match to LEHD data. The left-side panel presents the statistics for all individuals in ACS and the right-side panel presents the same statistics for covered workers in ACS.

Using these matched records, we compare the earnings distribution of four samples of ACS individuals in Figure C.5:

- Individuals with positive UI earnings, but no reported ACS earnings (dotted red line)
- Individuals with positive reported ACS earnings, but no UI earnings (solid red line with diamonds)
- Individuals with positive reported ACS earnings and positive UI earnings using ACS earnings to compute the earnings distribution (solid green line)
- Individuals with positive reported ACS earnings and positive UI earnings using UI earnings to compute the earnings distribution (dotted green line with circles).

We compute these distributions for both all workers and covered workers in the ACS. Note that for covered workers, having only UI earnings is vanishingly rare since all covered workers should report positive earnings in the ACS. The earnings distributions from these samples are compared to the one constructed from the eligible-workers frame in LEHD (dotted blue line with squares). Figure C. 5 plots the $5^{\text {th }}, 50^{\text {th }}$, and $95^{\text {th }}$ percentiles of these various earnings distributions. ${ }^{56}$

For workers whose earnings are both reported in the ACS and found in LEHD (matched workers), the percentiles computed using the ACS earnings are nearly identical to those computed using UI earnings. Specifically, notice that in Figure C. 5 the solid green line and the dotted

[^2]green line with circles are very close to each other in all subplots, and especially at and above the median. The differences in the CPS/ACS percentiles and the LEHD percentiles in Figure C. 2 are, therefore, very unlikely to be due to misreporting in household surveys. Instead, they must be due to differences in the workers who are surveyed and report earnings in the ACS and those who are found in LEHD. Workers who report positive ACS earnings, but do not match to LEHD (ACSonly) tend to have lower earnings than the workers who do match (solid red lines with diamonds in Figure C.5). However, this gap is less pronounced for workers at the top of the earnings distribution for both the all-workers and covered-workers samples in ACS.

While the ACS-only workers do not earn as much as the matched workers, they do earn significantly more than a large portion of workers in LEHD. This means that the LEHD eligibleworkers frame captures more workers in the bottom half of the earnings distribution than the ACS. To see this, notice in Figure C.5b, the $95^{\text {th }}$ percentiles of both the matched and ACS-only samples are nearly identical to the $95^{t h}$ percentile of the eligible-workers earnings distribution in LEHD. However, for the median and lower percentiles, the differences are not trivial. The median matched worker tends to make about $21.5 \%$ more than the median eligible-worker in LEHD ( $\approx \$ 4,770$ ), while the median ACS-only worker makes about $6.4 \%$ less $(\approx \$ 1,417)$. At the bottom, the differences in the $5^{t h}$ percentiles are most stark. A matched worker at the $5^{t h}$ percentile tends to make about 3.22 times as much as an eligible worker at the $5^{\text {th }}$ percentile in LEHD ( $\approx \$ 2,649$ ). Even an ACS-only worker at the $5^{\text {th }}$ percentile makes about 2.44 times as much as a corresponding eligible worker $(\approx \$ 1,459)$. Thus, the left tail of the earnings distribution in ACS is much shorter than the one for eligible workers in the LEHD data, resulting in the LEHD percentiles being less than those computed from household surveys.

Figure C.5: Comparison of Percentiles in the ACS and LEHD


Notes: This figure plots the $5^{\text {th }}, 50^{t h}$, and $95^{t h}$ percentiles of the earnings distributions from four samples of the ACS: (i) individuals with positive UI earnings, but no reported ACS earnings (dotted red line); (ii) individuals with positive reported ACS earnings, but no UI earnings (solid red line with diamonds); (iii) individuals with positive reported ACS earnings and positive UI earnings using ACS earnings to compute the earnings distribution (solid green line); and (iv) individuals with positive reported ACS earnings and positive UI earnings using UI earnings to compute the earnings distribution (dotted green line with circles). Subplots (a), (c), and (e) are the percentiles for all workers in ACS. Subplots (b), (d), and (f) are the percentiles for the covered workers in ACS. These are compared to the same percentiles from the eligible-workers frame in the LEHD data (dotted blue line with squares).

## D Inactive Workers and Inequality

In Section 2.2 we tracked both active and inactive workers in our eligible-workers frame. In Section 3.2 we briefly discussed how the treatment of inactivity affects measures of earnings inequality. This Appendix presents details supporting those analyses and conclusions.

In Appendix B, we excluded inactive workers from the analysis so that we could focus on trends in the ratios of top and bottom percentiles over time. While some inactive workers, given the wages and employment terms on offer, choose to be nonparticipants, others are involuntarily excluded from the labor market. In this section, we present an analysis of how including inactive workers, especially those who were recently employed, affects earnings inequality measures. We begin by analyzing how inactivity has changed in recent years considering comparisons with the employment to population ratio from the CPS/ACS data. Next, we turn our attention to the eligible workers in the LEHD data.

## D. 1 The Employment-to-Population Ratio

If the U.S. labor market tends to stay relatively close to full employment except for brief periods after the start of a recession, the resultant implied rapid employment growth during a recovery should generate a quick increase in the employment-to-population ratio and a quick decline in the unemployment rate to pre-recession levels. However, our results using annual CPS/ACS survey data show a different pattern around the Great Recession.

Figure D. 1 shows the estimated employment-to-population ratio by year from 1990-2013 for all workers (solid blue line) and covered workers (dashed red line) in the CPS/ACS. The NBER identifies three recessions during this period, beginning in the following years: 1990, 2001, and 2008 (December, 2007). Both CPS/ACS series show a dip in 1993 and then a sustained increase until 1999, when the covered-worker sample employment-population ratio begins to decline while the all-workers sample remains relatively flat. Until 1999, the trends for both series are similar, but then the two series diverge, with a decline in the covered workers as a proportion of all workerssuggesting a movement of workers into self-employment. At the beginning of the Great Recession, all three series show a large sustained drop in the employment to population ratio, bottoming out in 2009/2010, with only a mild recovery during the ensuing years. These results suggest that unlike previous recessions, substantial numbers of persons employed prior to the Great Recession did not return to employment even five or more years after the start of the Great Recession. While previous research focused only on employed persons, the large and persistent decrease in the employment-topopulation ratio for all workers and for covered workers only, during and after the Great Recession, argues strongly for an expansion of inequality measures to include at least some inactive but eligible workers.

Figure D.1: Employment to Population Ratio (Household Surveys)


Notes: This figure plots the estimated employment-to-population ratio by year from 1990-2013 for all workers (solid blue line) and covered workers (dashed red line) in CPS/ACS. Estimates based on the authors' calculations from the microdata. These are not the official statistics as released by the BLS from the CPS.

## D. 2 Inactivity-Adjusted Inequality Measures

We estimate three traditional measures of inequality (Gini, Hoover, and Theil), both with and without a category for inactive workers. Deciles of the earnings distribution, estimated as discussed in section 2.4, were used to compute each statistic, with an additional category added for eligible workers with no reported earnings (the inactive category). The earnings value for each person in the inactive category was set to $\$ 1$, a modification necessary to facilitate the consistent calculation of all measures (particularly the Theil index, which uses logarithms). ${ }^{57}$ We create three samples, each with a different set of eligible but inactive workers:

1. All eligible workers each year: This sample assumes all eligible workers are at risk to be employed. Note that this sample is complete and has no dependence on previous years, but the majority of the inactive eligible workers are probably not in the labor force.
2. Active workers and eligible workers most at risk to be employed: This sample includes all active workers and workers not active in the current year, but who were active in at least one of the past 4 years. For years prior to 2008 we do not have complete data for every state. In particular, workers with jobs in MA, DC, AR, NH, and MS will be slightly under-represented (see Table A.1). Some of these workers will have earnings in the previous fours year that we do not observe. An upper bound of the impact of this exclusion might be $5 \%$ of the jobs in 2004, but the actual impact is likely much less since the largest state, MA, entered in 2002Q1, and is therefore missing only two years of history in 2004Q1. In addition, employment in every state is at risk at some point during 2003, the year a worker not employed in 2004 is most likely to have previously been employed.
3. Only active workers: This sample includes only active workers, so no modifications are made to the standard formulas.
[^3]Table D. 1 shows the results. The first panel is for all eligible workers, while the second panel shows results for workers most at risk to be eligible workers. The last panel includes only active eligible workers.

All three of the inequality measures increase substantially as the proportion of eligible workers included in the calculation increases. Not surprisingly, including a large block of workers with only $\$ 1$ of annual earnings greatly increases measured inequality. Comparing our results with another administrative data source, estimates of inequality using SSA data, we find that ours are somewhat larger, although the exact source of the difference is unclear due to coverage differences imposed on the SSA estimation sample. For example, in 2004 the estimated Gini coefficient using a restricted sample of currently eligible SSA recipients is 0.471 , while in our data the estimated Gini is 0.510 (Kopczuk et al., 2010).

Table D.1: Inequality Measures with and without Inactive Workers

| Year | Persons | Gini | Hoover | Theil |
| :---: | :---: | :---: | :---: | :---: |
| All Eligible Workers |  |  |  |  |
| 2004 | $219,763,469$ | 0.696 | 0.538 | 2.379 |
| 2005 | $222,160,089$ | 0.697 | 0.538 | 2.379 |
| 2006 | $224,721,578$ | 0.698 | 0.539 | 2.377 |
| 2007 | $227,553,012$ | 0.699 | 0.540 | 2.386 |
| 2008 | $230,355,015$ | 0.702 | 0.544 | 2.416 |
| 2009 | $232,813,313$ | 0.714 | 0.558 | 2.535 |
| 2010 | $234,304,705$ | 0.720 | 0.564 | 2.576 |
| 2011 | $235,429,997$ | 0.720 | 0.563 | 2.563 |
| 2012 | $236,484,312$ | 0.719 | 0.560 | 2.547 |
| 2013 | $237,816,938$ | 0.716 | 0.558 | 2.532 |

Eligible Workers with Earnings in the Current or at Least One of the Past 4 Years

| 2004 | $164,243,214$ | 0.593 | 0.437 | 1.352 |
| :---: | :---: | :---: | :---: | :---: |
| 2005 | $165,892,505$ | 0.594 | 0.438 | 1.346 |
| 2006 | $167,417,542$ | 0.594 | 0.438 | 1.331 |
| 2007 | $168,988,105$ | 0.595 | 0.439 | 1.327 |
| 2008 | $170,229,709$ | 0.597 | 0.441 | 1.351 |
| 2009 | $170,241,870$ | 0.609 | 0.452 | 1.472 |
| 2010 | $170,617,692$ | 0.616 | 0.458 | 1.509 |
| 2011 | $171,015,983$ | 0.615 | 0.457 | 1.480 |
| 2012 | $170,986,772$ | 0.611 | 0.454 | 1.437 |
| 2013 | $170,735,917$ | 0.604 | 0.448 | 1.387 |
| Eligible Workers with |  |  |  | Earnings |
| in | the Current Year |  |  |  |
| 2004 | $136,562,515$ | 0.510 | 0.369 | 0.529 |
| 2005 | $138,340,770$ | 0.513 | 0.372 | 0.535 |
| 2006 | $140,363,860$ | 0.516 | 0.375 | 0.541 |
| 2007 | $142,034,418$ | 0.519 | 0.378 | 0.546 |
| 2008 | $142,109,590$ | 0.517 | 0.377 | 0.543 |
| 2009 | $137,948,364$ | 0.517 | 0.378 | 0.546 |
| 2010 | $137,345,658$ | 0.522 | 0.382 | 0.557 |
| 2011 | $138,810,297$ | 0.525 | 0.385 | 0.562 |
| 2012 | $140,415,325$ | 0.527 | 0.386 | 0.563 |
| 2013 | $141,665,611$ | 0.523 | 0.384 | 0.555 |

Notes: This table presents traditional measures of inequality (Gini, Hoover, and Theil) for three samples of persons: (i) all eligible workers (top panel), (ii) most at-risk eligible workers (middle panel), and (iii) only active workers (bottom panel).

Figure D. 2 shows the share of eligible workers who are inactive (solid blue line) and the share who are most at risk to be active (dashed red line) relative to the base year 2004. The solid blue
line represents the share of eligible workers not currently working each year-the difference between the number of workers in the first panel of Table D. 1 and the number of workers in the third panel. The dashed red line represents the share of workers most at risk to be active not currently working each year-the difference between the number of workers in the second panel of Table D. 1 and the number of workers in the third panel. The dashed red line is noticeably more responsive to changes in labor demand, suggesting that we chose a reasonable group to represent the workers most at risk to be active. However, a closer look at the source of the decline in the most at risk group (red line) from 2011 forward shows that the decline is due both to the growth in employment during the recovery and a lack of growth in the number of workers most at risk to be active. Many of the at risk workers who had positive earnings just prior to or at the start of the Great Recession have not had positive earnings in the subsequent four years. By 2011, they are dropping out of the at risk group. It is difficult to know the labor force status of these workers due to limitations of administrative data, it does highlight the benefit of having multiple measures of inactive status for the eligible-workers population.

Figure D.2: Inequality Measures


Notes: This figure shows the share of eligible workers who are inactive (solid blue line) and the share who are "at risk" (dashed red line) relative to the base year 2004. In a given year, a person is inactive if that person did not make positive earnings that year. In a given year, a person is "at risk" if that person did not make positive earnings that year, but did make positive earnings sometime in the last four years.

In Section 3 we documented the increase in inequality post-2000 using ratios of various percentiles of the earnings distribution. For the eligible-workers frame, the increase in earnings of the top $20 \%$ relative to the bottom $20 \%$ of earners accelerates during the Great Recession, with annual earnings increases for workers at the 80th percentile and small declines or no increases for at the 20th percentile. The increases for the $99 / 1$ ratio, the $95 / 5$ ratio, and the $90 / 10$ ratio are even larger, with the ratios increasing faster the more extreme the comparison (Figure 3). Here we have taken an alternative approach. Instead of comparing two specific points in the earnings distribution, the portmanteau inequality measures presented here weight the changes occurring across the earnings distribution and combine them to produce a single measure of overall inequality. Each measure uses different weights and combining rules, therefore it is useful to compare each approach.

The relative changes in the Gini coefficients for each of the three samples are presented in Figure D.3. The results for the first two samples are almost identical. The Gini coefficients for
the third sample, only active workers, grow faster before the Great Recession, but do not show the increase in inequality at the start of the Great Recession present for the eligible-workers and at-risk-workers samples. Part of the reason for this difference is that the Gini coefficient is very sensitive to changes in earnings at the top of the distribution. At the beginning of the recession, earnings at the top of the distribution declined or stagnated. In spite of the large number of workers moving from active to inactive status at the beginning of the Great Recession, the Gini coefficient for the active-only sample shows inequality declining, although it does start to climb as earnings growth at the top of the distribution resumes in 2009. In contrast, the Gini coefficients for the all-eligible, and most at-risk samples show increasing inequality at the start of the Great Recession, similar to the $80 / 20$ ratio (also shown in the figure).

Figure D.3: Inequality Measures - Gini Coefficient


Notes: This figure plots the Gini coefficient for three samples of eligible workers: (i) All Eligible Workers: includes active and all inactive workers, (ii) At Risk: includes active workers and inactive workers who made positive earnings sometime in the last four years, and (iii) Active Only: includes only active workers. The ratio of the $80^{t h}$ to the $20^{t h}$ percentile is also plotted for reference.

The results for the Hoover index, shown in Figure D.4, are similar to the Gini coefficient, although the relative increase in inequality during the Great Recession is larger when measured using the Hoover index. The increase before the Great Recession is also larger when using only active workers.

Figure D.4: Inequality Measures - Hoover Index


Notes: This figure plots the Hoover index for three samples of eligible workers: (i) All Eligible Workers: includes active and all inactive workers, (ii) At Risk: includes active workers and inactive workers who made positive earnings sometime in the last four years, and (iii) Active Only: includes only active workers. The ratio of the $80^{t h}$ to the $20^{t h}$ percentile is also plotted for reference.

The final measure we consider is the symmetric Theil index. The results using this measure are shown in Figure D.5. Over the entire period, the Theil measure is more responsive to earnings distribution changes than either the Gini coefficient or the Hoover index, but it is especially responsive to the addition of inactive workers. The relative change in the Theil index computed using all eligible workers (sample one) is almost identical to the $80 / 20$ ratio through 2009, with greater inequality after that 2009 reflecting the slow decline in inactive workers during the recovery. The relative change in the Theil index computed using only the most at-risk workers (sample two) could arguably be viewed as an exaggerated version of the $80 / 20$ ratio. The inclusion of inactive at-risk workers in sample two introduces additional information into the Theil index calculation, magnifying the decline in inequality prior to the Great Recession, the increase during the Great Recession, and the decline during the recovery.

Figure D.5: Inequality Measures - Theil Index


Notes: This figure plots the Theil index for three samples of eligible workers: (i) All Eligible Workers: includes active and all inactive workers, (ii) At Risk: includes active workers and inactive workers who made positive earnings sometime in the last four years, and (iii) Active Only: includes only active workers. The ratio of the $80^{t h}$ to the $20^{t h}$ percentile is also plotted for reference.

Introducing information about inactive but at-risk workers into the calculation of the Gini coefficient and Hoover index changes the trend, but the inequality levels in 2013 are largely the same relative to 2004 using either measure. The Theil index changes in similar ways with the addition of inactive but at-risk workers; however, the Theil index is much more sensitive to both changes in the earnings distribution and the addition of inactive workers. The growth in the Theil index using only active workers is larger than either the Gini or Hoover index. Similar to the Gini and the Hoover indices, by not including inactive workers the Theil index fails to capture the increase in inequality at the start of the Great Recession. Adding inactive workers to the Theil index (sample one), results in a measure similar to the 80/20 ratio through 2009; after 2009 the two measures diverge due to the slow decline in the number inactive workers during the recovery from the Great Recession. The Theil index for the most at-risk workers (sample two) shows the largest changes in inequality.

Although it is unclear which of the adjusted inequality measures correctly weights the inactive workers, it is worthwhile to consider adjusted measures that count at least some of the zero-earning workers as part of any general analysis of changes in earnings inequality.

## E Decomposing Changes in the Earnings Distribution

In Section 4 we presented the evolution of the earnings/inactivity distribution in terms of the year to year flows of workers across different parts of the earnings distributions and into and out of active status.

## E. 1 Worker Flows

Starting in 2005, each year we calculate the change in the number of workers between the current and the previous year for the four earnings/inactivity categories. The year-to-year change in the number of workers in a specific category is driven by changes in the number of workers entering (inflows) and the number of workers leaving (outflows). Specifically, to compute the flows between two employment states, let $A$ and $B$ be arrays of counts for each category in years $t-1$ and $t$, respectively:

$$
\begin{array}{rlrl}
\text { year } t-1: & & A & =\left[\begin{array}{lllll}
a_{0} & a_{1} & a_{2} & a_{3} & a_{4}
\end{array}\right] \\
\text { year } t: & B & =\left[\begin{array}{llllll}
b_{0} & b_{1} & b_{2} & b_{3} & b_{4}
\end{array}\right]
\end{array}
$$

In order to complete the decomposition and capture all possible transitions we must add an additional category, zero, representing workers not eligible to work in one of the two periods, but who are eligible to work in the other. Let $C_{A B}$ be the transition matrix of counts:

$$
C_{A B}=\left[\begin{array}{ccccc}
c_{00} & c_{01} & c_{02} & c_{03} & c_{04} \\
c_{10} & c_{11} & c_{12} & c_{13} & c_{14} \\
c_{20} & c_{21} & c_{22} & c_{23} & c_{24} \\
c_{30} & c_{31} & c_{32} & c_{33} & c_{34} \\
c_{40} & c_{41} & c_{42} & c_{43} & c_{44}
\end{array}\right]
$$

The rows of the transition matrix represent the origin state $(A)$ and the columns represent the destination state $(B)$. For example, $c_{21}$ is the number of workers who were in the bottom $20 \%$ of the overall-earnings distribution in year $t-1$ and transition to the eligible but no reported UI earnings category in year $t$.

To compute the total net inflows into an employment category, we first introduce some notation. Let $\iota$ be a $(5 \times 1)$ column vector of ones. Then:

$$
\begin{aligned}
& C_{A \bullet}=C_{A B} \cdot \iota=\text { outflows }+ \text { stayers } \\
& C \bullet B=C_{A B}^{T} \cdot \iota=\text { inflows }+ \text { stayers }
\end{aligned}
$$

Net inflows into each employment state $\Delta_{A B}^{C}$ are defined as:

$$
\begin{align*}
\Delta_{A B}^{C} & \equiv B-A \\
& =C \bullet B-C_{A \bullet} \\
& =\underbrace{C_{A B}^{T} \cdot \iota}_{\text {inflows }+ \text { stayers }}-\underbrace{C_{A B} \cdot \iota}_{\text {outflows }+ \text { stayers }} \\
& =\underbrace{\left(C_{A B}^{T}-C_{A B}\right) \cdot \iota}_{\text {inflows }- \text { outflows }=\text { net inflows }} \tag{E-1}
\end{align*}
$$

Note the position of the stayers on the main diagonal. When we take the difference between $C_{A B}^{T}$ and $C_{A B}$, the resulting matrix will have zeros on the main diagonal, showing that the stayers do not directly affect the earnings distribution except through changes in average earnings. It should also be noted there is a direct relationship between the number of outflows and the number of stayers. If more workers leave a given category then there will be fewer stayers, ceteris paribus.

Table E. 1 provides descriptive statistics on the individuals in each earnings category. It is expanded in the main text in Table 5, which shows the net change in workers between the previous year and the current year from 2005-2013.

The flows of workers affect the earnings distribution, but the average earnings of each category and the change in average earnings for stayers also affect the change in the earnings distribution. Here we show the complete decomposition of the change in the earnings distribution. Table E. 2 shows the earnings changes we decompose here. Unlike Table 5 in the main text, the decomposition for earnings does not include net inflows into the eligible-worker frame or net inflows to inactive status. As we show below, these flows have no associated earnings and therefore have a weight of zero.

The corresponding earnings transition matrix for a given transition matrix of counts $C_{A B}$ is:

$$
E_{A B}=\left[\begin{array}{ccccc}
0 & 0 & e_{02} & e_{03} & e_{04} \\
0 & 0 & e_{12} & e_{13} & e_{14} \\
e_{20} & e_{21} & e_{22} & e_{23} & e_{24} \\
e_{30} & e_{31} & e_{32} & e_{33} & e_{34} \\
e_{40} & e_{41} & e_{42} & e_{43} & e_{44}
\end{array}\right]
$$

Unlike the transition matrix of counts, each element of the transition matrix of earnings has two associated total earnings values, the total earnings for the workers in period $A$ and the total earnings for those same workers in period $B$. Each element of the earnings transition matrix is an ordered pair of elements. For example, $e_{23}=\left\{e_{23}^{A}, e_{23}^{B}\right\}$ represents the earnings of workers moving from the bottom $20 \%$ to the middle $60 \%$ of the earnings distribution. The first element is the total earnings in the previous period (when each worker is in the bottom $20 \%$ ) and the second element is the total earnings in the current period (when each worker is in the middle $60 \%$. Elements with an ordered pair of two zeros are shown as zeros in the earnings transition matrix.

Applying the net inflow formulas for the counts to the earnings transition matrix,

$$
\begin{equation*}
\Delta_{A B}^{E}=\underbrace{\left(E_{A B}^{\prime}-E_{A B}\right) \cdot i}_{\text {net inflows }} \tag{E-2}
\end{equation*}
$$

and choosing the appropriate earnings value from each tuple, using an $A$ or $B$ superscript to indicate the first or second element chosen, respectively, we have:

$$
\Delta_{A B}^{E}=\left[\begin{array}{ccccc}
(0-0) & (0-0) & \left(e_{20}^{B}-e_{02}^{A}\right) & \left(e_{30}^{B}-e_{03}^{A}\right) & \left(e_{40}^{B}-e_{04}^{A}\right) \\
(0-0) & (0-0) & \left(e_{21}^{B}-e_{12}^{A}\right) & \left(e_{31}^{B}-e_{13}^{A}\right) & \left(e_{41}^{B}-e_{14}^{A}\right) \\
\left(e_{02}^{B}-e_{20}^{A}\right) & \left(e_{12}^{B}-e_{21}^{A}\right) & \left(e_{22}^{B}-e_{22}^{A}\right) & \left(e_{32}^{B}-e_{23}^{A}\right) & \left(e_{42}^{B}-e_{24}^{A}\right) \\
\left(e_{03}^{B}-e_{30}^{A}\right) & \left(e_{13}^{B}-e_{31}^{A}\right) & \left(e_{23}^{B}-e_{32}^{A}\right) & \left(e_{33}^{B}-e_{33}^{A}\right) & \left(e_{43}^{B}-e_{34}^{A}\right) \\
\left(e_{04}^{B}-e_{40}^{A}\right) & \left(e_{14}^{B}-e_{41}^{A}\right) & \left(e_{24}^{B}-e_{42}^{A}\right) & \left(e_{34}^{B}-e_{43}^{A}\right) & \left(e_{44}^{B}-e_{44}^{A}\right)
\end{array}\right] \cdot i .
$$

The sum of each row in the matrix is the net inflow for each category of the earnings/inactivity

Table E.1: Descriptive Statistics by Earnings Categories

|  | 1: Eligible, No Earn | $\begin{array}{r} \text { 2: Bottom } \\ 20 \% \end{array}$ | 3: Middle $60 \%$ | $\begin{aligned} & \text { 4: Top } \\ & 20 \% \end{aligned}$ | Total |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Year | Number of Eligible Workers |  |  |  |  |
| 2004 | 83,200,954 | 27,062,314 | 82,821,341 | 26,678,860 | 219,763,469 |
| 2005 | 83,819,319 | 27,376,301 | 84,079,363 | 26,885,106 | 222,160,089 |
| 2006 | 84,357,718 | 27,598,826 | 84,946,369 | 27,818,665 | 224,721,578 |
| 2007 | 85,518,594 | 27,800,774 | 85,576,064 | 28,657,580 | 227,553,012 |
| 2008 | 88,245,425 | 28,120,283 | 85,548,690 | 28,440,617 | 230,355,015 |
| 2009 | 94,864,949 | 28,119,169 | 81,894,162 | 27,935,033 | 232,813,313 |
| 2010 | 96,959,047 | 28,154,014 | 81,314,722 | 27,876,922 | 234,304,705 |
| 2011 | 96,619,700 | 28,498,111 | 82,538,961 | 27,773,225 | 235,429,997 |
| 2012 | 96,068,987 | 28,269,636 | 83,930,862 | 28,214,827 | 236,484,312 |
| 2013 | 96,151,327 | 28,119,381 | 84,707,469 | 28,838,761 | 237,816,938 |
| Year | Total Earnings (Millions of Real (2000) Dollars) |  |  |  |  |
| 2004 | - | 76,178 | 1,959,201 | 2,351,882 | 4,387,260 |
| 2005 | - | 77,118 | 1,984,925 | 2,407,259 | 4,469,302 |
| 2006 | - | 77,653 | 2,006,111 | 2,529,269 | 4,613,033 |
| 2007 | - | 78,142 | 2,021,497 | 2,636,516 | 4,736,155 |
| 2008 | - | 78,716 | 2,012,397 | 2,576,185 | 4,667,298 |
| 2009 | - | 77,793 | 1,923,326 | 2,488,291 | 4,489,410 |
| 2010 | - | 77,788 | 1,901,588 | 2,524,307 | 4,503,683 |
| 2011 | - | 79,000 | 1,918,544 | 2,542,238 | 4,539,782 |
| 2012 | - | 78,880 | 1,947,808 | 2,625,836 | 4,652,524 |
| 2013 | - | 78,850 | 1,969,953 | 2,657,238 | 4,706,041 |
| Year | Average Earnings per Worker ( $e_{i t}>0$ ) |  |  |  |  |
| 2004 | - | 2,815 | 23,656 | 88,155 | 32,126 |
| 2005 | - | 2,817 | 23,608 | 89,539 | 32,306 |
| 2006 | - | 2,814 | 23,616 | 90,920 | 32,865 |
| 2007 | - | 2,811 | 23,622 | 92,001 | 33,345 |
| 2008 | - | 2,799 | 23,523 | 90,581 | 32,843 |
| 2009 | - | 2,767 | 23,486 | 89,074 | 32,544 |
| 2010 | - | 2,763 | 23,386 | 90,552 | 32,791 |
| 2011 | - | 2,772 | 23,244 | 91,536 | 32,705 |
| 2012 | - | 2,790 | 23,207 | 93,066 | 33,134 |
| 2013 |  | 2,804 | 23,256 | 92,141 | 33,219 |
| Variable | Cumulative Change (2004-2013) |  |  |  |  |
| Number of | 12,950,373 | 1,057,067 | 1,886,128 | 2,159,901 | 18,053,469 |
| Workers | 14.4\% | 3.8\% | 2.3\% | 7.8\% | 7.9\% |
| Total | - | 2,671 | 10,752 | 305,357 | 318,780 |
| Earnings | - | 3.4\% | 0.5\% | 12.2\% | 7.0\% |
| Average | - | -11 | -400 | 3,986 | -175 |
| Earnings | - | -0.4\% | -1.7\% | 4.4\% | -0.9\% |

Notes: The cumulative change in average earnings includes workers with $e_{i t}=0$ (column 1) in the denominator. The overall change for the entire period for workers with $e_{i t}>0$ is $3.3 \%$.

Table E.2: Earnings Associated with Exit from and Entry to Each Earnings Category

| Year | Earn $t-1$ | Earn <br> $t$ | Net <br> Change | Stayers | Outflows | Inflows | InflowsOutflows | Net Change |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Bottom 20\% of the Overall UI Earnings Distribution |  |  |  |  |  |  |  |  |
| 2005 | 76,178 | 77,118 | 939 | 1,625 | 41,849 | 41,164 | -685 | 939 |
| 2006 | 77,118 | 77,653 | 535 | 1,752 | 42,340 | 41,123 | -1,217 | 535 |
| 2007 | 77,653 | 78,142 | 489 | 1,553 | 42,415 | 41,351 | -1,065 | 489 |
| 2008 | 78,142 | 78,716 | 575 | 337 | 41,662 | 41,900 | 237 | 575 |
| 2009 | 78,716 | 77,793 | -923 | -1,193 | 41,681 | 41,951 | 270 | -923 |
| 2010 | 77,793 | 77,788 | -5 | 1,401 | 42,571 | 41,165 | -1,406 | -5 |
| 2011 | 77,788 | 79,000 | 1,212 | 1,948 | 42,359 | 41,622 | -736 | 1,212 |
| 2012 | 79,000 | 78,880 | -120 | 2,680 | 43,350 | 40,550 | -2,800 | -120 |
| 2013 | 78,880 | 78,850 | -30 | 2,637 | 42,914 | 40,246 | -2,668 | -30 |
| Middle $60 \%$ of the Overall UI Earnings Distribution |  |  |  |  |  |  |  |  |
| 2005 | 1,959,201 | 1,984,925 | 25,725 | 37,258 | 278,555 | 267,021 | -11,534 | 25,725 |
| 2006 | 1,984,925 | 2,006,111 | 21,186 | 55,382 | 292,830 | 258,634 | -34,196 | 21,186 |
| 2007 | 2,006,111 | 2,021,497 | 15,386 | 53,012 | 296,600 | 258,975 | -37,626 | 15,386 |
| 2008 | 2,021,497 | 2,012,397 | -9,101 | 15,411 | 288,018 | 263,506 | -24,512 | -9,101 |
| 2009 | 2,012,397 | 1,923,326 | -89,071 | 4,842 | 331,453 | 237,541 | -93,912 | -89,071 |
| 2010 | 1,923,326 | 1,901,588 | -21,738 | 23,095 | 289,271 | 244,438 | -44,833 | -21,738 |
| 2011 | 1,901,588 | 1,918,544 | 16,956 | 22,643 | 263,326 | 257,639 | -5,687 | 16,956 |
| 2012 | 1,918,544 | 1,947,808 | 29,264 | 47,349 | 266,666 | 248,581 | -18,085 | 29,264 |
| 2013 | 1,947,808 | 1,969,953 | 22,144 | 58,469 | 273,520 | 237,196 | -36,324 | 22,144 |
| Top $20 \%$ of the Overall UI Earnings Distribution |  |  |  |  |  |  |  |  |
| 2005 | 2,351,882 | 2,407,259 | 55,377 | 64,813 | 245,494 | 236,058 | -9,436 | 55,377 |
| 2006 | 2,407,259 | 2,529,269 | 122,010 | 88,284 | 227,727 | 261,453 | 33,726 | 122,010 |
| 2007 | 2,529,269 | 2,636,516 | 107,247 | 86,390 | 240,848 | 261,705 | 20,857 | 107,247 |
| 2008 | 2,636,516 | 2,576,185 | -60,330 | -15,291 | 271,995 | 226,955 | -45,040 | -60,330 |
| 2009 | 2,576,185 | 2,488,291 | -87,894 | -22,790 | 291,186 | 226,082 | -65,104 | -87,894 |
| 2010 | 2,488,291 | 2,524,307 | 36,016 | 67,434 | 246,006 | 214,587 | -31,418 | 36,016 |
| 2011 | 2,524,307 | 2,542,238 | 17,931 | 44,185 | 230,451 | 204,197 | -26,254 | 17,931 |
| 2012 | 2,542,238 | 2,625,836 | 83,598 | 78,243 | 214,172 | 219,527 | 5,355 | 83,598 |
| 2013 | 2,625,836 | 2,657,238 | 31,403 | 28,123 | 217,801 | 221,081 | 3,280 | 31,403 |

Notes: The estimates are based on the authors' calculations using transitions into and out of the eligible-workers frame and between categories of the earnings distributions, including inactive workers.
distribution. The sum of the first two rows is zero; each element of the first two rows is zero, there are no earning when not eligible or eligible but inactive. Multiplying each element of the next three rows by a conformable vector of ones we can separate each total earnings value into the product of average earnings and the counts for that value. For example, the net inflows between period $A$ and period $B$ for earnings category two is:

$$
\begin{align*}
\Delta_{A B}^{E 2} & =\left(\bar{e}_{02}^{B} \cdot c_{02}-\bar{e}_{20}^{A} \cdot c_{20}\right)+\left(\bar{e}_{12}^{B} \cdot c_{12}-\bar{e}_{21}^{A} \cdot c_{21}\right) \\
& +\left(\bar{e}_{22}^{B}-\bar{e}_{22}^{A}\right) \cdot c_{22}+\left(\bar{e}_{32}^{B} \cdot c_{32}-\bar{e}_{23}^{A} \cdot c_{23}\right) \\
& +\left(\bar{e}_{42}^{B} \cdot c_{42}-\bar{e}_{24}^{A} \cdot c_{24}\right) \tag{E-3}
\end{align*}
$$

The year-to-year change in the earnings associated with a given part of the earnings distribution is a linear function (weighted sum) of the average earnings and the transition counts. Table E. 2 shows the results, after first grouping the stayers, inflows, and outflows together for the bottom $20 \%$, middle $60 \%$, and top $20 \%$ categories.

The change in earnings reduces to a simple (signed) sum of the counts if the average earnings
is the same for each flow, i.e. $\left(\bar{e}_{2}^{*}=\bar{e}_{02}^{B}=\bar{e}_{20}^{A}=\bar{e}_{12}^{B}=\bar{e}_{21}^{A}=\bar{e}_{22}^{B}=\bar{e}_{22}^{A}=\bar{e}_{32}^{B}=\bar{e}_{23}^{A}=\bar{e}_{42}^{B}=\bar{e}_{24}^{A}\right)$.

$$
\begin{equation*}
\Delta_{A B}^{E 2}=\bar{e}_{2}^{*} \cdot[\underbrace{\left(c_{02}+c_{12}+c_{32}+c_{42}\right)}_{\text {inflows }}-\underbrace{\left(c_{20}+c_{21}+c_{23}+c_{24}\right)}_{\text {outflows }}] \tag{E-4}
\end{equation*}
$$

Although the simple formula will rarely hold in practice, it is useful as the earnings change for each category is now a scaled function of the counts. For the data in this paper a different constant average earnings value for each category does a reasonable job approximating the gross outflows and inflows. However, when using a constant the individual flows are not always scaled correctly since the weights (average earnings) differ substantially in some cases. Even though there are level differences across flows, the average earnings values are for the most part stable over time, allowing the counts to proxy for the change in the earnings distribution over time, once the appropriate scale factor is known for a given flow. The table below shows the average earnings and measures of variability for each of the flows.

Table E.3: Average Earnings and Variability by Transition Type

|  | Flows from Bottom 20\% |  |  |  |  | Flows to Bottom 20\% |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | et_2_2_A | et_2_0 | et_2_1 | et_2_3 | et_2_4 | et_2_2_B | et_0_2 | et_1_2 | et_3_2 | et_4_2 |
| Mean | 2,706 | 2,620 | 2,053 | 3,631 | 3,377 | 2,814 | 2,427 | 2,202 | 3,474 | 2,963 |
| IQR | 22 | 17 | 30 | 48 | 15 | 24 | 244 | 14 | 20 | 67 |
| Minimum | 2,657 | 2,600 | 2,015 | 3,571 | 3,362 | 2,712 | 2,267 | 2,150 | 3,365 | 2,804 |
| Maximum | 2,802 | 2,678 | 2,087 | 3,747 | 3,399 | 2,873 | 2,569 | 2,227 | 3,518 | 3,041 |
|  | Flows from Middle 60\% |  |  |  |  | Flows to Middle 60\% |  |  |  |  |
|  | et_3_3_A | et_3_0 | et_3_1 | et_3_2 | et_3_4 | et_3_3_B | et_0_3 | et_1_3 | et_2_3 | et_4_3 |
| Mean | 23,940 | 18,680 | 16,720 | 14,220 | 37,430 | 24,450 | 11,672 | 15,240 | 12,980 | 35,340 |
| IQR | 160 | 442 | 166 | 181 | 228 | 297 | 1,685 | 429 | 277 | 507 |
| Minimum | 23,540 | 18,110 | 16,560 | 13,850 | 36,910 | 24,220 | 10,370 | 14,950 | 12,540 | 34,260 |
| Maximum | 24,160 | 19,391 | 16,910 | 14,950 | 38,200 | 24,720 | 12,290 | 15,510 | 13,210 | 35,970 |
|  | Flows from Top 20\% |  |  |  |  | Flows to Top 20\% |  |  |  |  |
|  | et_4_4_A | et_4_0 | et_4_1 | et_4_2 | et_4_3 | et_4_4_B | et_0_4 | et_1_4 | et_2_4 | et_3_4 |
| Mean | 94,160 | 113,200 | 107,220 | 80,970 | 60,900 | 96,080 | 96,160 | 94,320 | 73,770 | 57,510 |
| IQR | 1,922 | 3,219 | 8,788 | 2,596 | 1,117 | 2,086 | 6,893 | 1,409 | 3,017 | 152 |
| Minimum | 91,810 | 107,440 | 97,650 | 78,800 | 59,930 | 93,720 | 78,540 | 91,320 | 69,450 | 56,820 |
| Maximum | 96,100 | 118,500 | 117,400 | 82,800 | 61,790 | 98,010 | 108,200 | 98,800 | 81,490 | 57,780 |

Notes: Dominant flows are in bold. The estimates are the weighted annual mean, inter-quartile Range (IQR), minimum, and maximum of the mean annual earnings in each category. Statistics are over nine pairs of years from 2004-2005 to 2012-2013.

Appendix Figures E.1-fig:TE4 repeat the analysis shown in the main text in Figures 10-12.

Figure E.1: Earnings Flows out from and in to the Bottom 20\%


Notes: The estimates are based on the authors' calculations using transitions into and out of categories of the earnings distribution, including inactivity.

Figure E.2: Earnings Flows out from and in to the Middle $60 \%$


Notes: The estimates based on the authors' calculations using transitions into and out of categories of the earnings distribution, including inactivity.

Figure E.3: Earnings Flows out from and in to the Top 20\%


Notes: The estimates are based on the authors' calculations using transitions into and out of categories of the earnings distribution, including inactivity.

## E. 2 AKM Decomposition

We estimate the following AKM model:

$$
\begin{equation*}
\ln y_{i j t}=x_{i t} \beta+\theta_{i}+\psi j+\varepsilon_{i j t} \tag{E-5}
\end{equation*}
$$

where $y_{i j t}$ is $\log$ real annual earnings of person $i$, employed at firm $j$ in year $t ; \theta_{i}$ is individual $i$ 's person effect; $\psi_{j}$ is firm $j$ 's fixed effect; and $x_{i t}$ includes controls for experience, labor force attachment, and aggregate labor market conditions detailed in Table E.4. Estimates of all these controls are in Table E.5.

Table E.4: AKM Model Specification

## Labor-Force Attachment

    weeks by hours categories ( 41 total, 40 hours by \(50-52\) weeks excluded)
    sixq dummies ( 9 total: sixq2-sixq6, sixq_4th, sixq_left, sixq_right, sixq_inter)
    \(\mathbb{1}\{\) female \(\} *[\) sixq dummies]
    \(\mathbb{1}\{\) black \(\} *\) [sixq dummies]
    \(\mathbb{1}\{\) Hispanic \(\} *[\) sixq dummies]
    \(\mathbb{1}\) \{foreign born\} \(*\) [sixq dummies]
    \(\mathbb{1}\{\) foreign born \(\} * \mathbb{1}\{\) female \(\} *\) [sixq dummies]
    \(\mathbb{1}\{\) foreign born \(\} * \mathbb{1}\{\) black \(\} *\) [sixq dummies]
    \(\mathbb{1}\{\) foreign born \(\} * \mathbb{1}\{\) Hispanic \(\} *[\) sixq dummies \(]\)
    ```
Aggregate Labor-Market Conditions
    [ut,\mathbb{1}{\mp@subsup{u}{t}{}>\mp@subsup{u}{t-1}{}}*\mp@subsup{u}{t}{}]
    \mathbb{1}{\mathrm{ female } * [ut , }\mathbb{1}{\mp@subsup{u}{t}{}>\mp@subsup{u}{t-1}{}}*\mp@subsup{u}{t}{}]
    \mathbb{1}\mathrm{ {lack} * [ut, 代{ut > u}\mp@subsup{u}{t-1}{}}*\mp@subsup{u}{t}{}]
    \mathbb{1}\mathrm{ Hispanic } * [ut, }\mathbb{1}{\mp@subsup{u}{t}{}>\mp@subsup{u}{t-1}{}}*\mp@subsup{u}{t}{}]
    \mathbb{1}{foreign born } * [ut,\mathbb{1}{\mp@subsup{u}{t}{}>\mp@subsup{u}{t-1}{}}*\mp@subsup{u}{t}{}]
```




```
    \mathbb{1}{\mathrm{ foreign born } * }\mathbb{1}{\mathrm{ Hispanic } * [ut, 弐{ut > u}\mp@subsup{u}{t-1}{}}*\mp@subsup{u}{t}{}]
```


## Incomplete 2014Q1 Data Controls

[right: indicator for incomplete data in 2014Q1 in one state and DC]
$\mathbb{1}\{$ female $\} *[$ right $]$
$\mathbb{1}\{$ black $\} *[$ right $]$
$\mathbb{1}\{$ Hispanic $\} *[$ right $]$
$\mathbb{1}$ \{foreign born\} $*$ [right]
$\mathbb{1}\{$ foreign born $\} * \mathbb{1}\{$ female $\} *[$ right $]$
$\mathbb{1}\{$ foreign born $\} * \mathbb{1}\{$ black $\} *$ [right $]$
$\mathbb{1}\{$ foreign born $\} * \mathbb{1}\{$ Hispanic $\} *$ [right $]$
Notes: The specification also includes a fixed worker effect for each individual in the eligible-workers frame and a fixed firm effect for each employer in that frame. The AKM estimation occurs only during date regime 4, which is the complete population; however, our labor-force attachment variables require an additional quarter to calculate (2014Q1), which is missing for one state and DC. The "right" variable controls for the case where a sixq variable is set to zero due to data availability instead of actual labor-force attachment.

Table E.5: AKM Model Estimates

| 1 | experience | 0.0973 | 86 | female_sixq5 | 0.0474 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 2 | experience_2 | -0.0379 | 87 | female_sixq6 | 0.0545 |
| 3 | experience_3 | 0.0070 | 88 | female_sixq_4th | -0.0149 |
| 4 | experience_4 | -0.0006 | 89 | female_sixq_left | -0.0028 |
| 5 | female_experience | -0.0122 | 90 | female_sixq_right | 0.0333 |
| 6 | female_experience_2 | 0.0044 | 91 | female_sixq_inter | 0.0265 |
| 7 | female_experience_3 | -0.0005 | 92 | black_sixq2 | 0.0896 |
| 8 | female_experience_4 | 0.0000 | 93 | black_sixq3 | 0.1388 |
| 9 | black_experience | -0.0470 | 94 | black_sixq4 | 0.1584 |
| 10 | black_experience_2 | 0.0211 | 95 | black_sixq5 | 0.1445 |
| 11 | black_experience_3 | -0.0046 | 96 | black_sixq6 | 0.2161 |
| 12 | black_experience_4 | 0.0004 | 97 | black_sixq_4th | 0.0250 |
| 13 | hispanic_experience | -0.0372 | 98 | black_sixq_left | -0.1004 |
| 14 | hispanic_experience_2 | 0.0201 | 99 | black_sixq_right | -0.0761 |
| 15 | hispanic_experience_3 | -0.0049 | 100 | black_sixq_inter | -0.0829 |
| 16 | hispanic_experience_4 | 0.0005 | 101 | hispanic_sixq2 | 0.0962 |
| 17 | fbstat_experience | -0.0424 | 102 | hispanic_sixq3 | 0.1280 |
| 18 | fbstat_experience_2 | 0.0238 | 103 | hispanic_sixq4 | 0.1386 |
| 19 | fbstat_experience_3 | -0.0058 | 104 | hispanic_sixq5 | 0.1101 |
| 20 | fbstat_experience_4 | 0.0005 | 105 | hispanic_sixq6 | 0.1990 |
| 21 | female_fbstat_experience | 0.0007 | 106 | hispanic_sixq_4th | 0.0269 |
| 22 | female_fbstat_experience_2 | 0.0008 | 107 | hispanic_sixq_left | -0.0927 |
| 23 | female_fbstat_experience_3 | -0.0004 | 108 | hispanic_sixq_right | -0.0933 |
| 24 | female_fbstat_experience_4 | 0.0000 | 109 | hispanic_sixq_inter | -0.1123 |
| 25 | black_fbstat_experience | 0.0179 | 110 | fbstat_sixq2 | -0.0108 |
| 26 | black_fbstat_experience_2 | -0.0086 | 111 | fbstat_sixq3 | -0.0361 |
| 27 | black_fbstat_experience_3 | 0.0021 | 112 | fbstat_sixq4 | -0.0602 |
| 28 | black_fbstat_experience_4 | -0.0002 | 113 | fbstat_sixq5 | -0.1155 |
| 29 | hispanic_fbstat_experience | 0.0146 | 114 | fbstat_sixq6 | -0.1533 |
| 30 | hispanic_fbstat_experience_2 | -0.0140 | 115 | fbstat_sixq_4th | 0.0377 |
| 31 | hispanic_fbstat_experience_3 | 0.0044 | 116 | fbstat_sixq_left | 0.0292 |
| 32 | hispanic_fbstat_experience_4 | -0.0005 | 117 | fbstat_sixq_right | 0.0079 |
| 33 | WKSHRS1 | -0.3017 | 118 | fbstat_sixq_inter | 0.0724 |
| 34 | WKSHRS2 | -0.2561 | 119 | female_fbstat_sixq2 | -0.0245 |
| 35 | WKSHRS3 | -0.2044 | 120 | female_fbstat_sixq3 | -0.0390 |
| 36 | WKSHRS4 | -0.1260 | 121 | female_fbstat_sixq4 | -0.0308 |
| 37 | WKSHRS5 | -0.0625 | 122 | female_fbstat_sixq5 | -0.0262 |
| 38 | WKSHRS6 | 0.0782 | 123 | female_fbstat_sixq6 | -0.0374 |
| 39 | WKSHRS7 | 0.1381 | 124 | female_fbstat_sixq_4th | 0.0067 |
| 40 | WKSHRS8 | -0.2907 | 125 | female_fbstat_sixq_left | 0.0100 |
| 41 | WKSHRS9 | -0.1951 | 126 | female_fbstat_sixq_right | -0.0028 |
| 42 | WKSHRS10 | -0.1122 | 127 | female_fbstat_sixq_inter | 0.0041 |
| 43 | WKSHRS11 | -0.0100 | 128 | black_fbstat_sixq2 | 0.0007 |
| 44 | WKSHRS12 | 0.0831 | 129 | black_fbstat_sixq3 | 0.0243 |
| 45 | WKSHRS13 | 0.1570 | 130 | black_fbstat_sixq4 | 0.0403 |
| 46 | WKSHRS14 | 0.1734 | 131 | black_fbstat_sixq5 | 0.0770 |
| 47 | WKSHRS15 | -0.3176 | 132 | black_fbstat_sixq6 | 0.0787 |
| 48 | WKSHRS16 | -0.1633 | 133 | black_fbstat_sixq_4th | -0.0270 |
| 49 | WKSHRS17 | -0.0929 | 134 | black_fbstat_sixq_left | 0.0341 |
| 50 | WKSHRS18 | -0.0090 | 135 | black_fbstat_sixq_right | 0.0483 |
| 51 | WKSHRS19 | 0.0628 | 136 | black_fbstat_sixq_inter | 0.0437 |
| 52 | WKSHRS20 | 0.1167 | 137 | hispanic_fbstat_sixq2 | 0.0099 |
| 53 | WKSHRS21 | 0.1404 | 138 | hispanic_fbstat_sixq3 | -0.0025 |
| 54 | WKSHRS22 | -0.3661 | 139 | hispanic_fbstat_sixq4 | 0.0027 |
| 55 | WKSHRS23 | -0.2028 | 140 | hispanic_fbstat_sixq5 | 0.0529 |
| 56 | WKSHRS24 | -0.1196 | 141 | hispanic_fbstat_sixq6 | 0.0141 |
| 57 | WKSHRS25 | -0.0685 | 142 | hispanic_fbstat_sixq-4th | -0.0252 |
| 58 | WKSHRS26 | -0.0223 | 143 | hispanic_fbstat_sixq_left | 0.0414 |
| 59 | WKSHRS27 | 0.0011 | 144 | hispanic_fbstat_sixq_right | 0.0278 |
| 60 | WKSHRS28 | 0.0161 | 145 | hispanic_fbstat_sixq_inter | 0.0434 |
| 61 | WKSHRS29 | -0.3451 | 146 | urate | -0.0095 |
| 62 | WKSHRS30 | -0.1839 | 147 | urate_up | 0.0017 |
| 63 | WKSHRS31 | -0.0999 | 148 | female_urate | 0.0034 |
| 64 | WKSHRS32 | -0.0550 | 149 | female_urate_up | 0.0006 |
| 65 | WKSHRS33 | -0.0145 | 150 | black_urate | 0.0045 |
| 66 | WKSHRS34 | 0.0028 | 151 | black_urate_up | -0.0001 |
| 67 | WKSHRS35 | 0.0183 | 152 | hispanic_urate | 0.0015 |
| 68 | WKSHRS36 | -0.3237 | 153 | hispanic_urate_up | 0.0005 |
| 69 | WKSHRS37 | -0.1716 | 154 | fbstat_urate | -0.0000 |
| 70 | WKSHRS38 | -0.0929 | 155 | fbstat_urate_up | 0.0003 |
| 71 | WKSHRS39 | -0.0361 | 156 | female_fbstat_urate | -0.0004 |
| 72 | WKSHRS41 | 0.0223 | 157 | female_fbstat_urate_up | -0.0001 |
| 73 | WKSHRS42 | 0.0320 | 158 | black_fbstat_urate | -0.0059 |
| 74 | sixq2 | 1.1170 | 159 | black_fbstat_urate_up | 0.0009 |
| 75 | sixq3 | 2.2170 | 160 | hispanic_fbstat_urate | -0.0032 |
| 76 | sixq4 | 2.7750 | 161 | hispanic_fbstat_urate_up | -0.0003 |
| 77 | sixq5 | 3.2910 | 162 | right | 0.2083 |
| 78 | sixq6 | 3.6920 | 163 | female_right | 0.0319 |
| 79 | sixq-4th | 0.0323 | 164 | black_right | -0.0181 |
| 80 | sixq_left | -0.2940 | 165 | hispanic_right | -0.0051 |
| 81 | sixq_right | -0.1401 | 166 | fbstat_right | 0.0060 |
| 82 | sixq_inter | -0.7029 | 167 | female_fbstat_right | -0.0273 |
| 83 | female_sixq2 | 0.0250 | 168 | black_fbstat_right | 0.0545 |
| 84 | female_sixq3 | 0.0563 | 169 | hispanic_fbstat_right | -0.0139 |
| 85 | female_sixq4 | 0.0544 |  |  |  |

Notes: The table presents the coefficient estimates of all the controls listed in Table E.4. $N=2,014,000,000$, Jobs $=825,900,000$, Persons $=200,700,000$, Firms $=14,650,000$. Intercept $=6.098$, calculated after estimation. The equation includes one person effect for each person and firm effects for all firms, save one. Estimation and identification performed as described in Abowd et al. (4000) All observations in the complete frame, which has universal coverage over the period 2004-2013, were used. Finite population standard errors are zero. The estimates and their associated standard errors have not been corrected for edit, imputation, and post-processing uncertainty.

## E. 3 Analyzing Earnings Inequality Changes Using Only Firm-Type and Non-Firm-Type

In the main text Section 5, we use the AKM decomposition to create firm, non-firm, and skill components of earnings. These components are used to create firm-type, non-firm-type and skilltype bins that we subsequently employ to characterize the worker and firm contributions to changes in earnings inequality.

An earlier version of this paper used the non-firm-type bins in a manner similar to the use of the skill-type bins in the main text. We discuss these results here for each non-firm-type separately. We remind the reader that the non-firm-component contains the effects of changes in the labor-force attachment, macroeconomic conditions, date regime boundaries, and the residual, all of which are excluded from the skill-type in the main text.

Table E. 6 presents outcomes for workers in the bottom bin of the non-firm component. Table E. 7 presents outcomes for workers in the middle bin of the non-firm component distribution. Table E. 8 presents outcomes for workers in the top bin of the non-firm component distribution.

The tables were created as follows. They are based on classifying workers in the previous year, i.e., year $t-1$. Beginning in 2004 and ending in 2012, for every year that an eligible worker has positive earnings a single observation is added to one of the three tables. The appropriate table classification for each observation is determined by the non-firm type for that year, which can vary over time as workers accumulate experience, work more/less hours during the quarter, receive a positive or negative aggregate demand shock, or have a large positive or negative residual. Within each non-firm type, the earnings record is further classified based on the firm type, resulting in each earnings observation being classified into one of nine possible cells. ${ }^{58}$ Within each of the non-firmtype $\times$ firm-type cells, we break down the results by the three possible overall-earnings outcomes (bottom, middle and top). There are, thus, twenty-seven cells for which we present information on the number of workers, average earnings for the previous year $(t-1)$, and average earnings for the current year $(t)$ by flow type. ${ }^{59}$

To fix ideas, we will take a detailed look at two rows in Table E.6. To be recorded in this table, the person must have been in the bottom bin (lowest bin) of the non-firm-component distribution in the "Previous Year," i.e., $t-1$.

Consider the first row of the table. This row is in the panel labeled "Bottom Firm," indicating that this person is employed at a firm in the bottom bin of the firm component distribution in $t-1$. Persons in this row are also in the bottom bin of the overall-earnings distribution in year $t-1$, and the share of such persons (relative to those in the middle or top of the overall-earnings distribution) is 1.000 , indicating that no person in the bottom of the non-firm component distribution and the bottom of the firm component distribution is employed outside of the bottom bin of the overallearnings distribution. The flow labeled " $2 \_0$ " is the movement from the bottom of the overallearnings distribution (bin 2) to ineligible; that is, this is the flow out of the frame for persons at the bottom of the overall-earnings distribution. There were, on average, 59,554 such persons each "previous year" $(t-1)$. They represent $0.7 \%$ of the flows from bin 2 of the overall-earnings distribution. Average earnings in $t-1$ were $\$ 1,381$ of which $-\$ 1,463$ are attributed to the firm component of our decomposition and $\$ 2,844$ are attributed to the non-firm component of our decomposition. There were no earnings in the current year $(t)$, because the person has moved out of the frame in $t$.

Next, consider the row labeled "Middle" in the "All Earnings" column in the "Middle Firm" panel with a " $3 \_3$ " flow. All persons in this row were, once again, at the bottom of the non-

[^4]firm component distribution in year $t-1$. Of all such persons, $56 \%$ are employed by a firm in the middle of the firm component distribution. Of all persons at the bottom of the non-firm component distribution and in the middle of the firm component distribution in year $t-1$, the proportion 0.159 were in the middle of the overall-earnings distribution. Among such persons, the "3_3" row shows those who remain in the middle of the overall-earnings distribution in the current year, $t$, of which there were, on average, $1,470,659$ in the 9 pairs of years for which the table was constructed. Those who stayed in the middle of the overall-earnings distribution represented $58.9 \%$ of all persons who were in the middle of the overall-earnings distribution in year $t-1$, on average. In year $t-1$, their earnings averaged $\$ 8,498$ of which $\$ 2,180$ is attributed to the firm component in our decomposition and $\$ 6,318$ is attributed to the non-firm component. In the current year, year $t$, average earnings were $\$ 15,688$ of which $\$ 3,555$ is associated with the firm component and $\$ 12,132$ is associated with the non-firm component.

We use these tables to investigate worker sorting directly by looking at the interaction of the non-firm and firm type for each worker-year-earnings observation. If there were no sorting, the distribution of earnings observations across firm types would be similar for all three tables, because outcomes would be unaffected by which part of the non-firm component distribution an individual occupied, given his place in the overall-earnings distribution. This hypothesis is clearly not supported by the data. For example, again using Table E. 6 showing the bottom of the non-firmtype distribution, about $33 \%$ of the earnings observations are in firms at the bottom of the firm-type distribution, $56 \%$ are in firms of the middle type, and only $11 \%$ are in top firms. In comparison, Tables E. 7 and E. 8 show that persons in the middle and top of the non-firm type distributions are much less likely to be employed at firms in the bottom type ( $14 \%$ and $24 \%$ respectively), and much more likely to be employed at top firms ( $23 \%$ and $20 \%$ respectively). Interestingly, the relationship is not monotonic; workers in the middle are more likely to work at both middle and top firms relative to top workers.

Next, we focus on each non-firm-type in turn, starting with the earnings observations for workers in the bottom of the non-firm component distribution in Table E.6. For workers at the bottom of the non-firm-component distribution, working at a high-paying firm has two advantages: higher earnings than otherwise and a greater chance of moving to a higher bin in the overall-earnings distribution. For example, a worker at the bottom of the non-firm-component and overall-earnings distributions has a probability of moving to the middle of the overall-earnings distribution of $18 \%$ at a low paying firm, $29.5 \%$ at a middle paying firm, and $27.5 \%$ at a high paying firm. Prior to the transition the average worker with a low non-firm component at a low-, middle- and high-paying firm earns $\$ 2,084, \$ 3,556$, and $\$ 3,806$, respectively. ${ }^{60}$ After the transition the average worker at a low-, middle- and high-paying firm earns $\$ 11,640, \$ 13,752$, and $\$ 18,017$, respectively. Most of the additional increase in earnings for workers employed at a top-paying employer in the previous year is due to working at a top-paying employer in the next year.

[^5]Table E.6: Earnings Associated with Flows by Firm Bin for Persons in the Bottom-Type Non-Firm Category)

|  |  |  |  |  | Previous Year |  |  | Current Year |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| All <br> Earnings | Share | Flow | Average Count | Pct | Total | Firm | NonFirm | Total | Firm | NonFirm |
| Bottom Firm (33\%) |  |  |  |  |  |  |  |  |  |  |
| Bottom | 1.000 | 2_0 | 59,554 | 0.7\% | 1,381 | -1,463 | 2,844 | - | - | - |
|  |  | 2_1 | 2,441,375 | 26.8\% | 1,102 | -1,099 | 2,201 |  |  |  |
|  |  | 2_2 | 4,962,828 | 54.5\% | 1,635 | -1,588 | 3,223 | 2,466 | -1,981 | 4,447 |
|  |  | 2_3 | 1,641,446 | 18.0\% | 2,084 | -1,738 | 3,823 | 11,640 | -4,017 | 15,657 |
|  |  | 2_4 | 8,640 | 0.1\% | 1,513 | -1,558 | 3,071 | 78,157 | 8,958 | 69,199 |
| Middle | 0.000 | 3_0 | 0 | 0.0\% | - | - | - | - | - |  |
|  |  | 3_1 | 0 | $0.0 \%$ | _ | $\ldots$ | _ | - | - | - |
|  |  | 3_2 | 0 | 0.0\% | - | _ | - | - | - | - |
|  |  | 3_3 | 0 | 0.0\% | - | - | - | - | - | - |
|  |  | 3_4 | 0 | 0.0\% | - | - | - |  | - |  |
| Top | 0.000 | 4_0 | 0 | 0.0\% | $\qquad$ | $[$ | - | - | - | - |
|  |  | 4.1 | 0 | $0.0 \%$ | $\qquad$ | $\qquad$ | $\qquad$ | - | - | - |
|  |  | 4.2 | 0 | 0.0\% | $\qquad$ | _ | $\square$ | - | - | - |
|  |  | 4_3 | 0 | 0.0\% | $\qquad$ | _ | - | - | - | - |
|  |  | 4 -4 | 0 | 0.0\% | - | - | - | - | - | - |
| Middle Firm (56\%) |  |  |  |  |  |  |  |  |  |  |
| Bottom | 0.841 | 2_0 | $116,724$ | 0.9\% | $2,660$ | 72 | $2,588$ | - | - | - |
|  |  | 2_1 | $3,613,606$ | 27.3\% | $2,289$ | 42 | $2,247$ |  | - | - |
|  |  | 2_2 | 5,565,538 | 42.0\% | 2,784 | -145 | 2,929 | 2,799 | -594 | 3,392 |
|  |  | 2_3 | 3,911,555 | 29.5\% | 3,556 | -27 | 3,583 | 13,752 | 561 | 13,191 |
|  |  | 2_4 | 36,073 | 0.3\% | 3,392 | 469 | 2,923 | 69,402 | 19,672 | 49,730 |
| Middle | 0.159 | 3_0 | $21,191$ | $0.8 \%$ | $8,381$ | $2,189$ | $6,193$ | - | - | - |
|  |  | 3_1 | 428,729 | 17.2\% | $8,384$ | $2,249$ | $6,135$ |  | - |  |
|  |  | 3.2 | 554,068 | 22.2\% | 8,153 | 1,893 | 6,260 | 3,321 | -14 | 3,336 |
|  |  | 3_3 | 1,470,659 | 58.9\% | 8,498 | 2,180 | 6,318 | 15,688 | 3,555 | 12,132 |
|  |  | 3-4 | 21,549 | 0.9\% | 8,955 | 2,823 | 6,132 | 64,566 | 22,919 | 41,647 |
| Top | 0.000 | 4_0 | 0 | 0.0\% | - | - | - | $\square$ | - | - |
|  |  | 4_1 | 0 | 0.0\% | - | - | - | - | - | - |
|  |  | $4 \_2$ | 0 | 0.0\% | - | - | - | - | - | - |
|  |  | 4_3 | 0 | $0.0 \%$ | $\qquad$ | $\qquad$ | - | - | - | - |
|  |  | 4-4 | 0 | 0.0\% | $\ldots$ | - |  | - | - | - |
|  | Top Firm (11\%) |  |  |  |  |  |  |  |  |  |
| Bottom | 0.396 | 2_0 | 17,420 | 1.4\% | 2,913 | 1,598 | 1,314 | $\square$ | $\square$ | - |
|  |  | 2_1 | 469,324 | 38.3\% | 2,758 | 1,515 | 1,243 | - | - | - |
|  |  | 2 _2 | 377,303 | 30.8\% | 3,174 | 1,740 | 1,433 | 2,905 | 806 | 2,099 |
|  |  | 2_3 | 337,787 | 27.5\% | 3,806 | 2,034 | 1,771 | 18,017 | 7,482 | 10,535 |
|  |  | 2_4 | 24,607 | 2.0\% | 3,701 | 2,058 | 1,642 | 76,278 | 41,316 | 34,962 |
| Middle | 0.596 | 3_0 | 16,910 | 0.9\% | 12,121 | 7,299 | 4,822 | - | - | - |
|  |  | 3_1 | 375,155 | 20.3\% | 12,082 | 7,280 | 4,802 | - | - | - |
|  |  | 3_2 | 243,668 | 13.2\% | 11,573 | 6,756 | 4,817 | 3,134 | 774 | 2,360 |
|  |  | 3_3 | 1,108,356 | 60.0\% | 13,551 | 8,484 | 5,067 | 20,030 | 10,749 | 9,281 |
|  |  | 3_4 | 102,240 | 5.5\% | 15,785 | 10,529 | 5,255 | 70,117 | 42,636 | 27,481 |
| Top | 0.008 | 4_0 | 172 | 0.7\% | 97,790 | 93,786 | 4,005 | - | - | - |
|  |  | 4_1 | 1,924 | 7.6\% | 96,408 | 92,599 | 3,809 | - | - | - |
|  |  | 4_2 | 498 | 2.0\% | 93,690 | 89,588 | 4,103 | 2,376 | -53 | 2,430 |
|  |  | 4_3 | 4,217 | 16.7\% | 65,091 | 60,669 | 4,422 | 32,200 | 28,313 | 3,887 |
|  |  | 4-4 | 18,478 | 73.1\% | 108,698 | 104,482 | 4,216 | 117,522 | 110,839 | 6,683 |

Notes: The estimates are based on the authors' calculations using the nine paired years from 2004-2005 to 2012-2013. The first year in the pair is the "previous year" in the table, and the second year in the pair is the "current year." Bins associated with the flows are " 0 " inflow/outflow from the eligible-workers frame, " 1 " inactive but eligible, " 2 " bottom of the overall-earnings distribution, " 3 " middle of the overall-earnings distribution, and " 4 " top of the overall-earnings distribution. "Average count" is the average number of persons in the row during the year labeled "previous year" $(t-1)$. Pct is the percent distribution of transitions for all persons who started the year in the same overall-earnings distribution bin. For "Previous Year" and "Current Year," "Total" is the average real earnings in 2000 dollars, "Firm" is the average real earnings associated with the firm component in our decomposition, and "Non-Firm" is the average real earnings associated with the non-firm component in our decomposition.

The vast majority ( $63 \%$ ) of workers in the middle of the non-firm component distribution are employed at middle-paying firms, as Table E. 7 shows. The next most prevalent outcomes for such workers are employment at top- and bottom-paying firms, $23 \%$ and $14 \%$ respectively. Similar to workers at the bottom of the non-firm type distribution, who also generally appear at the bottom of the overall-earnings distribution ( $84 \%$ ) when employed by middle-paying firms, the majority of workers in the middle of the non-firm type distribution, no matter the firm type, are in the middle of the overall-earnings distribution. However, in spite of the majority of earnings observations being in the middle of the overall-earnings distribution, average earnings differ substantially across firm types. A middle-type worker in bin 3 of the overall-earnings distribution who stays in bin 3 of the overall distribution (a "3_3" flow) at a bottom-type firm has $t-1$ earnings of $\$ 12,356$, a middle-type worker in a middle-type firm has $t-1$ earnings of $\$ 22,978$, and a middle-type worker at a top firm has earnings of $\$ 32,321$. Most of the difference is due to a larger firm effect, although the non-firm component declines somewhat as a middle-type person is found in increasing firm types, giving back some of the gains. Similar to bottom-type workers, one of the additional benefits of finding employment at a high-paying firm is a greater probability of moving to the top of the earnings distribution ( $0.2 \%$ vs. $2.7 \%$ vs. $11.9 \%$ in rows 5,25 , and 40 , respectively).

Similar to bottom and middle non-firm-type workers, Table E. 8 shows that about $64 \%$ of top non-firm type workers are also in the top of the overall-earnings distribution, but there is also a substantial minority in the middle. The differences between working at a middle- compared to a bottom-type firm are relatively small, but the gains from working at a top-type firm are very large. Somewhat surprisingly perhaps, there are a relatively large number of top-type workers at bottom- and middle-type firms. On average, these workers, especially in the middle, are employed at worse-paying firms than middle non-firm type workers.

Table E.7: Earnings Associated with Flows by Firm Bin for Persons in the Middle-Type Non-Firm Category

|  |  |  |  |  | Previous Year |  |  | Current Year |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| All <br> Earnings | Share | Flow | Average Count | Pct | Total | Firm | Non- <br> Firm | Total | Firm | Non- <br> Firm |
| Bottom Firm (14\%) |  |  |  |  |  |  |  |  |  |  |
| Bottom | 0.313 | 2_0 | 26,241 | 0.7\% | 4,596 | -7,583 | 12,179 | - | - | - |
|  |  | 2_1 | 352,516 | 9.4\% | 4,627 | -6,141 | 10,768 |  | - |  |
|  |  | 2_2 | 2,005,303 | 53.6\% | 4,676 | -6,627 | 11,303 | 3,590 | -4,861 | 8,452 |
|  |  | 2_3 | 1,352,780 | 36.2\% | 5,199 | -5,937 | 11,137 | 11,278 | -7,815 | 19,093 |
|  |  | 2.4 | 4,255 | 0.1\% | 4,649 | -7,524 | 12,173 | 79,652 | -16,894 | 96,546 |
| Middle | 0.687 | 3_0 | 39,797 | 0.5\% | 11,198 | -9,855 | 21,053 | - | - | - |
|  |  | 3_1 | 312,400 | 3.8\% | 10,551 | -8,857 | 19,408 |  | - |  |
|  |  | 3_2 | 1,331,161 | 16.2\% | 9,726 | -8,550 | 18,275 | 3,798 | -3,326 | 7,124 |
|  |  | 3_3 | 6,493,717 | 79.2\% | 12,356 | -9,762 | 22,118 | 14,200 | -9,400 | $23,600$ |
|  |  | 3.4 | $18,706$ | 0.2\% | 14,081 | -10,636 | 24,717 | 70,391 | -30,622 | $101,013$ |
| Top | 0.000 | 4_0 | 0 | 0.0\% | - |  | - | $\square$ | - |  |
|  |  | 4.1 | 0 | $0.0 \%$ | $\qquad$ | $\qquad$ | $\qquad$ | - | - | - |
|  |  | 4_2 | 0 | 0.0\% | $\qquad$ | $\qquad$ | _ | - | - | - |
|  |  | 4_3 | 0 | 0.0\% | $\qquad$ | $\qquad$ | - | - | - |  |
|  |  | 4-4 | 0 | 0.0\% | - | - | - | - | - | - |
| Middle Firm (63\%) |  |  |  |  |  |  |  |  |  |  |
| Bottom | 0.010 | 2_0 | $3,160$ | 0.6\% | $6,108$ | $-2,046$ | $8,154$ | - | - | - |
|  |  | 2.1 | $56,529$ | 10.7\% | $6,093$ | $-2,039$ | $8,132$ | - | - |  |
|  |  | 2_2 | 211,504 | 39.9\% | 6,079 | -2,081 | 8,160 | 3,753 | -1,489 | 5,242 |
|  |  | 2_3 | 257,664 | 48.7\% | 6,122 | -2,062 | 8,185 | 12,008 | -2,577 | 14,585 |
|  |  | 2_4 | 730 | 0.1\% | 6,121 | -2,031 | 8,152 | 68,020 | 5,440 | 62,580 |
| Middle | 0.958 | 3_0 | $170,775$ | 0.3\% | $18,829$ | $1,971$ |  | - | - | - |
|  |  | 3-1 | 1,789,911 | 3.6\% | $16,909$ | $2,210$ | $14,699$ |  |  |  |
|  |  | 3.2 | 3,467,732 | 6.9\% | 15,078 | 884 | 14,194 | 3,439 | -520 | 3,958 |
|  |  | 3_3 | 43,259,502 | 86.4\% | 22,978 | 3,475 | 19,503 | 23,517 | 3,506 | 20,012 |
|  |  | 3_4 | 1,370,036 | 2.7\% | 35,902 | 10,130 | 25,772 | 57,122 | 16,550 | 40,572 |
| Top | 0.031 | 4_0 | 2,532 | 0.2\% | 51,159 | 18,825 | 32,335 | $\square$ | - | - |
|  |  | 4.1 | 17,159 | 1.0\% | 51,191 | 19,006 | 32,185 | - | - |  |
|  |  | 4_2 | 13,212 | 0.8\% | 50,902 | 18,745 | 32,156 | 3,202 | 569 | 2,632 |
|  |  | 4_3 | 437,317 | 26.6\% | 49,933 | 17,999 | 31,934 | 37,792 | 13,081 | 24,711 |
|  |  | 4_4 | 1,174,019 | 71.4\% | 51,694 | 19,249 | 32,445 | 55,357 | 20,583 | 34,775 |
| Top Firm (23\%) |  |  |  |  |  |  |  |  |  |  |
| Bottom | 0.000 | 2_0 | 0 | 0.0\% | - | - | - | - | - | - |
|  |  | 2.1 | 0 | 0.0\% | - | - | - | - | - | - |
|  |  | 2_2 | 0 | 0.0\% | - | - | - | - | - | - |
|  |  | 2_3 | 0 | 0.0\% | - | $\qquad$ | - | - | - | - |
|  |  | 2-4 | 0 | 0.0\% | - |  | - | - | - | $\underline{+}$ |
| Middle | 0.569 | 3_0 | 30,130 | 0.3\% | 29,438 | 15,459 | 13,980 | - | - | - |
|  |  | 3.1 | 445,548 | 4.0\% | 27,758 | 14,834 | 12,923 | - | - | - |
|  |  | 3-2 | 343,349 | 3.1\% | 27,352 | 14,186 | 13,166 | 3,111 | 789 | 2,322 |
|  |  | 3_3 | $8,891,952$ | 80.7\% | 32,321 | 16,549 | 15,772 | 31,657 | 15,654 | 16,003 |
|  |  | 3.4 | 1,306,028 | 11.9\% | 38,938 | 20,918 | 18,021 | 58,297 | 31,112 | 27,185 |
| Top | 0.431 | 4_0 | $12,388$ | 0.1\% | 64,410 | 38,198 | 26,213 | - | - | - |
|  |  | 4.1 | $129,141$ | 1.5\% | 64,268 | 39,009 | 25,258 | - | - | - |
|  |  | 4_2 | $69,540$ | 0.8\% | 61,384 | 35,782 | 25,602 | 2,939 | 974 | 1,965 |
|  |  | 4_3 | $1,055,443$ | 12.6\% | 56,162 | 31,142 | 25,020 | 34,895 | 17,929 | 16,965 |
|  |  | 4-4 | 7,085,455 | 84.8\% | 64,675 | 37,688 | 26,987 | 68,632 | 39,649 | 28,983 |

Notes: The estimates are based on the authors' calculations using the nine paired years from 2004-2005 to 2012-2013. The first year in the pair is the "previous year" in the table, and the second year in the pair is the "current year." Bins associated with the flows are " 0 " inflow/outflow from the eligible-workers frame, " 1 " inactive but eligible, " 2 " bottom of the overall-earnings distribution, " 3 " middle of the overall-earnings distribution, and " 4 " top of the overall-earnings distribution. "Average count" is the average number of persons in the row during the year labeled "previous year" $(t-1)$. Pct is the percent distribution of transitions for all persons who started the year in the same overall-earnings distribution bin. For Previous Year" and "Current Year," "Total" is the average real earnings in 2000 dollars, "Firm" is the average real earnings associated with the firm component in our decomposition, and "Non-Firm" is the average real earnings associated with the non-firm component in our decomposition.

Table E.8: Earnings Associated with Flows by Firm Bin for Persons in the High-Type Non-Firm Category

|  |  |  |  |  | Previous Year |  |  | Current Year |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| All <br> Earnings | Share | Flow | Average Count | Pct | Total | Firm | NonFirm | Total | Firm | Non- <br> Firm |
| Bottom Firm (24\%) |  |  |  |  |  |  |  |  |  |  |
| Bottom | 0.005 | 2_0 | 679 | 2.0\% | 4,353 | -62,991 | 67,344 | - | - | - |
|  |  | 2_1 | 2,485 | 7.2\% | 4,316 | -64,774 | 69,090 |  |  |  |
|  |  | 2_2 | 23,108 | 67.2\% | 4,275 | -64,476 | 68,750 | 3,484 | -51,189 | 54,673 |
|  |  | 2_3 | 7,784 | 22.7\% | 5,099 | -60,143 | 65,241 | 12,205 | -94,469 | 106,674 |
|  |  | 2_4 | 307 | 0.9\% | 3,928 | -91,664 | 95,592 | 193,962 | -307,444 | 501,406 |
| Middle | 0.768 | 3_0 | 16,791 | 0.3\% | 24,421 | -41,367 | 65,788 | - | - | - |
|  |  | 3_1 | $93,523$ | $1.8 \%$ | $25,157$ | $-38,303$ | $63,460$ |  |  |  |
|  |  | 3_2 | $162,321$ | $3.2 \%$ | $21,323$ | $-37,554$ | $58,877$ | 3,476 | -8,143 | 11,619 |
|  |  | 3_3 | $4,657,816$ | 90.6\% | $28,233$ | -34,400 | $62,633$ | 27,573 | -31,979 | 59,552 |
|  |  | 3_4 | 211,635 | 4.1\% | 39,258 | -46,667 | 85,925 | 56,165 | -60,567 | 116,732 |
| Top | 0.227 | 4_0 | 4,499 | 0.3\% | 92,980 | -133,477 | 226,457 | - | - | - |
|  |  | 4_1 | 15,036 | 1.0\% | 93,061 | -171,801 | 264,862 |  |  |  |
|  |  | 4_2 | 9,282 | 0.6\% | 80,095 | -133,162 | 213,257 | 3,008 | -6,393 | 9,401 |
|  |  | 4.3 | 192,250 | 12.6\% | 58,951 | -75,598 | 134,549 | 36,197 | -41,533 | 77,730 |
|  |  | 4_4 | 1,299,404 | 85.5\% | 79,502 | -109,491 | 188,992 | 80,704 | -107,069 | 187,773 |
| Middle Firm (56\%) |  |  |  |  |  |  |  |  |  |  |
| Bottom | 0.000 | 2_0 | 0 | 0.0\% | - | - | - | - | - | - |
|  |  | 2_1 | 0 | 0.0\% | - | - | - | - | - | - |
|  |  | 2_2 | 0 | 0.0\% | - | - | - | - | - | - |
|  |  | 2_3 | 0 | 0.0\% | - | - | - | - | - | - |
|  |  | 2 -4 | 0 | 0.0\% | - | - | - | - | - | - |
| Middle | 0.310 | 3_0 | 9,579 | 0.2\% | 37,680 | -5,362 | 43,042 | - | - | - |
|  |  | 3_1 | 58,065 | 1.2\% | 37,365 | -5,593 | 42,957 | - | - |  |
|  |  | 3_2 | 61,221 | 1.3\% | 36,524 | -5,940 | 42,464 | 3,195 | -1,035 | 4,230 |
|  |  | 3_3 | 4,173,530 | 85.7\% | 37,519 | -5,487 | 43,005 | 35,402 | -5,028 | 40,430 |
|  |  | 3-4 | 570,086 | 11.7\% | 42,684 | -2,446 | 45,130 | 54,161 | -1,779 | 55,940 |
| Top | 0.690 | 4_0 | 25,565 | 0.2\% | 111,359 | 17,448 | 93,911 | - | - | - |
|  |  | 4_1 | 103,830 | 1.0\% | 96,076 | 16,375 | 79,701 | - | - |  |
|  |  | $4 \_2$ | 60,924 | 0.6\% | 76,673 | 11,055 | 65,618 | 3,037 | -346 | 3,383 |
|  |  | 4_3 | $1,192,613$ | 11.0\% | 61,122 | $7,031$ | 54,091 | 35,535 | 2,971 | 32,564 |
|  |  | 4_4 | 9,463,943 | 87.3\% | 88,559 | 15,037 | 73,522 | 89,570 | 15,434 | 74,136 |
| Top Firm (20\%) |  |  |  |  |  |  |  |  |  |  |
| Bottom | 0.000 |  | 0 |  |  |  |  | - | - | - |
|  |  | 2_1 | 0 | $0.0 \%$ | $\ldots$ | $\qquad$ | $\qquad$ | $\square$ | - | - |
|  |  | 2_2 | 0 | $0.0 \%$ | $\qquad$ | $\qquad$ | - | - | - | - |
|  |  | $2 \_3$ | 0 | $0.0 \%$ | $\qquad$ | $\qquad$ | - | - | - | - |
|  |  | 2_4 | 0 | 0.0\% | - |  | - |  |  |  |
| Middle | 0.000 | 3_0 |  |  | [ | - | - | - | - | - |
|  |  | 3_1 | 0 | $0.0 \%$ | _ | _ | - | - | - | - |
|  |  | 3_2 | 0 | $0.0 \%$ | _ | - | - | - | - | - |
|  |  | 3_3 | 0 | 0.0\% | - |  | - | - | - | - |
|  |  | 3-4 | 0 | 0.0\% | - |  |  |  |  |  |
| Top | 1.000 | 4_0 | $9,962$ | 0.2\% |  |  | $88,225$ | - | - | - |
|  |  | 4_1 | $73,693$ | 1.4\% | $214,392$ | $127,870$ | $86,521$ | - | - | - |
|  |  | 4_2 | 27,036 | 0.5\% | $155,772$ | $88,384$ | $67,388$ | 2,733 | 610 | 2,123 |
|  |  | 4_3 | 163,477 | 3.0\% | 121,408 | 66,733 | 54,675 | 29,337 | 13,458 | 15,879 |
|  |  | 4_4 | 5,145,974 | 94.9\% | 158,370 | 90,525 | 67,845 | 158,948 | 90,228 | 68,720 |

Notes: The estimates are based on the authors' calculations using the nine paired years from 2004-2005 to 2012-2013. The first year in the pair is the "previous year" in the table, and the second year in the pair is the "current year." Bins associated with the flows are " 0 " inflow/outflow from the eligible-workers frame, " 1 " inactive but eligible, " 2 " bottom of the overall-earnings distribution, " 3 " middle of the overall-earnings distribution, and " 4 " top of the overall-earnings distribution. "Average count" is the average number of persons in the row during the year labeled "previous year" $(t-1)$. Pct is the percent distribution of transitions for all persons who started the year in the same overall-earnings distribution bin. For Previous Year" and "Current Year," "Total" is the average real earnings in 2000 dollars, "Firm" is the average real earnings associated with the firm component in our decomposition, and "Non-Firm" is the average real earnings associated with the non-firm component in our decomposition.


[^0]:    ${ }^{53}$ See http://www.bls.gov/opub/hom/pdf/homch5.pdf for more information
    ${ }^{54}$ The BLS QCEW estimates account for about $95 \%$ of the BEA wage and salary component of the NIPA tables. See http://www.bea.gov/faq/index.cfm?faq_id=104 for more information.

[^1]:    ${ }^{55}$ See Spletzer (2014) for a very similar comparison.

[^2]:    ${ }^{56}$ For similar comparisons of the 10 th, 20th, 80 th, and 90 th percentiles, see Figure C.4.

[^3]:    ${ }^{57}$ The results for the Gini and Hoover measures using $\$ 0$ show very small differences in levels and identical trends compared with setting the earnings value to $\$ 1$.

[^4]:    ${ }^{58}$ The estimated AKM firm effects do not vary during the period, but workers can and do change employers.
    ${ }^{59}$ The earnings observation we used for classification are labeled "previous year" in the tables.

[^5]:    ${ }^{60}$ Notice that the non-firm component of earnings declines as we move up the firm type distribution. Although it is unclear exactly which covariate is primarily responsible for this decline (fewer hours worked during the year perhaps), the impact of working at a higher paying firm would be much greater if the non-firm component of earnings were the same across firm types.

